

Lecture 13 : Gaussian comparison inequalities.

- ① Motivation.
- ② Slepian's inequality and Sudakov-Fernique inequality.
- ③ Examples: Wishart matrix and spiked GOE matrix.
- ④ Proof of Slepian's inequality if time permitted.

- ① Motivating example: the LASSO problem

Signal: $x_0 \in \mathbb{R}^d$, Noise: $w \in \mathbb{R}^n$, Sensing matrix: $A \in \mathbb{R}^{n \times d}$.

Response: $y = Ax_0 + w \in \mathbb{R}^n$, $A_{ij} \sim \text{iid } N(0, \frac{1}{n})$, $w_i \sim \text{iid } N(0, \sigma^2)$, important.

LASSO estimator: $\hat{x} = \arg \min_x \frac{1}{2n} \|y - Ax\|_2^2 + \frac{\lambda}{n} \|x\|_1$,

Goal: derive $m_x \equiv \lim_{d \rightarrow \infty} \mathbb{E} \left[\frac{1}{d} \sum_{i=1}^d \psi(\hat{x}_i, x_{0,i}) \right]$. when $n/d \rightarrow s$.

Free energy approach ($\beta = \infty$): the perturbed LASSO objective.

$$f(h) \equiv \lim_{d \rightarrow \infty} \mathbb{E} \left[\min_x \left\{ \frac{1}{2n} \|y - Ax\|_2^2 + \frac{\lambda}{n} \|x\|_1 + h \frac{1}{d} \sum_{i=1}^d \psi(x_i, x_{0,i}) \right\} \right].$$

$$\Rightarrow f'(0) = m_x. \quad (\text{need to argue why limit and derivative can be exchanged})$$

Question: how to derive $f(h)$?

$$\|u\|_2^2 = \sup_v [2\langle u, v \rangle - \|v\|^2]$$

$$\text{Intuition: } F = \min_x \left\{ \frac{1}{2} \|y - Ax\|_2^2 + \lambda \|x\|_1 \right\}.$$

$$= \min_x \max_v \underbrace{\langle A(x_0 - x) + w, v \rangle}_{\substack{\uparrow \\ \text{fixed}}} - \frac{1}{2} \|v\|_2^2 + \lambda \|x\|_1$$

Gaussians

$$\Phi(x, v) \equiv \langle A(x_0 - x) + w, v \rangle - \frac{1}{2} \|v\|_2^2 + \lambda \|x\|_1$$

Gaussian process on $\mathbb{R}^d \times \mathbb{R}^n$ indexed by x, v .

Need some machinery to analyze the extremum of Gaussian processes.

Gaussian comparison inequalities.

$\mathbb{E}[\text{ext Gaussian process}]$

② Slepian's inequality and its applications.

Definition (Gaussian process). A stochastic process $(X_t)_{t \in T}$

is called Gaussian process, if \forall finite subset $T_0 \subseteq T$,
the random vector $(X_t)_{t \in T_0}$ has normal distribution.

Remark: a Gaussian process can be completely characterized by its mean function $\mu(t) = \mathbb{E}[X_t]$ and covariance function $\Sigma(t, s) = \mathbb{E}[X_t X_s]$. We will often consider continuous Gaussian process where X_t is a continuous function of t a.s., and T is a separable space. This leads to $\mathbb{E}[\sup_{t \in T} X_t] = \sup_{\substack{T_0 \subseteq T \\ \text{finite}}} \mathbb{E}[\sup_{t \in T_0} X_t]$.

