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7 April 04

Association of (stationary) time series.

Measuring and asking if it is due to chance.

Some characteristics of dependence:

$$E(XY) \neq E(X)E(Y)$$

$$E(Y|X) = g(X)$$

$$X = g(\mathcal{B}), Y = h(\mathcal{B}) \quad \mathcal{B}: \text{latent}$$

$$f_{XY}(x, y) \neq f_X(x)f_Y(y) \quad \text{for some } (x, y)$$

Mutual information, $I_{XY} \neq 0$

$$I_{XY} = E_{X,Y} \left\{ \log \frac{f_{XY}(X,Y)}{f_X(X)f_Y(Y)} \right\}$$

For a test, need to keep alternatives in mind.

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Time series $(X(t), Y(t)), t=0, \pm 1, \pm 2, \dots$

stationary, mixing

1. Time-side statistic

$$\text{cov}\{X(t+u), Y(t)\} = c_{XY}(u)$$

Is it O at $\log u$? (Alternative $\log u$ dependence)

$$c_{XY}^T(u) = \frac{1}{T} \sum_{0 \leq t, t+u \leq T-1} [X(t+u) - \bar{X}][Y(t) - \bar{Y}]$$

asymptotic distribution

DRB book (7.6.11)

$$\begin{aligned} \frac{1}{2\pi} \text{var } c_{XY}^T(u) &\sim \int_0^{2\pi} f_{XX}(\alpha) f_{YY}(\alpha) d\alpha \\ &+ \int_0^{2\pi} e^{-i2u\alpha} |f_{XY}(\alpha)|^2 d\alpha \\ &+ \int_0^{2\pi} \int_0^{2\pi} e^{i(u\alpha - v\beta)} f_{XXYY}(\alpha, -\alpha, -\beta) d\alpha d\beta \end{aligned}$$

2 April 07

Crude result for unrelated linear processes

$$\text{var} \left\{ \frac{c_{xy}^T(\omega)}{\sqrt{c_{xx}^T(0)c_{yy}^T(0)}} \right\} \sim \frac{1}{T}$$

(but need to consider behavior under alternative)

Other way to get $\text{var } c_{xx}^T(0)$:

1. Split sample
2. Via spectrum of product $z(t) = x(t)y(t)$
3. Bootstraps

Note. Could have $c_{xy}(0) = 0$, but X and Y dependent.

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2. Frequency-side statistic.

Coherence. A measure of linear time invariant association as a function of frequency.

$$|R_{yx}(\omega)|^2 = \frac{|f_{xy}(\omega)|^2}{f_{xx}(\omega) f_{yy}(\omega)} = |R_{xy}(\omega)|^2$$

if denom $\neq 0$

= 0 if denom = 0

$$0 \leq |R_{xy}(\omega)|^2 \leq 1$$

Proof. Consider

$$d_x^T(\omega) = \sum_{t=0}^{T-1} e^{-i\omega t} x(t), \quad d_y^T(\omega) =$$

$$\text{var } d_x^T(\omega) = \int |\Delta^T(\omega - \alpha)|^2 f_{xx}(\alpha) d\alpha$$

from Gamér.

$$\sim 2\pi T f_{xx}(\omega) \quad \text{for large } T$$

$$\text{cov}\{d_x^T(\omega), d_y^T(\omega)\} = \int \Delta^T(\omega - \alpha)^2 f_{xy}(\alpha) d\alpha$$

$$\sim 2\pi T f_{xy}(\omega)$$

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But $|\text{cov}\{X, Y\}|^2 \leq \text{var} X \text{var} Y$

Schwarz

So

$$|\text{corr}\{d_X^T(\omega), d_Y^T(\omega)\}|^2 \leq 1$$

Take limit as $T \rightarrow \infty$.

(an estimate coherence via

$$\frac{|f_{xy}^T(\omega)|^2}{f_{xx}^T(\omega) f_{yy}^T(\omega)}$$

Density of large sample distribution

$$(1 - |R|^2)^m {}_2F_1(m, m; 1, |R|^2 |\hat{R}|^2) \frac{\Gamma(n)}{\Gamma(n-1)\Gamma(1)} (1 - |\hat{R}|^2)^{n-2}$$

