

# Mean Field Asymptotics in Statistical Learning.

Feb 3rd.

Lecture 5 .  $\mathbb{Z}_2$  synchronization and the free energy approach.

① Asymptotic risk in  $\mathbb{Z}_2$  synchronization.

a) Model setup.

$$\theta = (\theta_1, \theta_2, \dots, \theta_n)^T \in \mathbb{R}^n, \quad \theta_i \stackrel{i.i.d.}{\sim} \text{Unif}(\mathbb{Z}_2 = \{\pm 1\}) \quad \theta \in \mathbb{D} = \{\pm 1\}^n.$$

$$Y = \frac{1}{n} \theta \theta^T + W \in \mathbb{R}^{n \times n}, \quad W \sim \text{GOE}(n).$$

$$W_{ij} \sim_{i.i.d.} N(0, \frac{1}{n}), \quad 1 \leq i < j \leq n, \quad W_{ii} \sim_{i.i.d.} N(0, \frac{2}{n}), \quad 1 \leq i \leq n, \quad W_{ij} = W_{ji}$$

Observe  $Y \in \mathbb{R}^{n \times n}$ , estimate  $\theta$ .

$$\text{Expected risk : } R(\hat{\theta}) = \mathbb{E}_{\substack{\theta \sim \text{Unif}(\mathbb{Z}_2^n) \\ Y \sim P(Y|\theta)}} \| \hat{\theta}(Y) - \theta \theta^T \|_F^2 / n^2$$

★ Connection to stochastic block model.

People form two groups (more generally,  $k$  groups).

People from same group form an edge w/ prob  $P$  independently  
 People from different group form an edge w/ prob  $q$  independently.  $(P > q)$   $P = \frac{a}{n}$   
 $q = \frac{b}{n}$ .

Observe the adjacency matrix  $A$ . Infer the two groups.

$\theta \sim \text{Unif}(\{\pm 1\}^n)$ . Give  $\theta$ , generate a graph

$$G = (V, E), \quad V = \{1, 2, \dots, n\},$$

$$V_+ = \{i \in V : \theta_i = +1\}, \quad V_- = \{i \in V : \theta_i = -1\}.$$

$$A_{ij} \sim \begin{cases} \text{Ber}(P) & i, j \text{ in the same group} \\ \text{Ber}(q) & i, j \text{ in different group.} \end{cases} \quad A_{ii} = 0.$$

$$\mathbb{E}[A | \theta] = \frac{1}{2}(P+q)I + \frac{1}{2}(P-q)\theta\theta^T - P I.$$

$$Y \equiv A - \frac{1}{2}(P+q)I - \frac{1}{2}(P-q)\theta\theta^T + P I. \quad \text{then} \quad \mathbb{E}[Y | \theta] = \frac{1}{2}(P-q)\theta\theta^T = \frac{a-b}{2n}\theta\theta^T$$

$$W = Y - \frac{a-b}{2n}\theta\theta^T \quad (\text{mean 0 noise}).$$

$$\text{Var}(W_{ij}) = \frac{1}{2}P(1-P) + \frac{1}{2}q(1-q) \approx \frac{a+b}{2n}$$

$$Y = \frac{a-b}{2n}\theta\theta^T + W$$

$$\text{Effective SNR: } \lambda \equiv \frac{a-b}{\sqrt{2(a+b)}}$$

b) Estimators in  $\mathbb{Z}_2$  sync

$$\text{MLE} : \hat{\theta}_{\text{ML}}(Y) = \underset{\sigma \in \{\pm 1\}^n}{\operatorname{argmax}} \langle \sigma, Y \sigma \rangle. \quad \begin{matrix} \text{Computationally} \\ \text{NP hard.} \end{matrix}$$

$$\{\pm 1\}^n \subseteq S^{n-1}(\sqrt{n})$$

$$\text{Spectral estimator} : \hat{\theta}_{\text{spec}}(Y) = \underset{\sigma \in S^{n-1}(\sqrt{n})}{\operatorname{argmax}} \langle \sigma, Y \sigma \rangle = \sqrt{n} v_{\max}(Y). \quad \begin{matrix} \uparrow \\ \text{leading eigenvector.} \end{matrix}$$