Example: $X_t = \langle t, g \rangle$, $t \in S^{n-1}$, $g \sim N(0, I_d)$.

$$\mathbb{E}[\sup_t X_t] = \mathbb{E}[\|g\|_2] \approx \sqrt{d}. \quad W(S) = \mathbb{E}[\sup_{t \in S} \langle t, g \rangle]$$

Example: $A \in \mathbb{R}^{n \times m}$, $A_{ij} \stackrel{iid}{\sim} N(0, 1)$.

$$\mathbb{E}[\|A\|_{op}] = \mathbb{E}[\sup_{\substack{u \in S^{n-1} \\ v \in S^{m-1}}} \langle u, Av \rangle]$$

$$X_{uv} = \langle u, Av \rangle = \langle A, uv^T \rangle. \quad T = S^{n-1} \times S^{m-1}.$$

Thm [Slepian's inequality and Sudakov-Fernique inequality]

Let $T \subseteq \mathbb{R}^d$, Let $(X_t)_{t \in T}$ $(Y_t)_{t \in T}$ be two GP.

with $\mathbb{E}X_t = \mathbb{E}Y_t \quad \forall t \in T$. (to ①).

and $\mathbb{E}[(X_t - X_s)^2] \leq \mathbb{E}[(Y_t - Y_s)^2]$. Then

$$\textcircled{1} \quad \mathbb{E}[\sup_{t \in T} X_t] \leq \mathbb{E}[\sup_{t \in T} Y_t]. \quad [\text{Sudakov-Fernique}]$$

$$\textcircled{2} \quad \text{If additionally } \mathbb{E}[X_t^2] = \mathbb{E}[Y_t^2] \quad \forall t \in T.$$

$$\Rightarrow \forall z \in \mathbb{R}, \quad \mathbb{P}(\sup_{t \in T} X_t \geq z) \leq \mathbb{P}(\sup_{t \in T} Y_t \geq z) \quad [\text{Slepian}].$$

③ Example 1. (Norm of Gaussian random matrix).

Proposition: $A \in \mathbb{R}^{m \times n}$, $A_{ij} \sim \text{iid } N(0, 1)$.

Then $\mathbb{E}[\|A\|_{op}] \leq \sqrt{n} + \sqrt{m}$.

$$\mathbb{E}[\|A/\sqrt{n}\|_{op}] \leq 1 + \sqrt{\frac{m}{n}} \quad m \propto n.$$

Proof: $\mathbb{E}[\|A\|_{op}] = \mathbb{E}[\sup_{\substack{u \in S^{n-1} \\ v \in S^{m-1}}} \underbrace{\langle u, Av \rangle}_{X_{uv}}]$.

$$\mathbb{E}[X_{uv}] = 0$$

$$\mathbb{E}[(X_{uv} - X_{wz})^2] = \mathbb{E}[(\langle A, vu^\top - zw^\top \rangle)^2]$$

$$\mathbb{E}[\langle g, t \rangle^2] = \|v u^\top - z w^\top\|_F^2 = 2 - 2\langle v, z \rangle \langle u, w \rangle.$$

$$= \|t\|_2^2$$

Define $(Y_{uv})_{(u,v) \in T}$: $Y_{uv} = \langle g, u \rangle + \langle h, v \rangle$
 $g \sim N(0, I_m)$, $h \sim N(0, I_n)$ $g \perp h$.

$$\mathbb{E}[Y_{uv}] = 0$$

$$\begin{aligned}\mathbb{E}[(Y_{uv} - Y_{wz})^2] &= \mathbb{E}[(\langle g, u-w \rangle + \langle h, v-z \rangle)^2] \\ &= \mathbb{E}[\langle g, u-w \rangle^2] + \mathbb{E}[\langle h, v-z \rangle^2] \\ &= \|u-w\|_2^2 + \|v-z\|_2^2 \\ &= 4 - 2\langle u, w \rangle - 2\langle v, z \rangle.\end{aligned}$$

$$\begin{aligned}\mathbb{E}[(Y_{uv} - Y_{wz})^2] - \mathbb{E}[(X_{uv} - X_{wz})^2] \\ &= 2(1 - \langle u, w \rangle - \langle v, z \rangle + \langle u, w \rangle \langle v, z \rangle) \\ &= 2(1 - \langle u, w \rangle)(1 - \langle v, z \rangle) \geq 0\end{aligned}$$