if n periodogram ordinates averaged to obtain $\hat{f}^s$

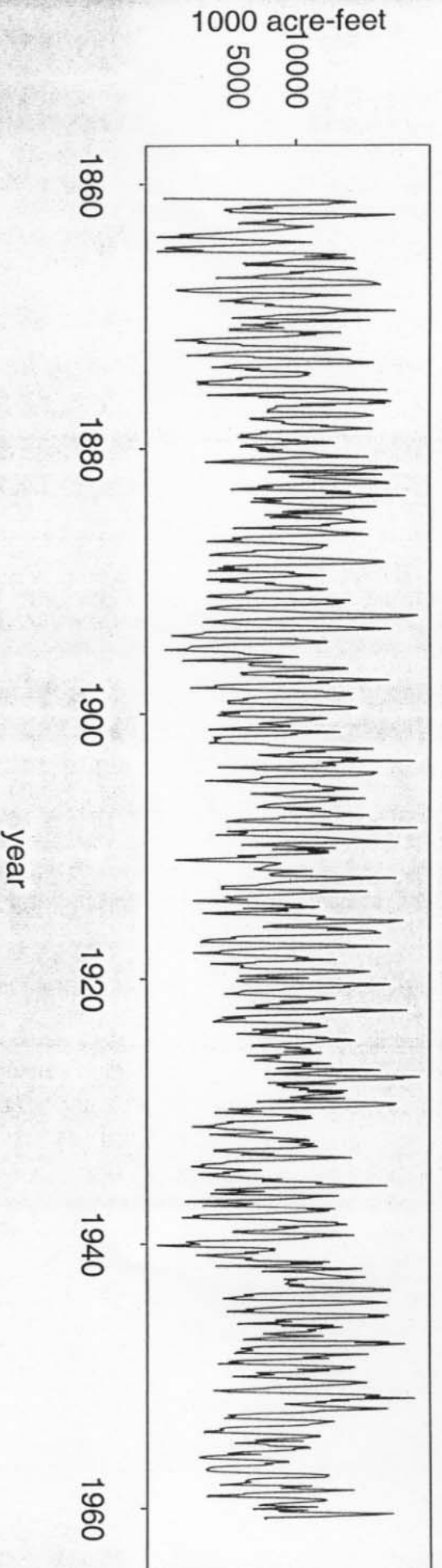
If $|R|^2 = 0$, then

$$E |\hat{R}|^2 \sim \frac{1}{n}$$

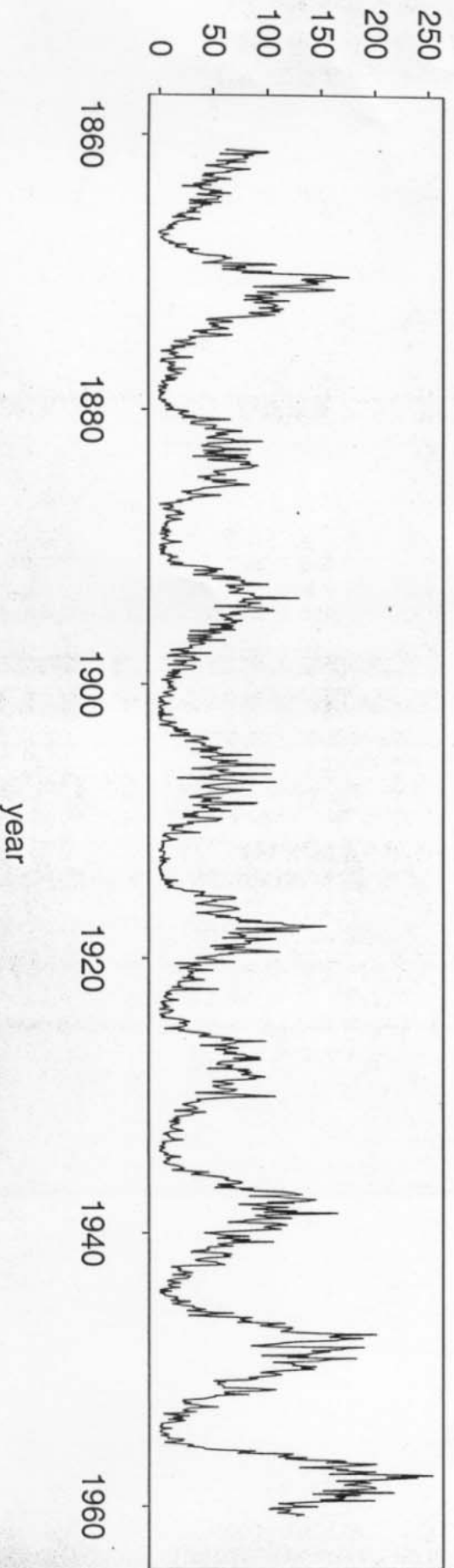
$$100\alpha\% \text{ point} \quad 1 - (1 - \alpha)^{1/(n-1)}$$

