### **Introduction to Time Series Analysis. Lecture 25.**

- 1. Lagged regression models.
- 2. Review: lagged regression in the time domain
- 3. Cross spectrum. Coherence.
- 4. Lagged regression in the frequency domain.

# **Lagged regression models**

Consider a lagged regression model of the form

$$Y_t = \sum_{h=-\infty}^{\infty} \beta_h X_{t-h} + V_t,$$

where  $X_t$  is an observed input time series,  $Y_t$  is the observed output time series, and  $V_t$  is a stationary noise process.

This is useful for

- Identifying the (best linear) relationship between two time series.
- Forecasting one time series from the other.

# Lagged regression models: Agenda

- Multiple, jointly stationary time series in the time domain: cross-covariance function, sample CCF.
- Lagged regression in the time domain: model the input series, extract the white time series driving it ('prewhitening'), regress with transformed output series.
- Multiple, jointly stationary time series in the frequency domain: cross spectrum, coherence.
- Lagged regression in the frequency domain: Calculate the input's spectral density, and the cross-spectral density between input and output, and find the transfer function relating them, in the frequency domain. Then the regression coefficients are the inverse Fourier transform of the transfer function.

#### **Review: Cross-covariance**

The cross-covariance function of two jointly stationary processes  $\{X_t\}$  and  $\{Y_t\}$  is

$$\gamma_{xy}(h) = \mathbf{E}\left[ (X_{t+h} - \mu_x)(Y_t - \mu_y) \right].$$

Their cross-correlation function is

$$\rho_{xy}(h) = \frac{\gamma_{xy}(h)}{\sqrt{\gamma_x(0)\gamma_y(0)}}.$$

So  $\rho_{xy}(h) = \rho_{yx}(-h)$ .

**Example:** For  $Y_t = \beta X_{t-\ell} + W_t$  for stationary  $\{X_t\}$ , white uncorr.  $\{W_t\}$ ,

$$\gamma_{xy}(h) = \beta^2 \gamma_x(h+\ell).$$

If  $\ell > 0$ , we say  $X_t$  leads  $Y_t$ .

If  $\ell < 0$ , we say  $X_t$  lags  $Y_t$ .

### **Review: lagged regression in the time domain**

Suppose we wish to fit a lagged regression model of the form

$$Y_t = \alpha(B)X_t + \eta_t = \sum_{j=0}^{\infty} \alpha_j X_{t-j} + \eta_t,$$

where  $X_t$  is an observed input time series,  $Y_t$  is the observed output time series, and  $\eta_t$  is a stationary noise process, uncorrelated with  $X_t$ .

- 1. Fit  $\theta_x(B)$ ,  $\phi_x(B)$  to model the input series  $\{X_t\}$ .
- 2. Prewhiten the input series by applying the inverse operator  $\phi_x(B)/\theta_x(B)$ :

$$\tilde{Y}_t = \frac{\phi_x(B)}{\theta_x(B)} Y_t = \alpha(B) W_t + \frac{\phi_x(B)}{\theta_x(B)} \eta_t.$$

# **Review: Lagged regression in the time domain**

3. Calculate the cross-correlation of  $\tilde{Y}_t$  with  $W_t$ ,

$$\gamma_{\tilde{y},w}(h) = \mathbf{E}\left(\sum_{j=0}^{\infty} \alpha_j W_{t+h-j} W_t\right) = \sigma_w^2 \alpha_h,$$

to give an indication of the behavior of  $\alpha(B)$  (for instance, the delay).

4. Estimate the coefficients of  $\alpha(B)$  and hence fit an ARMA model for the noise series  $\eta_t$ .

#### Coherence

To analyze lagged regression in the frequency domain, we'll need the notion of *coherence*, the analog of cross-correlation in the frequency domain.

Define the cross-spectrum as the Fourier transform of the cross-correlation,

$$f_{xy}(\nu) = \sum_{h=-\infty}^{\infty} \gamma_{xy}(h)e^{-2\pi i\nu h},$$

$$\gamma_{xy}(h) = \int_{-1/2}^{1/2} f_{xy}(\nu) e^{2\pi i \nu h} d\nu,$$

(provided that  $\sum_{h=-\infty}^{\infty} |\gamma_{xy}(h)| < \infty$ ).

