Section 7. 1 Nov. 2001

Nonparametric methods of uncertainty estimination

interpenetrating samples,

δ-method, jackknife, bootstrap, cross-validation

Techniques broadly applicable, complex statistics
Generally justified by asymptotics
There exist singular and inappropriate cases
NECESSARY

δ-method AKA method of linearization, propagation of error, Taylor series method

Gauss (1815)

Basically one approximates functions by Taylor expansions (usually linear) of basic random variables.

Rao, Section 6a.2

Sequence of k-dimensional statistics

$$\underset{\sim}{T}_n = (T_{1n}, \dots, T_{kn}) \qquad n = 1, 2, \dots$$

eg. $(\bar{X}, \bar{Y})$

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Derived statistic

$$g(\underset{\sim}{T}_n) = g(T_{1n}, \ldots, T_{kn})$$

eg. $$g(\bar{X}, \bar{Y}) = \frac{\bar{Y}}{\bar{X}}$$

Suppose $\boxed{\sqrt{n}\,(\underset{\sim}{T}_n - \underset{\sim}{\theta}) \xrightarrow{d} N_k(0, \underset{\sim}{\Sigma})}$ $\xrightarrow{L}$

Write

$$g(\underset{\sim}{T}_n) = g(\underset{\sim}{\theta}) + \frac{\partial g(\underset{\sim}{\theta})}{\partial \underset{\sim}{\theta}} \cdot (\underset{\sim}{T}_n - \theta) + \ldots$$

The entity of principal interest is typically $g(\underset{\sim}{\theta})$

Theorem. If $g(\,)$ has a continuous first derivative

$$\boxed{\sqrt{n}\,\{g(\underset{\sim}{T}_n) - g(\theta)\} \xrightarrow{d} N\left(\underset{\sim}{0}, \frac{\partial g}{\partial \theta}^T \underset{\sim}{\Sigma} \frac{\partial g}{\partial \theta}\right)}$$

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Convergence in distribution (law)

Sequence of r.v.'s $\{X_n\}$

$$F_n(x) = \text{Prob}\{X_n \le x\}$$

$X_n \xrightarrow{d} X$ $\quad (X_n \xrightarrow{L} X)$ $\quad$ (weakly)

if $F_n(x) \to F(x)$ at all continuity points of F

equivalent to

$\int g\, dF_n \to \int g\, dF$ for all bounded continuous g

Doesn't always hold if g is unbounded.

e.g. $X_n = \mu + \sigma Z + \frac{1}{n} C$ $\quad$ C: Cauchy

$$X_n \xrightarrow{d} X \sim N(\mu, \sigma^2)$$

$$EX_n = \infty, \quad EX = \mu$$

$\text{var}\, h(X_n) \not\to \text{var}\, h(X)$ generally

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Result suggests estimating the covariance matrix of $g(\underset{\sim}{T}_n)$ by

$$\left(\left.\frac{\partial g}{\partial \theta}\right|_{\theta = T_n}\right)' \hat{\underset{\sim}{\Sigma}} \left(\frac{\partial g}{\partial \theta}\right)$$

but...

Procedure "works" provided $\underset{\sim}{T}_n$ is a neighborhood of θ

Gives an approximating distribution.
Could use for confidence intervals.

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Example where not too useful (Cramér's book)
{ but more later } !

$$r = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum (x_i - \bar{x})^2 \sum (y_i - \bar{y})^2}}$$

$$= \frac{\bar{t}_4 - \bar{t}_1 \bar{t}_2}{\sqrt{(\bar{t}_3 - \bar{t}_1^2)(\bar{t}_4 - \bar{t}_2^2)}} = g(\bar{t}_1, \bar{t}_2)$$

$t_1 = x/\sqrt{n}$ $\quad t_3 = x^2$ $\quad t_5 = y^2$

$t_2 = y/\sqrt{n}$ $\quad t_4 = xy$

Answer. An estimate of var r

$$\frac{r^2}{4n}\left[\frac{\hat{\mu}_{40}}{\hat{\mu}_{20}^2} + \frac{\hat{\mu}_{04}}{\hat{\mu}_{02}^2} + \frac{2\hat{\mu}_{22}}{\hat{\mu}_{20}\hat{\mu}_{02}} + \frac{4\hat{\mu}_{22}}{\hat{\mu}_{11}} - \frac{4\hat{\mu}_{31}}{\hat{\mu}_{11}\hat{\mu}_{20}} - \frac{4\hat{\mu}_{13}}{\hat{\mu}_{11}\hat{\mu}_{02}} \right]$$

But in normal case $\quad (1-r^2)^2/n$

$$\hat{\mu}_{gh} = \frac{1}{n} \sum (x_i - \bar{x})^g (y_i - \bar{y})^h$$

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Remarks

1. Is n large enough?

2. Is g sufficiently linear? (differentiable)

3. Is the algebra correct?

4. Is it approximately normal? (CI's etc)

5. Note this is not giving expected values

6. The derivatives appearing may be approximated by finite differences

$$\frac{\partial g}{\partial \theta_j} = \frac{g(\theta_1, \ldots, \theta_{j-1}, \theta_j + \delta, \theta_{j+1}, \ldots, \theta_p) - g(\theta_1, \ldots, \theta_j - \delta, \ldots \theta_p)}{2\delta}$$