Degrees of freedom are $2m$

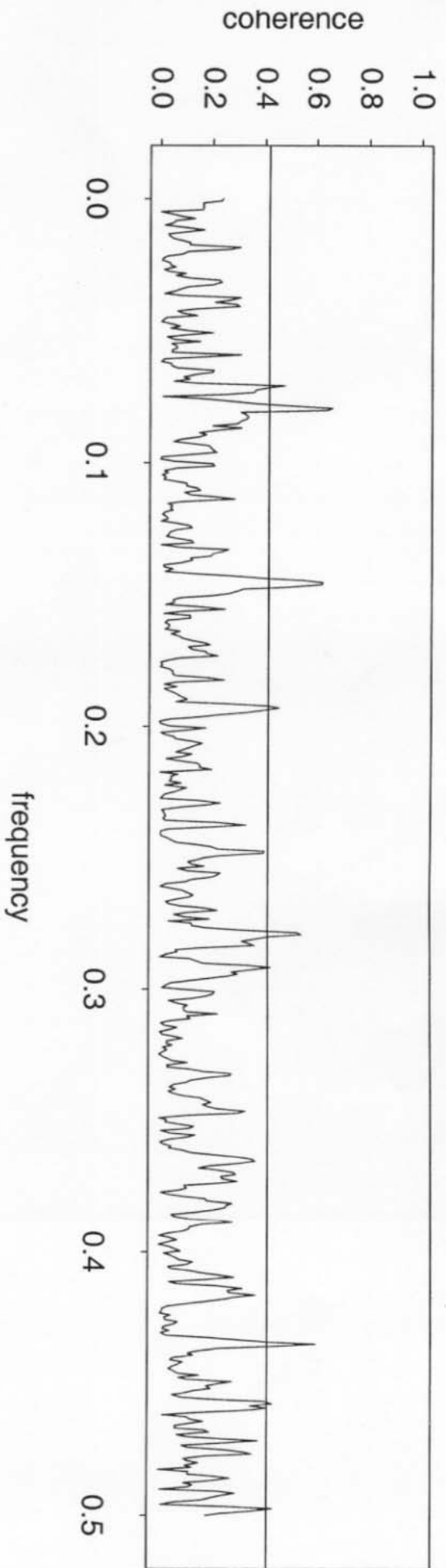
Monthly Mississippi river runoff, 1861-1960