Notice that  $f_{xy}(\nu)$  can be complex-valued.

Also, 
$$\gamma_{yx}(h) = \gamma_{xy}(-h)$$
 implies  $f_{yx}(\nu) = f_{xy}(\nu)^*$ .

#### Coherence

The squared coherence function is

$$\rho_{y,x}^{2}(\nu) = \frac{|f_{yx}(\nu)|^{2}}{f_{x}(\nu)f_{y}(\nu)}.$$

Compare this with the correlation  $\rho_{y,x} = \text{Cov}(Y,X)/\sqrt{\sigma_x^2 \sigma_y^2}$ . We can think of the squared coherence at a frequency  $\nu$  as the contribution to squared correlation at that frequency.

(Recall the interpretation of spectral density at a frequency  $\nu$  as the contribution to variance at that frequency.)

## **Estimating squared coherence**

Recall that we estimated the spectral density using the smoothed squared modulus of the DFT of the series,

$$\hat{f}_x(\nu_k) = \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} |X(\nu_k - l/n)|^2$$

$$= \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} X(\nu_k - l/n) X(\nu_k - l/n)^*.$$

We can estimate the cross spectral density using the same sample estimate,

$$\hat{f}_{xy}(\nu_k) = \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} X(\nu_k - l/n) Y(\nu_k - l/n)^*.$$

Coherence

Also, we can estimate the squared coherence using these estimates,

$$\hat{\rho}_{y,x}^{2}(\nu) = \frac{|\hat{f}_{yx}(\nu)|^{2}}{\hat{f}_{x}(\nu)\hat{f}_{y}(\nu)}.$$

(Knowledge of the asymptotic distribution of  $\hat{\rho}_{y,x}^2(\nu)$  under the hypothesis of no coherence,  $\rho_{y,x}(\nu)=0$ , allows us to test for coherence.)

Consider a lagged regression model of the form

$$Y_t = \sum_{j=-\infty}^{\infty} \beta_j X_{t-j} + V_t,$$

where  $X_t$  is an observed input time series,  $Y_t$  is the observed output time series, and  $V_t$  is a stationary noise process.

We'd like to estimate the coefficients  $\beta_j$  that determine the relationship between the lagged values of the input series  $X_t$  and the output series  $Y_t$ .

The projection theorem tells us that the coefficients that minimize the mean squared error,

$$\mathbf{E}\left[\left(Y_t - \sum_{j=-\infty}^{\infty} \beta_j X_{t-j}\right)^2\right]$$

satisfy the orthogonality conditions

$$\mathbf{E}\left[\left(Y_t - \sum_{j=-\infty}^{\infty} \beta_j X_{t-j}\right) X_{t-k}\right] = 0, \qquad k = 0, \pm 1, \pm 2, \dots$$

$$\sum_{j=-\infty}^{\infty} \beta_j \gamma_x (k-j) = \gamma_{yx}(k), \qquad k = 0, \pm 1, \pm 2, \dots$$

We could solve these equations for the  $\beta_j$  using the sample autocovariance and sample cross-covariance. But it is more convenient to use estimates of the spectra and cross-spectrum.

(Convolution with  $\{\beta_j\}$  in the time domain is equivalent to multiplication by the Fourier transform of  $\{\beta_j\}$  in the frequency domain. Let's verify this.)

We replace the autocovariance and cross-covariance with the inverse Fourier transforms of the spectral density and cross-spectral density in the orthogonality conditions,

$$\sum_{j=-\infty}^{\infty} \beta_j \gamma_x(k-j) = \gamma_{yx}(k), \qquad k = 0, \pm 1, \pm 2, \dots$$

This gives, for  $k = 0, \pm 1, \pm 2, \ldots$ ,

$$\int_{-1/2}^{1/2} \sum_{j=-\infty}^{\infty} \beta_j e^{2\pi i \nu (k-j)} f_x(\nu) d\nu = \int_{-1/2}^{1/2} e^{2\pi i \nu k} f_{yx}(\nu),$$

$$\int_{-1/2}^{1/2} e^{2\pi i \nu k} B(\nu) f_x(\nu) d\nu = \int_{-1/2}^{1/2} e^{2\pi i \nu k} f_{yx}(\nu),$$

where  $B(\nu) = \sum_{j=-\infty}^{\infty} e^{-2\pi i\nu j} \beta_j$  is the Fourier transform of the coefficient sequence  $\beta_j$ .