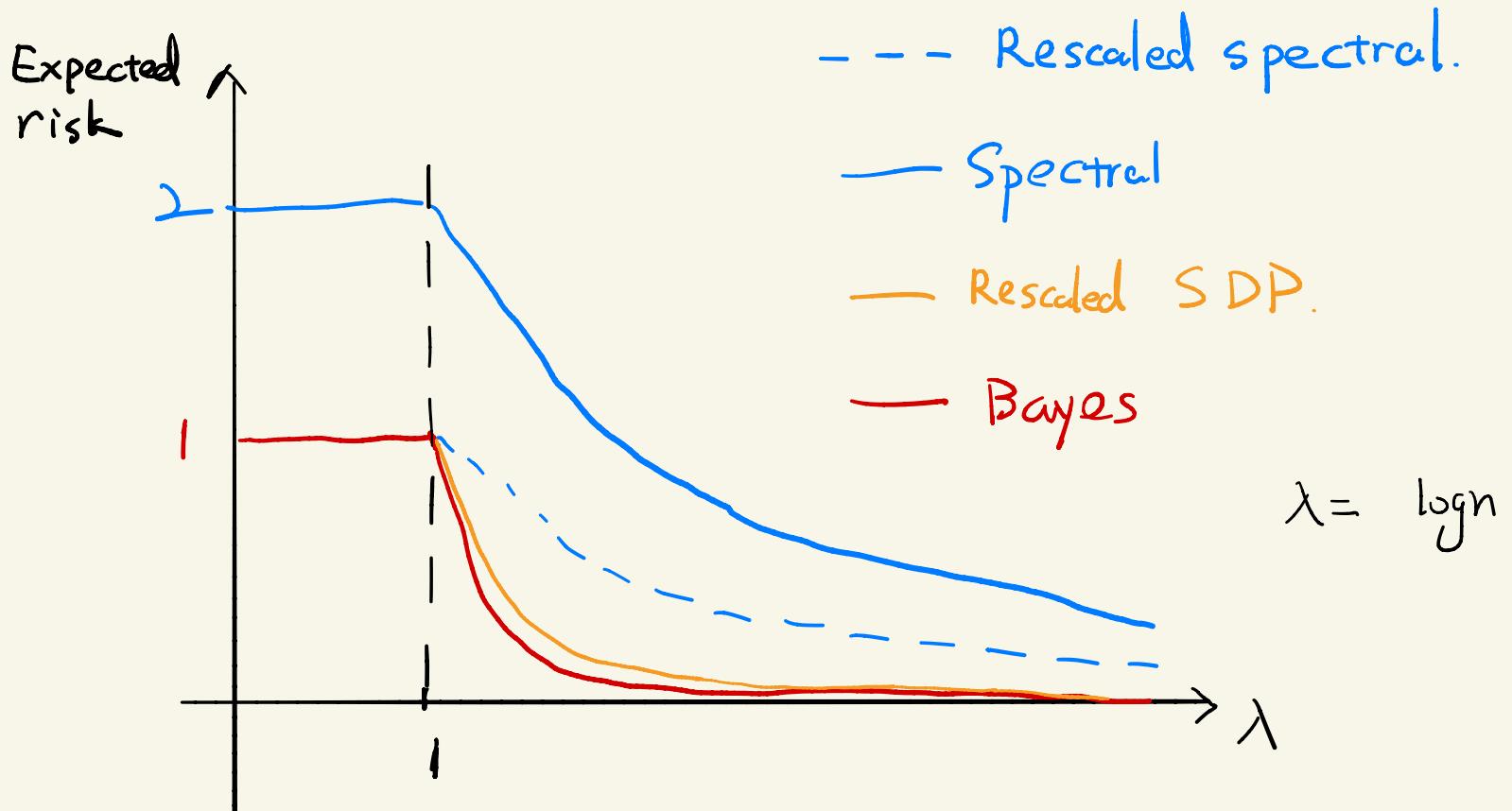
$$\hat{\Theta}_{\text{spec}}(Y) = \hat{\theta}_{\text{spec}} \hat{\theta}_{\text{spec}}^T \in \mathbb{R}^{n \times n}.$$