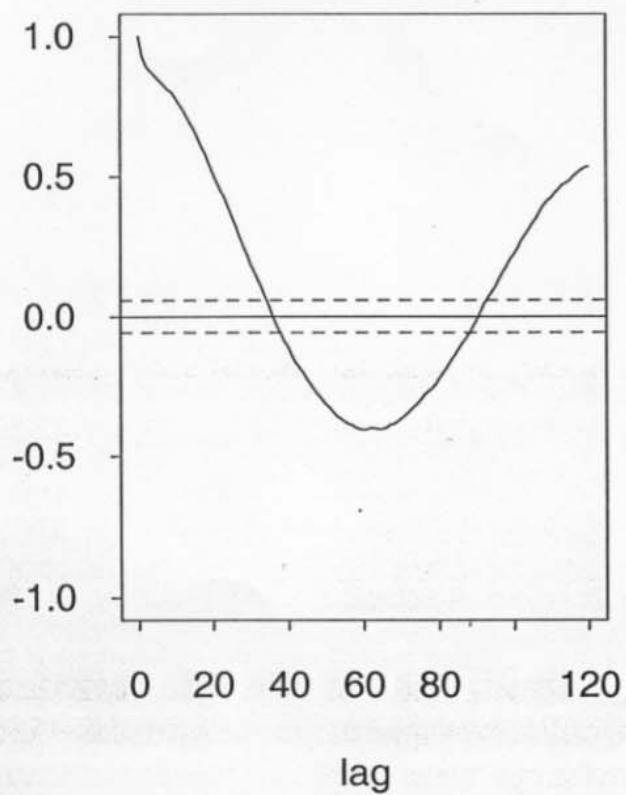
Monthly mean sunspot numbers, 1861-1960



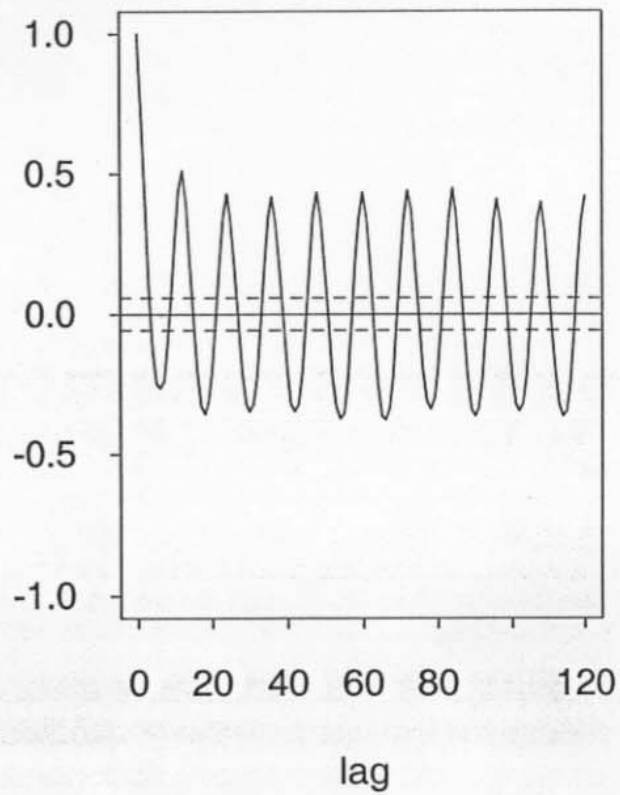
Coherence



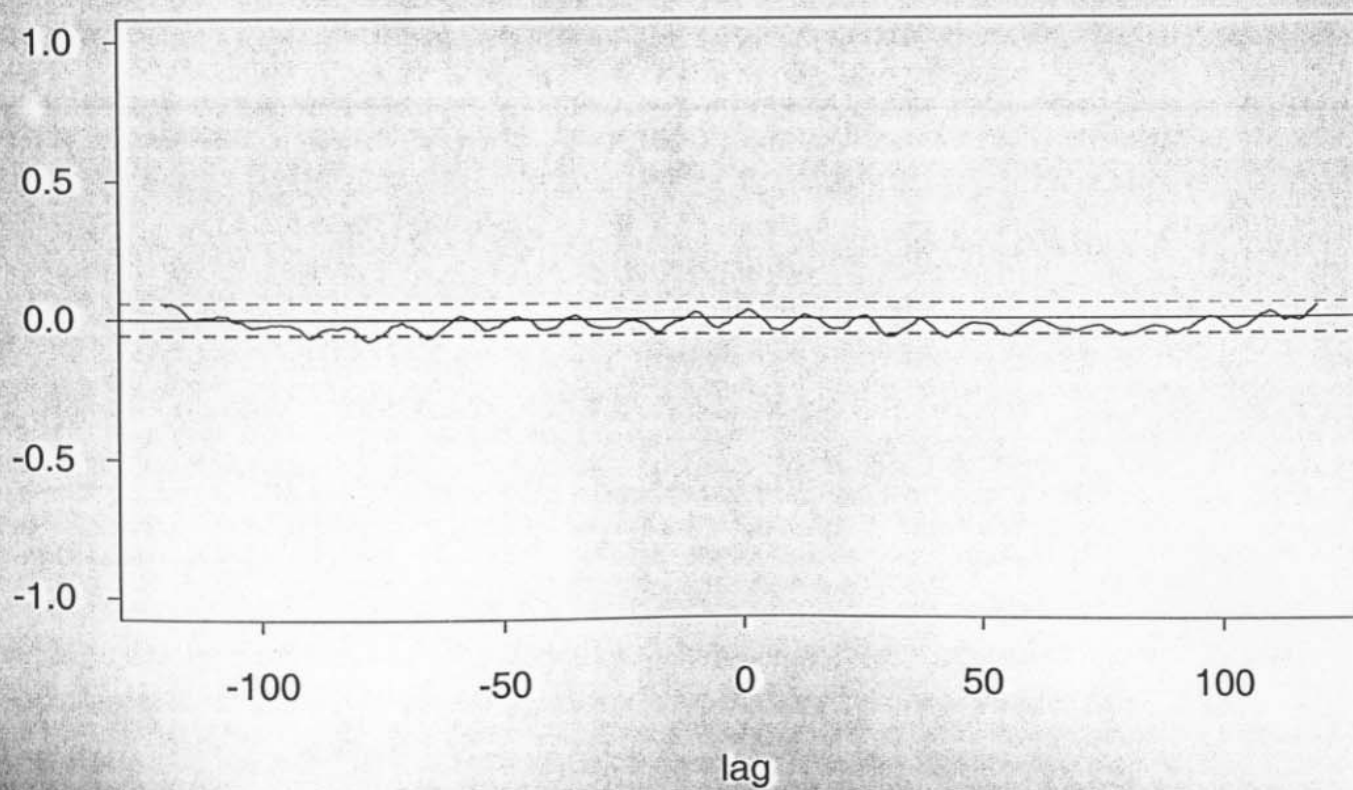