7. While often forgotten "learning" the bias of an estimate can be important.

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Example where useful

Ratio estimate, as in survey sampling

$$\hat{R} = \bar{y}/\bar{x} \qquad (\text{sampling without replacement})$$

$$\hat{R} - R = \frac{\bar{y}}{\bar{x}} - R = (\bar{y} - R\bar{x})(\bar{X} + (\bar{x} - \bar{X}))^{-1}$$

Writing cap letters for population values

$$\hat{R} - R = \frac{\bar{y} - R\bar{x}}{\bar{X}}\left[1 - \frac{\bar{x} - \bar{X}}{\bar{X}} + \left(\frac{\bar{x} - \bar{X}}{\bar{X}}\right)^2 + \dots\right]$$

$$\simeq \frac{\bar{y} - R\bar{x}}{\bar{X}} \qquad \text{linear approx}$$

$$\simeq \frac{\bar{y} - R\bar{x}}{\bar{X}} - \frac{\bar{y} - R\bar{x}}{\bar{X}}\left(\frac{\bar{x} - \bar{X}}{\bar{X}}\right) \qquad \text{quadratic}$$

$$\operatorname{var}\hat{R} \simeq \operatorname{var}\left(\frac{\bar{y} - R\bar{x}}{\bar{X}}\right) = \frac{1-f}{n\bar{X}^2}\sum_{i=1}^{N}(y_i - Rx_i)^2/(N-1)$$

$$f = \frac{n}{N}$$

bias $$E(\hat{R} - R) \simeq E\left(\frac{\bar{y} - R\bar{x}}{\bar{X}}\right) - E\,\frac{(\bar{y} - R\bar{x})(\bar{x} - \bar{X})}{\bar{X}^2}$$

$$\simeq -\frac{1-f}{n\bar{X}^2}\left(\rho S_y S_x - R S_x^2\right)$$

14 April 04

Can use more terms for a (possibly) better approximation:

$$g(T_n) = g(\theta) + g'(\theta)(T_n - \theta) + \frac{1}{2} g''(\theta)(T_n - \theta)^2 + \dots$$

$\overrightarrow{\text{ave}}\, g(T_n)$?

$$E\, T_n = \theta + \frac{b_\theta}{n} + \dots \qquad \text{biased}$$

$$\operatorname{var} T_n = \frac{\sigma^2}{n} + \dots$$

$$E(T_n - \theta)^2 = \left(\frac{b_\theta}{n}\right)^2 + \frac{\sigma^2}{n} + \dots$$

$$\overrightarrow{\text{ave}}\, g(T_n) = g(\theta) + g'(\theta)\frac{b_\theta}{n} + \frac{1}{2} g''(\theta)\frac{\sigma^2}{n} + \dots$$

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Variance stabilizing transformation

Correlation coefficient

$$\sqrt{n}(r-\rho) \xrightarrow{d} N(0, (1-\rho^2)^2)$$

Look for $g()$ such that variance of the large sample distribution is approximately constant

Wish $[g'(\rho)]^2(1-\rho^2)^2 = c$

Take $g'(\rho) = \dfrac{c}{1-\rho^2}$

$$\begin{aligned} g(\rho) &= \int \frac{c}{1-\rho^2}\,d\rho \\ &= \frac{c}{2}\int\left(\frac{1}{1+\rho} + \frac{1}{1-\rho}\right)d\rho \\ &= \frac{c}{2}\left(\log(1+\rho) - \log(1-\rho)\right) \\ &= c \tanh^{-1}\rho \end{aligned}$$

Claims: makes variable more gaussian
relationships more additive

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<u>An oddity</u> r^2 when $\rho = 0$

$$\text{var}\, r^2 \sim [2\rho]^2 \frac{1}{n}(1-\rho^2)^2$$

$$= 0 \quad \text{when } \rho = 0$$

Need more terms in expansion

$$\text{var}\, r^2 \sim \frac{2}{n^2} \quad \text{when } \rho = 0$$

Singular $\left.\frac{\partial g}{\partial \theta}\right|_{\theta_0}$ does occur in practical situations

e.g. estimating MI

Distributions become χ^2, not normal

There are functional forms of these results, e.g. using Frechet or Grateaux derivatives

Consider $\bar{y}/\bar{x}$

Suppose c.d.f. of (x,y) is $F(x,y)$ and empirical c.d.f. is

$$F_n(x,y) = \frac{1}{n}\#\{i \mid x_i \le x, y_i \le y\}$$

Then

$$\theta = \frac{\mu_y}{\mu_x} = \frac{\iint y \, dF(x,y)}{\iint x \, dF(x,y)}$$

and

$$\hat{\theta} = \frac{\bar{y}}{\bar{x}} = \frac{\iint y \, dF_n(x,y)}{\iint x \, dF_n(x,y)}$$