By Sudakov-Fernique

$$\begin{aligned}\mathbb{E}[\|A\|_{op}] &\leq \mathbb{E}\left[\sup_{\substack{u \in S^{m-1} \\ v \in S^{n-1}}} \langle g, u \rangle + \langle h, v \rangle\right]. \\ &= \mathbb{E}\left[\sup_{u \in S^{m-1}} \langle g, u \rangle\right] + \mathbb{E}\left[\sup_{v \in S^{n-1}} \langle h, v \rangle\right] \\ &= \mathbb{E}[\|g\|_2] + \mathbb{E}[\|h\|_2] \leq \sqrt{m} + \sqrt{n}.\end{aligned}$$

□

Example 2: (The Spiked GOE matrix). $W = (G + G^\top)/\sqrt{2n}$

Prop : $\theta \in S^{n-1}$, $\lambda \geq 0$. $W \sim \text{GOE}(n)$. ($W_{ii} \sim N(0, \frac{2}{n})$, $W_{ij} \sim N(0, \frac{1}{n})$)

$Y = \lambda \theta \theta^\top + W \in \mathbb{R}^{n \times n}$. Denote $\lambda_{\max}(Y) = \sup_{\sigma \in S^{n-1}} \langle \sigma, Y \sigma \rangle$.

Then $\limsup_{n \rightarrow \infty} \mathbb{E}[\|Y\|_{op}] \leq \begin{cases} 2, & \lambda \leq 1. \\ \lambda + \frac{1}{\lambda}, & \lambda \geq 1. \end{cases}$

(replica $\lim_{n \rightarrow \infty} \mathbb{E}[\|Y\|_{op}] = \uparrow$).

Proof : $\lambda_{\max}(Y) = \sup_{t \in S^{m-1}} \langle t, Y t \rangle = \sup_{t \in S^{m-1}} \underbrace{\langle t, W t \rangle}_{X_t} + \lambda \langle t, \theta \rangle^2$.

$$\mathbb{E} X_t = \lambda \langle t, \theta \rangle^2.$$

$$\mathbb{E}[(X_t - X_s)^2] = \frac{4}{n}(1 - \langle t, s \rangle^2).$$

$$(Y_t)_{t \in T} : Y_t = \frac{2}{\sqrt{n}} \langle g, t \rangle + \lambda \langle t, \theta \rangle^2, \quad g \sim N(0, I_n)$$

$$\mathbb{E}[Y_t] = \lambda \langle t, \theta \rangle^2.$$

$$\mathbb{E}[(Y_t - Y_s)^2] = \frac{4}{n}(2 - 2\langle t, s \rangle).$$

$$\mathbb{E}[(Y_t - Y_s)^2] - \mathbb{E}[(X_t - X_s)^2] = \frac{4}{n}(1 - \langle t, s \rangle^2) \geq 0.$$

$$\Rightarrow \mathbb{E}\|Y\|_{op} \leq \mathbb{E}\left[\sup_{t \in S^{n-1}} \frac{2}{\sqrt{n}} \langle g, t \rangle + \lambda \langle \theta, t \rangle^2\right].$$

① $\lambda = 0$.

$$\mathbb{E}\left[\sup_{t \in S^{n-1}} \frac{2}{\sqrt{n}} \langle g, t \rangle\right] = \frac{2}{\sqrt{n}} \mathbb{E}\|\langle g \rangle\|_2 \leq 2.$$

$$\begin{aligned} & \sup_{t \in S^{n-1}} \left[\frac{2}{\sqrt{n}} \langle g, t \rangle + \lambda \langle \theta, t \rangle^2 \right] \\ &= \inf_{\zeta \geq \lambda} \sup_{t \in \mathbb{R}^n} \frac{2}{\sqrt{n}} \langle g, t \rangle + \lambda \langle \theta, t \rangle^2 - \zeta (\|t\|_2^2 - 1) \quad \zeta < \lambda \\ &= \inf_{\zeta > \lambda} \left[\frac{1}{n} \langle g, (\zeta I - \lambda \theta \theta^\top)^{-1} g \rangle + \zeta \right]. \end{aligned}$$