Sunspot autocorrelation



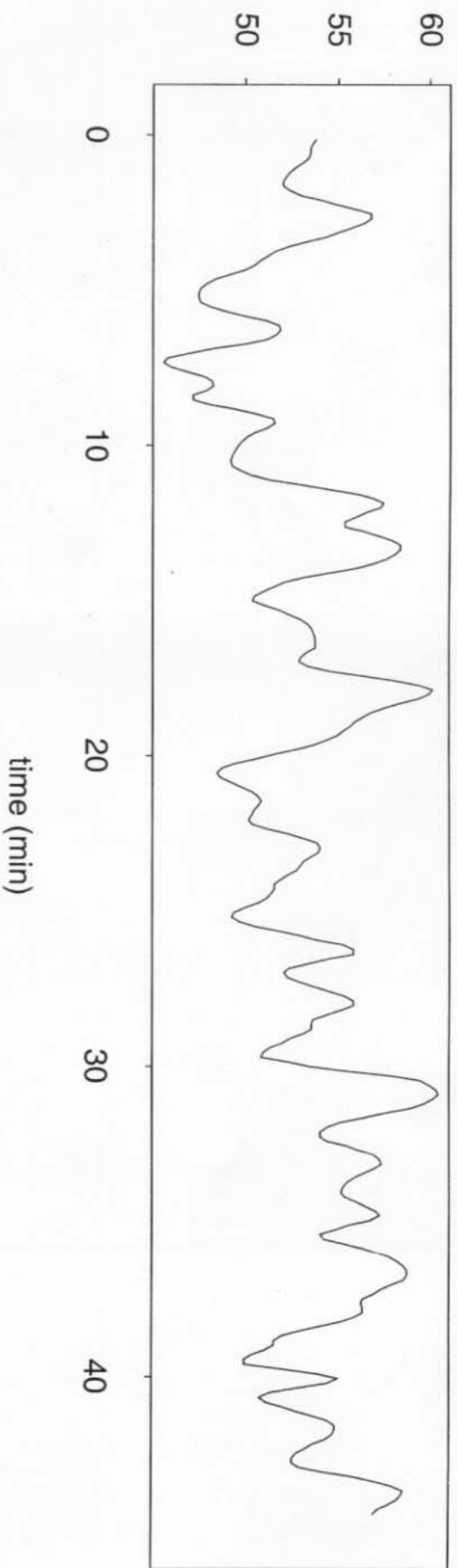
Mississippi autocorrelation



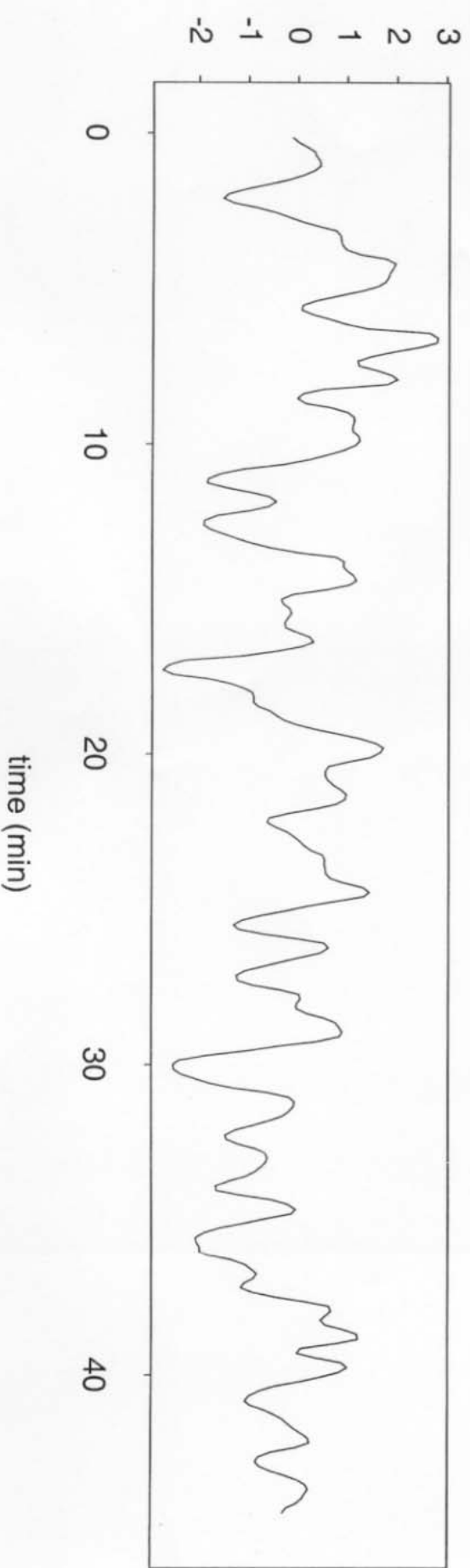
Crosscorrelation