$$\text{SDP estimator} : \hat{\Theta}_{\text{SDP}}(Y) = \underset{\substack{X \\ \text{s.t.} \\ X \geq 0 \\ X_{ii}=1 \\ (\text{rank}(X)=1)}}{\operatorname{argmax}} \langle Y, X \rangle$$

$$\text{Bayes estimator} : \hat{\Theta}_{\text{Bayes}}(Y) = \mathbb{E} [\theta \theta^T | Y] = \sum_{\sigma \in \{\pm 1\}^n} \sigma \sigma^T p(\sigma | Y)$$

$$p(\sigma | Y) \propto \exp \left\{ \lambda \langle \sigma, Y \sigma \rangle / 2 \right\}.$$

c) Expected risk  $\mathbb{E} \| \hat{\Theta}(Y) - \theta \theta^T \|_F^2 / n^2$  v.s.  $\lambda$ .



(No statistical-computational gap in this model)

## d) Asymptotic formula.

Proposition :

(1) Spectral estimator

$$\lim_{n \rightarrow \infty} \langle \hat{\theta}_{\text{spec}}, \theta \rangle^2 / n^2 = \begin{cases} 0, & \text{for } \lambda \leq 1 \\ 1 - \frac{1}{\lambda^2}, & \text{for } \lambda > 1. \end{cases}$$

BBP phase transition.

almost surely convergence.

$$\lim_{n \rightarrow \infty} \| \hat{\theta}_{\text{spec}} \hat{\theta}_{\text{spec}}^\top - \theta \theta^\top \|_F^2 / n = \begin{cases} 2 & \text{for } \lambda \leq 1. \\ \frac{2}{\lambda^2} & \text{for } \lambda > 1 \end{cases}$$

To show this

- ① Show concentration.
- ② Calculate limit.

(2) Bayes estimator.

$$\lim_{n \rightarrow \infty} \langle \theta, \hat{\theta}_{\text{Bayes}} \theta \rangle / n^2 = \begin{cases} 0, & \text{for } \lambda \leq 1 \\ q_x(\lambda)^2 & \lambda > 1 \end{cases}$$

$$\lim_{n \rightarrow \infty} \| \hat{\theta}_{\text{Bayes}} - \theta \theta^\top \|_F^2 / n^2 = \begin{cases} 1 & \text{for } \lambda \leq 1 \\ 1 - q_x(\lambda)^2, & \text{for } \lambda > 1. \end{cases}$$

where  $q_x(\lambda)$  is the unique non-negative solution of

$$q = \mathbb{E}_{G \sim N(0,1)} [\tanh((\lambda^2 q + \lambda \sqrt{q} G)^2)].$$

$$\tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}}$$

② Derivation of asymptotic risk.

Quantity of interest:

$$\begin{aligned} & \lim_{n \rightarrow \infty} \mathbb{E} [\| \hat{\theta}_{\text{Bayes}}(Y) - \theta \theta^\top \|_F^2] / n^2 \\ &= 1 - 2a_x + b_x \end{aligned}$$

$$a_x = \lim_{n \rightarrow \infty} \mathbb{E} [\langle \theta, \hat{\theta}_{\text{Bayes}} \theta \rangle] / n^2.$$

$$b_x = \lim_{n \rightarrow \infty} \mathbb{E} [\| \hat{\theta}_{\text{Bayes}}(Y) \|_F^2] / n^2$$

General recipe for calculating  $m_x = \lim_{n \rightarrow \infty} \mathbb{E} [?]$ .  
Free energy approach.