$$\mathbb{E}\|Y\|_{op} \leq \mathbb{E}\left[\inf_{\zeta > \lambda} \left(\frac{1}{n} \langle g, (\zeta I - \lambda \theta \theta^\top)^{-1} g \rangle + \zeta \right)\right]$$

$$\begin{aligned} & \leq \inf_{\zeta > \lambda} \mathbb{E}\left[\frac{1}{n} \langle g, (\zeta I - \lambda \theta \theta^\top)^{-1} g \rangle + \zeta \right] \\ &= \inf_{\zeta > \lambda} \underbrace{\frac{1}{n} \text{tr}((\zeta I - \lambda \theta \theta^\top)^{-1})}_{\frac{1}{n} \text{tr}((\zeta I - \lambda \theta \theta^\top)^{-2})} + \zeta. \end{aligned}$$

$$\boxed{\frac{1}{n} \text{tr}((\zeta I - \lambda \theta \theta^\top)^{-2}) = 1.}$$

— When $\lambda < 1$, we take $\zeta_x = 1$.

$$\mathbb{E}\|Y\|_{op} \leq \frac{1}{n} \text{tr}((1 - \lambda \theta \theta^\top)^{-1}) + 1 = 2 + \frac{\lambda}{n(1-\lambda)}.$$

$$\limsup_{n \rightarrow \infty} \mathbb{E}\|Y\|_{op} \leq 2.$$

— When $\lambda \geq 1$, we take $\zeta_x = \lambda + \frac{1}{\sqrt{n}}$.

$$\mathbb{E}\|Y\|_{op} \leq \frac{n-1}{n} \frac{1}{\lambda + \frac{1}{\sqrt{n}}} + \frac{\sqrt{n}}{n} + \lambda + \frac{1}{\sqrt{n}}.$$

$$\limsup_{n \rightarrow \infty} \mathbb{E}\|Y\|_{op} \leq \lambda + \frac{1}{\lambda}. \quad \square$$

④ Proof of Slepian's inequality.

[Slepian 1962]

(Sudakov 1970s)
(Fernique)

* Lemma: [Kahane, 1986].

Let $X = (X_i)_{i \in [n]}$ and $Y = (Y_i)_{i \in [n]}$ be two Gaussian vectors, with $\mathbb{E}X = \mathbb{E}Y$.

Denote $A \equiv \{(j, k) \in [n]^2 : \mathbb{E}[X_j X_k] < \mathbb{E}[Y_j Y_k]\}$.

$B \equiv \{(j, k) \in [n]^2 : \mathbb{E}[X_j X_k] > \mathbb{E}[Y_j Y_k]\}$.

Let $f(x) = f(x_1, \dots, x_n) \in C^2(\mathbb{R}^n)$ satisfies $(\mathbb{E}[\partial_i \partial_j f(x)])$ can be well defined for Gaussian X .

$$\begin{cases} \partial_j \partial_k f(x) \geq 0, & (j, k) \in A, \\ \partial_j \partial_k f(x) \leq 0, & (j, k) \in B. \end{cases}$$

Then we have $\mathbb{E}[f(X)] \leq \mathbb{E}[f(Y)]$.

Proof idea: interpolation method.

Proof: W.L.O.G, we assume that $\mathbb{E}[X] = \mathbb{E}[Y] = 0$.