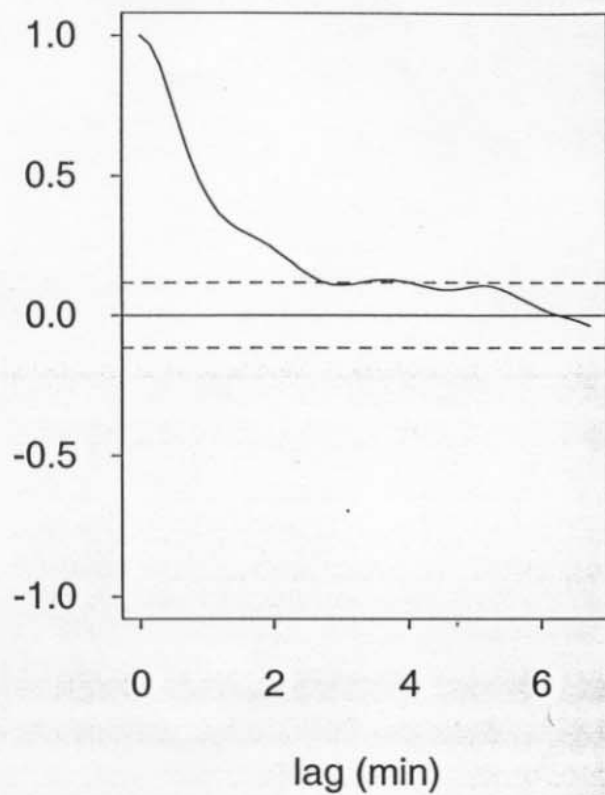
Percent CO₂ in outlet gas



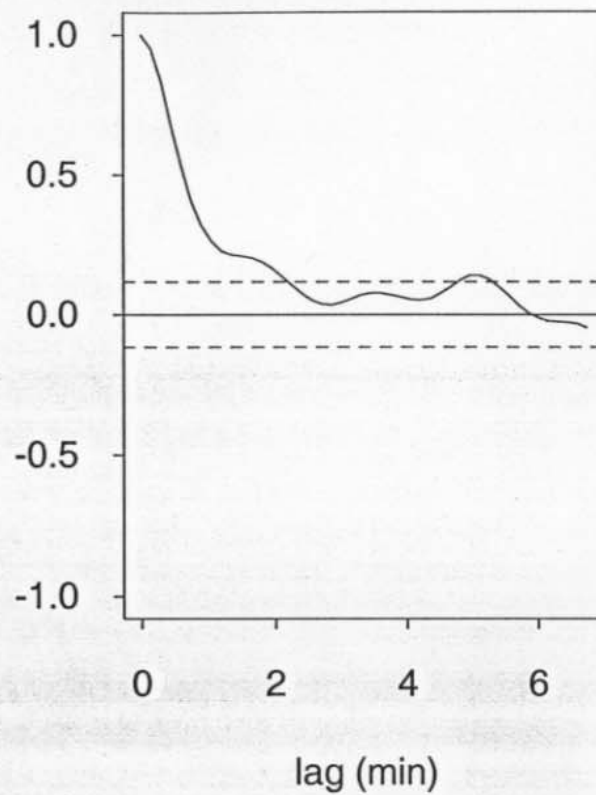
(.6 - methane feed)/.04



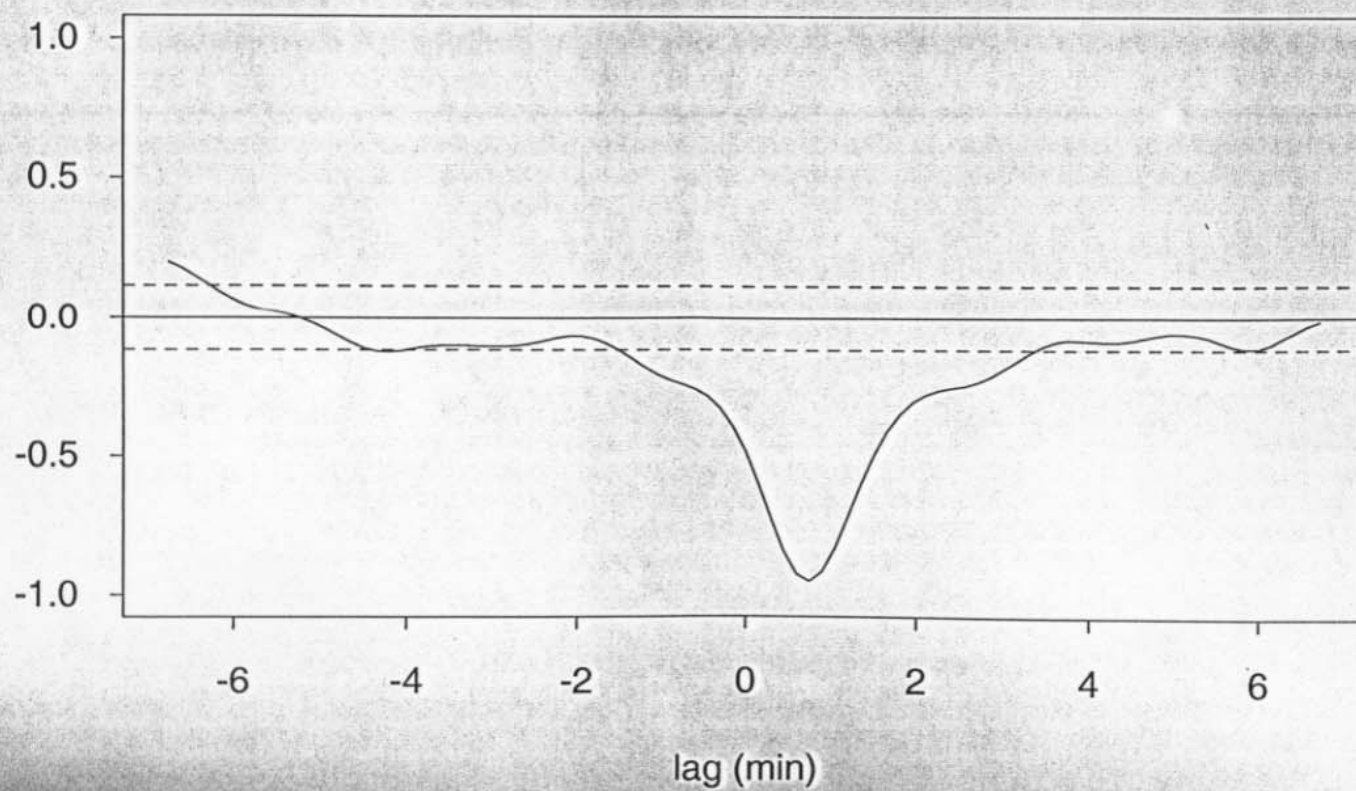