(1). Find a configuration space  $\Omega$ ,  $v_0$ .

a observable  $M: \Omega \rightarrow \mathbb{R}$ .

a perturbed Hamiltonian  $H_\lambda: \Omega \rightarrow \mathbb{R}$ .

$$H_\lambda(\sigma) = H_0(\sigma) + \lambda M(\sigma)$$

$$P_{\beta, \lambda}(\sigma) \propto \exp\{-\beta H_\lambda(\sigma)\}.$$

s.t.  $\mathbb{E}[?] = \langle M \rangle_{\beta, \lambda} / n$  for some  $\beta$  and  $\lambda$ .

$$(2) f(\beta, \lambda) = \lim_{n \rightarrow \infty} -\frac{1}{n\beta} \mathbb{E} [\log \int \exp\{-\beta H_\lambda(\sigma)\} v_0(d\sigma)].$$

Calculate this analytic.

$$(3) m_x = \lim_{n \rightarrow \infty} \mathbb{E} [\langle M \rangle_{\beta, \lambda}] / n = \partial_\lambda f(\beta, \lambda).$$

$$\begin{aligned}
a) \text{ (1)} \quad a_* &= \lim_{n \rightarrow \infty} \mathbb{E} [ \langle \theta, \hat{\theta}_{\text{Bayes}} \rangle ] / n^2 \\
&= \lim_{n \rightarrow \infty} \mathbb{E} [ \langle \theta, \underbrace{\sum_{\sigma \in \{\pm 1\}^n} \sigma \sigma^\top P(\sigma | Y) \theta}_{\downarrow} \rangle ] / n^2 \\
&= \sum_{\sigma \in \{\pm 1\}^n} (\langle \sigma, \theta \rangle^2 / n^2) P(\sigma | Y).
\end{aligned}$$

$$\begin{aligned}
P(\sigma | Y) &\propto \exp \{ \lambda \langle \sigma, Y \sigma \rangle / 2 \} \\
&= \exp \{ -\lambda [ -\langle \sigma, W \sigma \rangle / 2 - \frac{\lambda}{2n} \langle \theta, \sigma \rangle^2 ] \}.
\end{aligned}$$

Define  $\Omega = \{\pm 1\}^n$ ,  $v_0 = \text{Unif.}$

$$M(\sigma) = -\langle \sigma, \theta \rangle^2 / (2n)$$

$$H_\lambda(\sigma) = -\langle \sigma, W \sigma \rangle / 2 + \lambda M(\sigma).$$

$$P_{\beta, \lambda}(\sigma) \propto \exp \{ -\beta H_\lambda(\sigma) \}.$$

$$P(\sigma | Y) = P_{\beta, \lambda}(\sigma) \Big|_{\beta=\lambda}.$$

$$a_* = \lim_{n \rightarrow \infty} -2 \cdot \mathbb{E} [\langle M \rangle_{\lambda, \lambda}] / n$$

$$(2) \quad f(\beta, \lambda) = \lim_{n \rightarrow \infty} -\frac{1}{n\beta} \mathbb{E} [\log \exp(-\beta H_\lambda(\sigma))]$$

$$f(\beta, \lambda) = \max_{b, q} f_{mf}(b, q; \beta, \lambda)$$

$$\begin{aligned}
f_{mf}(b, q; \beta, \lambda) &\equiv -\frac{1}{4} \beta (1-q)^2 + \frac{1}{2} \lambda b^2 - \frac{1}{\beta} \mathbb{E} [\log 2 \cosh(\beta(\lambda b + \sqrt{q} G))] \\
&\quad \left. \begin{array}{l} \\ \uparrow \\ G \sim N(0, 1). \end{array} \right. \\
\begin{cases} b_* = \mathbb{E} [\tanh(\beta(\lambda b_* + \sqrt{q_*} G))] \\ q_* = \mathbb{E} [\tanh(\beta(\lambda_* + \sqrt{q_*} G))^2] \end{cases}
\end{aligned}$$

$$\begin{aligned}
m_*(\beta, \lambda) &= \partial_\lambda f(\beta, \lambda) = \frac{1}{2} b^2 - \frac{1}{\beta} \mathbb{E} [\tanh(\beta(\lambda b + \sqrt{q} G))] \cdot \cancel{\beta b} \Big|_{b=q} \\
&= \frac{1}{2} b_*^2 - b_*^2 = -\frac{1}{2} b_*^2
\end{aligned}$$