Define $Z(\theta) = X \cos \theta + Y \sin \theta$ for $0 \leq \theta \leq \frac{\pi}{2}$.

and $\varphi(\theta) = \mathbb{E}[f(Z(\theta))]$.

$$\Rightarrow \varphi(0) = \mathbb{E}[f(X)], \quad \varphi(1) = \mathbb{E}[f(Y)].$$

Hope to show $\varphi'(\theta) \geq 0$, $\forall \theta \in [0, \pi/2]$.

$$\varphi'(\theta) = \mathbb{E}[\langle \nabla f(Z(\theta)), Z'(\theta) \rangle]$$

$$= \sum_{i=1}^N \left[-\sin \theta \mathbb{E}[X_i \partial_i f(Z(\theta))] + \cos \theta \cdot \mathbb{E}[Y_i \partial_i f(Z(\theta))] \right].$$

Integration by part:

$$\mathbb{E}[X_i \partial_i f(X \cos \theta + Y \sin \theta)] = \cos \theta \sum_{j=1}^n \mathbb{E}[X_i X_j] \cdot \mathbb{E}[\partial_i \partial_j f(Z(\theta))].$$

$$\mathbb{E}[Y_i \partial_i f(X \cos \theta + Y \sin \theta)] = \sin \theta \sum_{j=1}^n \mathbb{E}[Y_i Y_j] \cdot \mathbb{E}[\partial_i \partial_j f(Z(\theta))].$$

$$= \cos \theta \cdot \sin \theta \cdot \sum_{i,j=1}^n \mathbb{E}[\partial_i \partial_j f(Z(\theta))] \{ \mathbb{E}[Y_i Y_j] - \mathbb{E}[X_i X_j] \} \geq 0$$

By assumption \square

Remark: If $\mathbb{E}[X] = \mathbb{E}[Y] = \mu \neq 0$, just define $\tilde{f}(x) = f(x - \mu)$.

Corollary: Let $X, Y \in \mathbb{R}^n$ be n -dimensional Gaussian random vectors, with $\mathbb{E}[X] = \mathbb{E}[Y]$. Assume that $\mathbb{E}[X_i^2] = \mathbb{E}[Y_i^2]$, $\forall i \in [n]$, and $\mathbb{E}[X_i X_j] \geq \mathbb{E}[Y_i Y_j] \quad \forall i, j \in [n]$.

Then

(1) For any real η_1, \dots, η_n ,

$$\mathbb{P}(X_i < \eta_i, \forall i) \geq \mathbb{P}(Y_i < \eta_i, \forall i),$$

(2) $\mathbb{P}(\max_{i \in [n]} X_i \geq \eta) \leq \mathbb{P}(\max_{i \in [n]} Y_i \geq \eta)$.

(3) $\mathbb{E}[\max_{i \in [n]} X_i] \leq \mathbb{E}[\max_{i \in [n]} Y_i]$.

Proof: Clearly (1) \Rightarrow (2) \Rightarrow (3).

To prove (1), conceptually we want to take

$$f(x_1, \dots, x_n) = \prod_{i=1}^n \mathbf{1}_{\{X_i < \eta_i\}} \quad \text{Not smooth.}$$

The $\psi \in C^\infty(\mathbb{R})$ with $\psi(t) = 0, t < 0, \psi(t) = 1, t \geq 1, \psi'(t) \geq 0$.

Take $f_\varepsilon(x_1, \dots, x_n) = \prod_{i=1}^n \psi((x_i - \eta_i)/\varepsilon)$. $\Rightarrow \partial_i \partial_j f_\varepsilon \geq 0$.

$$\Rightarrow \mathbb{E}[f_\varepsilon(X)] \leq \mathbb{E}[f_\varepsilon(Y)]$$

\Rightarrow send $\varepsilon \rightarrow 0$ and apply monotone convergence thm. \square

Remark: For $(X_t)_{t \in T}$ Gaussian process,

when T is finite, it directly gives the theorem.

When T is uncountable but separable (has a countable dense subset), when X_t is a continuous Gaussian process,

we have $\mathbb{E}[\sup_{t \in T} X_t] = \sup_{\substack{t_0 \subseteq T \\ \text{finite}}} \mathbb{E}[\sup_{t \in T_0} X_t]$.