Output autocorrelation



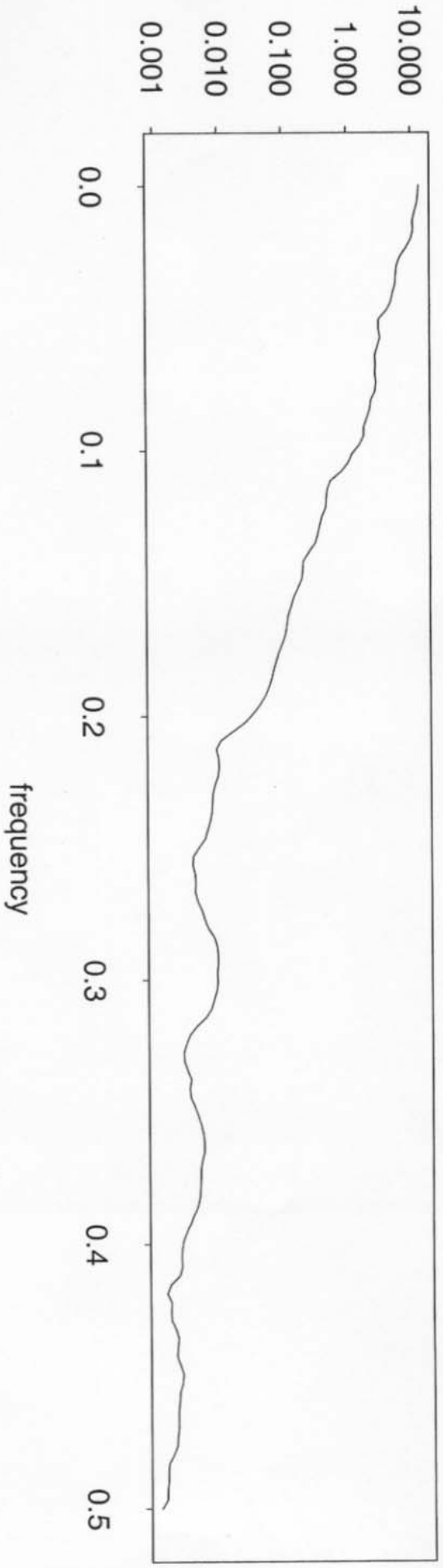
Input autocorrelation



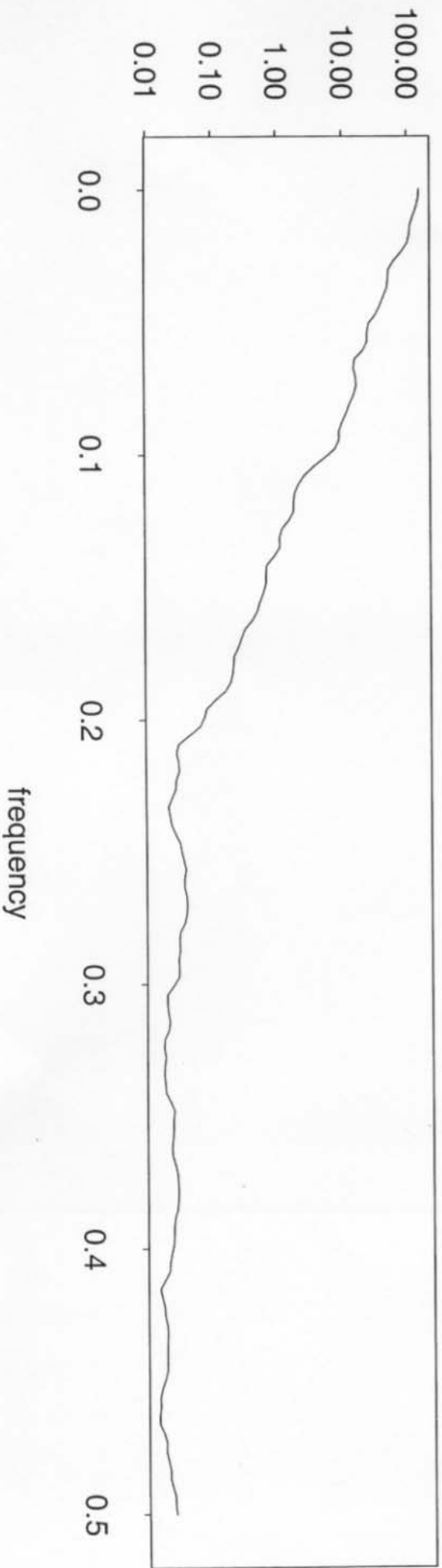
Crosscorrelation