Since the Fourier transform is unique, the orthogonality conditions are equivalent to

$$B(\nu)f_x(\nu) = f_{yx}(\nu).$$

We can write the mean squared error at the solution as follows. (Why?)

$$\mathbf{E}\left[\left(Y_{t} - \sum_{j=-\infty}^{\infty} \beta_{j} X_{t-j}\right) Y_{t}\right] = \gamma_{y}(0) - \sum_{j=-\infty}^{\infty} \beta_{j} \gamma_{xy}(-j)$$

$$= \int_{-1/2}^{1/2} \left(f_{y}(\nu) - B(\nu) f_{xy}(\nu)\right) d\nu$$

$$= \int_{-1/2}^{1/2} f_{y}(\nu) \left(1 - \frac{f_{yx}(\nu) f_{xy}(\nu)}{f_{x}(\nu) f_{y}(\nu)}\right) d\nu$$

$$= \int_{-1/2}^{1/2} f_{y}(\nu) \left(1 - \frac{|f_{yx}(\nu)|^{2}}{f_{x}(\nu) f_{y}(\nu)}\right) d\nu$$

$$= \int_{-1/2}^{1/2} f_{y}(\nu) \left(1 - \rho_{yx}^{2}(\nu)\right) d\nu.$$

$$MSE = \int_{-1/2}^{1/2} f_y(\nu) \left( 1 - \rho_{yx}^2(\nu) \right) d\nu.$$

Thus,  $\rho_{yx}(\nu)^2$  indicates how the component of the variance of  $\{Y_t\}$  at a frequency  $\nu$  is accounted for by  $\{X_t\}$ . Compare this with the corresponding decomposition for random variables:

$$\mathbf{E}(y - \beta x)^2 = \sigma_y^2 (1 - \rho_{xy}^2).$$

We can estimate the  $\beta_j$  in the frequency domain:

$$\hat{B}(\nu_k) = \frac{\hat{f}_{yx}(\nu_k)}{\hat{f}_x(\nu_k)}.$$

We can approximate the inverse Fourier transform of  $\hat{B}(\nu)$ ,

$$\hat{\beta}_j = \int_{-1/2}^{1/2} e^{2\pi i \nu j} \hat{B}(\nu) d\nu$$

via the sum,

$$\hat{\beta}_j = \frac{1}{M} \sum_{k=0}^{M-1} \hat{B}(\nu_k) e^{2\pi i \nu_k j}.$$

This gives a periodic sequence—we might truncate at j = M/2.

#### Here is the approach:

- 1. Estimate the spectral density and cross-spectral density.
- 2. Compute the transfer function  $\hat{B}(\nu)$ .
- 3. Take the inverse Fourier transform to obtain the impulse response function  $\beta_j$ .

It is often useful to consider both representations

$$Y_t = \sum_{j=-\infty}^{\infty} \alpha_j X_{t-j}, \qquad X_t = \sum_{j=-\infty}^{\infty} \beta_j Y_{t-j},$$

since there might be a more parsimonious representation in terms of one than the other. (Just as a small AR model often cannot be well approximated by a small MA model.)

In the  $X_t = SOI/Y_t = Recruitment$  example (Example 4.23), we obtain

$$Y_t = -22X_{t-5} - 15X_{t-6} - 11X_{t-7} - 10X_{t-8} - 7X_{t-9} - \dots + W_t,$$

$$X_t = 0.012Y_{t+4} - 0.018Y_{t+5} + V_t,$$

and the latter is equivalent to

$$(1 - 0.667B)Y_t = -56B^5X_t + U_t.$$

Overview.

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