One is considering

$$\theta = t(F) \quad \text{and} \quad \hat{\theta} = t(F_n) \qquad (*)$$

Might use a density estimate f_n instead of F_n

Using $(*)$ defines $\hat{\theta}$, consistently, for all n

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Replicated subsamples
Interpenetrating subsamples (Mahalanobis)

One computes $\hat{\theta}_i = t(\hat{F}_{in})$ for the i-th subsample.

$$\mathcal{S} = \bigcup_i \mathcal{S}_i \ , \quad \mathcal{S}_i \cap \mathcal{S}_{i'} = \emptyset \quad i \neq i'$$

Purposes.
1. To estimate sampling variances when sample design is compicated and exact estimators are unavailable or cumbersome

2. To control field work

3. To measure components of nonsampling variances (e.g. enumerators)
Particularly useful for the study of correlated errors.

$\bar{Y}$ overall sample mean
$\bar{Y}_i$ mean of i-th subsample

Estimate var $\bar{Y}$ by

$$\frac{1}{I} \sum_i (\bar{Y}_i - \bar{Y})^2 / (I-1)$$

Faster

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Anova

	SS	DF	MS
Between subsamples	$\sum_j \sum_i (\bar{Y}_i - \bar{Y})^2$	$I-1$	s_b^2
Within subsamples	$\sum_j \sum_i (Y_{ij} - \bar{Y}_i)^2$	$I(J-1)$	s_w^2
Total	$\sum_j \sum_i (Y_{ij} - \bar{Y})^2$	$IJ-1$	

Perhaps $Y_{ij} = \mu + \tau_i + \epsilon_{ij}$ $\quad \tau_i$ random

τ_i make values in i-th subsample correlated,

$$\text{corr}\{Y_{ij}, Y_{ij'}\} = \frac{\sigma_\tau^2}{\sigma_\epsilon^2 + \sigma_\tau^2}$$

$$E\, s_w^2 = \sigma_\epsilon^2$$

$$E\, s_b^2 = \sigma_\epsilon^2 + J\sigma_\tau^2$$

20 April 04

Advantages of the jackknife.

"Like the Boy Scout's knife, it can be used to do many jobs ..."

Just need a program to evaluate the estimate of interest, $\hat{\theta}$

15 April 04

<u>The jackknife</u>. $n = IJ$, I groups of J

$\hat\theta$ based on all the data

$\hat\theta_{-i}$ " " " ", but the i-th group

$\hat\theta_{pi} = I\hat\theta - (I-1)\hat\theta_{-i}$

$$\operatorname{ave}_i \left(I\hat\theta - (I-1)\hat\theta_{-i} \right) = \theta + \frac{c}{(IJ)^2} + \cdots$$

$\bar\theta$ = <u>ave</u> $\hat\theta_{pi}$ has <u>reduced bias</u> in an asymptotic sense

$$\hat\theta - \bar\theta = -(I-1)\Big[\hat\theta - \sum_i \hat\theta_{-i}/n\Big]$$

estimates the bias

$$s^2 = \sum_i \left(\hat\theta_{pi} - \bar\theta\right)^2 / (I-1)$$

<u>Estimate of var $\hat\theta$</u>, var $\bar\theta$ provided by

$$\frac{s^2}{I}$$

J. W. Tukey, M. Quenouille

15 April 04

E.g. $\hat{\theta} = \dfrac{X_1 + \ldots + X_n}{n} = \bar{X}$

$$= \frac{J\bar{X}_1 + \ldots + J\bar{X}_I}{IJ}$$

$$= \frac{\bar{X}_1 + \ldots + \bar{X}_I}{I}$$

$$\hat{\theta}_{-i} = \frac{J\bar{X}_1 + \ldots + J\bar{X}_I - J\bar{X}_i}{(I-1)J}$$

$$= \frac{\bar{X}_1 + \ldots + \bar{X}_I - \bar{X}_i}{(I-1)J}$$

$\hat{\theta}_{pi} = \bar{X}_i$

$\bar{\theta} = \hat{\theta} = \bar{X}$ here

$\hat{\theta} - \bar{\theta} = 0$

$s^2 = \sum_i (\bar{X}_i - \bar{X})^2 / (I-1)$

var $\bar{X}$ estimated by s^2/I

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Other justifications are asymptotic, e.g. for $\hat{\theta} = g(\bar{A}, \bar{B}, \ldots)$ (function of means)

Which asymptotics?

I fixed, $J \to \infty$ easy

$I \to \infty$, J fixed harder

e.g. $J = 1$

The estimate is inconsistent for the sample median when $J = 1$. (Not a regular enough functional)

Might compute for: histogram, qq plot, ...

There are also weighted jackknives, e.g. for regression

Tukey suggested $I = 10$