$$C_* = -2 m_*(\beta, \lambda) \Big|_{\beta=\lambda} = b_*^2(\lambda, \lambda) = q_*^2(\lambda)$$

$$q_* = \mathbb{E} [\tanh(\beta(\lambda_* + \sqrt{q_*} G))^2].$$

$$f_*(\lambda) = \max_q f(q, \lambda) \quad \text{implicit differentiation}$$

$$\partial_\lambda f_*(\lambda) = \partial_\lambda \left[ \max_q f(q, \lambda) \right] = \partial_\lambda f(q, \lambda) \Big|_{q=q_*}.$$

$$q_* = \arg \max_q f(q, \lambda)$$

$$(b) \quad b_{\infty} = \lim_{n \rightarrow \infty} \mathbb{E} [\|\hat{\mathbb{H}}_{\text{Bayes}}(Y)\|_F^2] / n^2.$$

$$\begin{aligned} & \|\hat{\mathbb{H}}_{\text{Bayes}}\|_F^2 / n^2 \quad (\mathbb{H} = \{\pm 1\}^n) \\ &= \left\langle \sum_{\sigma_1 \in \mathbb{H}} \sigma_1 \sigma_1^\top P(\sigma_1 | Y), \sum_{\sigma_2 \in \mathbb{H}} \sigma_2 \sigma_2^\top P(\sigma_2 | Y) \right\rangle / n^2 \\ &= \sum_{\substack{\sigma_1 \in \mathbb{H} \\ \sigma_2 \in \mathbb{H}}} \langle \sigma_1 \sigma_1^\top, \sigma_2 \sigma_2^\top \rangle P(\sigma_1 | Y) P(\sigma_2 | Y) / n^2. \\ &= \sum_{(\sigma_1, \sigma_2) \in \mathbb{H} \times \mathbb{H}} \langle \sigma_1, \sigma_2 \rangle^2 \mu(\sigma_1, \sigma_2 | Y) / n^2 \\ &\sigma = (\sigma_1, \sigma_2) \in \mathbb{H} \times \mathbb{H}. \end{aligned}$$

$$\begin{aligned} \mu(\sigma_1, \sigma_2 | Y) &= P(\sigma_1 | Y) P(\sigma_2 | Y) \\ &\propto \exp \{ \lambda \langle \sigma_1, Y \sigma_1 \rangle / 2 + \lambda \langle \sigma_2, Y \sigma_2 \rangle / 2 \}. \end{aligned}$$

$$\text{Define } \Omega = \mathbb{H} \times \mathbb{H}. \quad v_0 \sim \text{Unif}(\mathbb{H}) \times \text{Unif}(\mathbb{H}).$$

$$\sigma \in \Omega \quad \sigma = (\sigma_1, \sigma_2) \quad \sigma_1, \sigma_2 \in \mathbb{H}.$$

$$M(\sigma) = \langle \sigma_1, \sigma_2 \rangle^2 / n.$$

$$\begin{aligned} H_{\lambda, h}(\sigma) &= -\frac{1}{2} \langle \sigma_1, W \sigma_1 \rangle - \frac{\lambda}{2n} \langle \sigma_1, \theta \rangle^2 \\ &\quad - \frac{1}{2} \langle \sigma_2, W \sigma_2 \rangle - \frac{\lambda}{2n} \langle \sigma_2, \theta \rangle^2 \\ &\quad + h \langle \sigma_1, \sigma_2 \rangle^2 / n. \end{aligned}$$

$$P_{\beta, \lambda, h}(\sigma) \propto \exp \{ -\beta H_{\lambda, h}(\sigma) \}$$

$$b_{\infty} = \lim_{n \rightarrow \infty} \mathbb{E} [\langle M \rangle_{\beta, \lambda, h}] \Big|_{\substack{\beta=\lambda \\ h=0}}.$$

$$(2). \quad f(\beta, \lambda, h) = \text{formula} = \max_{\eta, \gamma} f_{\text{inf}}(\eta, \gamma; \beta, \lambda, h).$$