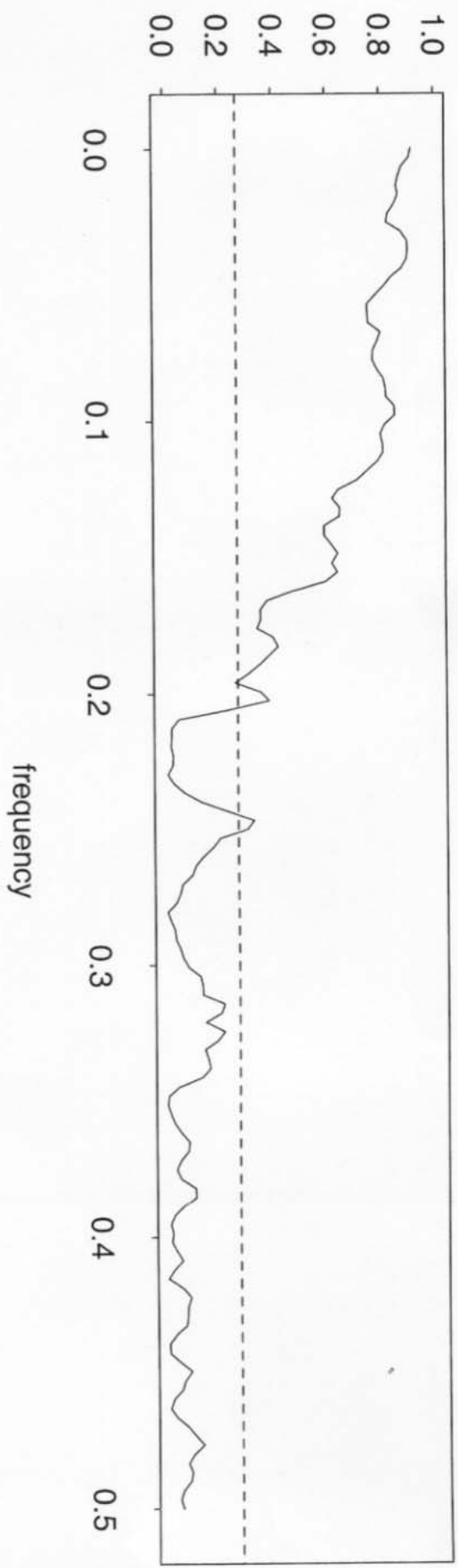
Input spectrum



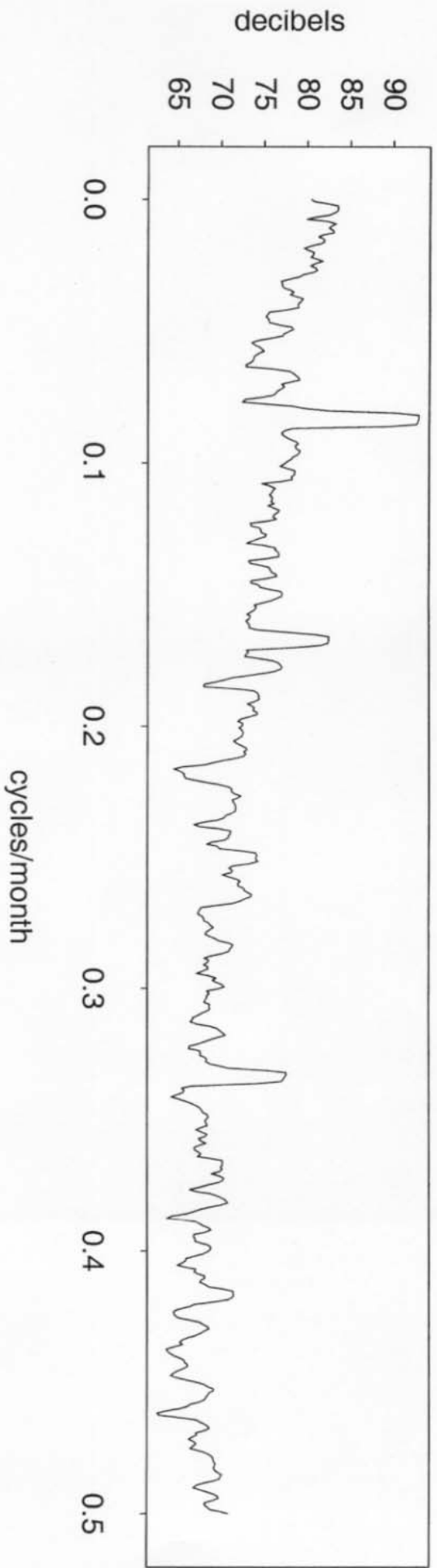
Output spectrum



Coherence



Runoff spectrum



Sunspot spectrum