⑤ Proof of Sudakov - Fernique.

Proposition [Sudakov - Fernique's inequality].

Let $X, Y \in \mathbb{R}^n$ be two Gaussian random vectors with $\mathbb{E}[X] = \mathbb{E}[Y]$, and $\mathbb{E}[(X_i - X_j)^2] \leq \mathbb{E}[(Y_i - Y_j)^2]$. Then.

$$\mathbb{E}[\max_{i \in [n]} X_i] \leq \mathbb{E}[\max_{i \in [n]} Y_i].$$

Proof: Let $\mu = \mathbb{E}X = \mathbb{E}Y$.

Define $\tilde{X} = X - \mu$, $\tilde{Y} = Y - \mu \in \mathbb{R}^n$.

and $Z(\theta) \equiv \cos\theta \cdot \tilde{X} + \sin\theta \cdot \tilde{Y} + \mu$

Fix $\beta > 0$, define $F_\beta : \mathbb{R}^n \rightarrow \mathbb{R}$

$$F_\beta(x) \equiv \beta^{-1} \log \left(\sum_{i=1}^n e^{\beta x_i} \right).$$

For $\theta \in [0, \pi/2]$, let $\varphi(t) = \mathbb{E}[F_\beta(Z(\theta))]$.

$$\varphi'(\theta) = \cos\theta \sin\theta \sum_{i,j=1}^n \mathbb{E} \left[\frac{\partial^2 F_\beta}{\partial x_j \partial x_i}(Z(\theta)) \right] \cdot \{ \mathbb{E}[\tilde{Y}_i \tilde{Y}_j] - \mathbb{E}[\tilde{X}_i \tilde{X}_j] \}$$

↑ same calculation as before

Define $P_i(x) = \partial_{x_i} F_\beta(x) = \frac{e^{\beta x_i}}{\sum_{j=1}^n e^{\beta x_j}}$, prob dist on \mathbb{R}^n .

$$\Rightarrow \frac{\partial^2 F_\beta}{\partial x_i \partial x_j}(x) = \begin{cases} \beta (P_i(x) - P_j(x))^2 & \text{if } i=j \\ -\beta P_i(x) P_j(x) & \text{if } i \neq j. \end{cases}$$

$$\begin{aligned} & \sum_{i,j=1}^n \frac{\partial^2 F_\beta}{\partial x_i \partial x_j}(x) \{ \mathbb{E}[\tilde{Y}_i \tilde{Y}_j] - \mathbb{E}[\tilde{X}_i \tilde{X}_j] \} \\ &= \beta \underbrace{\sum_{i=1}^n P_i(x) [\mathbb{E}[\tilde{Y}_i^2] - \mathbb{E}[\tilde{X}_i^2]]}_{\text{blue bracket}} - \beta \sum_{i,j=1}^n P_i(x) P_j(x) [\mathbb{E}[\tilde{Y}_i \tilde{Y}_j] - \mathbb{E}[\tilde{X}_i \tilde{X}_j]] \\ &= \frac{\beta}{2} \sum_{i,j=1}^n P_i(x) P_j(x) [\mathbb{E}[\tilde{Y}_i^2] - \mathbb{E}[\tilde{X}_i^2] + \mathbb{E}[\tilde{Y}_j^2] - \mathbb{E}[\tilde{X}_j^2]]. \\ &= \frac{\beta}{2} \sum_{i,j=1}^n P_i(x) P_j(x) [\mathbb{E}[(Y_i - Y_j)^2] - \mathbb{E}[(X_i - X_j)^2]] \geq 0. \end{aligned}$$

$$\Rightarrow \mathbb{E}[F_\beta(X)] \leq \mathbb{E}[F_\beta(Y)].$$

$$\beta \rightarrow \infty \Rightarrow \mathbb{E}[\max_{i \in [n]} X_i] \leq \mathbb{E}[\max_{i \in [n]} Y_i].$$

□