$$b_{\infty} = \partial_h f(\beta, \lambda, h) \Big|_{\substack{\beta=\lambda \\ h=0}}$$

$$b_{\infty} = q_{\infty}(\lambda)^2.$$

$$\textcircled{3} \quad m_{\hat{\theta}} = \lim_{n \rightarrow \infty} \mathbb{E} [ \langle \hat{\theta}_{\text{spec}}(Y), \theta \rangle^2 / n^2 ].$$

$$\hat{\theta}_{\text{spec}}(Y) = \sup_{\sigma \in S^{n-1}(\mathbb{R})} \langle \sigma, Y \cdot \sigma \rangle.$$

$$M(\sigma) = \langle \sigma, \theta \rangle^2 / n.$$

$$H_{\lambda}(\sigma) = -\langle \sigma, Y \cdot \sigma \rangle / 2$$

$$P_{\beta, \lambda, h}(\sigma) \propto \exp \{ -\beta [H_{\lambda}(\sigma) + h M(\sigma)] \}.$$

$$\begin{aligned} m_{\hat{\theta}} &= \lim_{n \rightarrow \infty} \mathbb{E} [ M(\hat{\theta}_{\text{spec}}) ] / n \\ &= \lim_{n \rightarrow \infty} \mathbb{E} [ M(\arg \min_{\sigma} H_{\lambda}(\sigma)) ] / n \\ &= \lim_{n \rightarrow \infty} \lim_{\beta \rightarrow \infty} \mathbb{E} [ \langle M \rangle_{\beta, \lambda, h} ] / n \Big|_{h=0} \\ &\stackrel{?}{=} \lim_{\beta \rightarrow \infty} \lim_{n \rightarrow \infty} \mathbb{E} [ \langle M \rangle_{\beta, \lambda, h} ] / n \Big|_{h=0} \\ &= \lim_{\beta \rightarrow \infty} \lim_{n \rightarrow \infty} \mathbb{E} [ \partial_h F(\beta, \lambda, h) ] / n \Big|_{h=0} \\ &= \lim_{\beta \rightarrow \infty} \partial_h f(\beta, \lambda, h) \Big|_{h=0}. \end{aligned}$$

Exercise : Suppose  $\beta_0 \sim N(0, \sigma_0^2 I_d)$ .

$$Y = X\beta_0 + \varepsilon, \quad X_{ij} \sim \text{i.i.d. } N(0, \frac{1}{d}), \quad \varepsilon_i \sim \text{i.i.d. } N(0, \sigma^2).$$

$$\text{Denote } \hat{\beta} = \arg \min_{\beta} \|Y - X\beta\|_2^2 + \lambda \|\beta\|_2^2 = (X^T X + \lambda I)^{-1} X^T Y$$

\textcircled{1} Figure out a configuration space  $\Omega$ ,  $\nu_0$   
an observable  $M: \Omega \rightarrow \mathbb{R}$

and a perturbed Hamiltonian  $H_{\lambda}: \Omega \rightarrow \mathbb{R}$ .

$$\text{st. defining } F(\beta, \lambda) = -\frac{1}{\beta} \log \int \exp\{-\beta H_{\lambda}(\sigma)\} \nu_0(d\sigma).$$

$$\text{we have } \langle \hat{\beta}, \beta_0 \rangle / n = \lim_{\beta \rightarrow 0} \partial_{\lambda} \lim_{n \rightarrow \infty} F(\beta, \lambda) / n$$

\textcircled{2} Do similar things for  $\|\hat{\beta}\|_2^2 / n$ .

\textcircled{3} Hopefully, your  $H_{\lambda}(\sigma)$  is a quadratic function of  $\sigma \in \mathbb{R}^d$   
and  $\nu_0(\sigma)$  is Lebesgue measure.

In this case,  $\int \exp\{-\beta H_{\lambda}(\sigma)\} \nu_0(d\sigma)$  is a Gaussian  
integration, and can be written out explicitly.

Please simplify  $\mathbb{E}_{X, \beta} [F(\beta, \lambda)]$  as much as possible

