

Events  $A$  and  $B$  are **independent** if:  
knowing whether  $A$  occurred does not change the probability of  $B$ .

Mathematically, can say in two equivalent ways:

$$P(B|A) = P(B)$$
$$P(A \text{ and } B) = P(B \cap A) = P(B) \times P(A).$$

Important to distinguish independence from **mutually exclusive** which would say  $B \cap A$  is empty (cannot happen).

**Example. Deal 2 cards from deck**

- A first card is Ace
- C second card is Ace

$$P(C|A) = \frac{3}{51}$$
$$P(C) = \frac{4}{52} \text{ (last class).}$$

So  $A$  and  $C$  are **dependent**.

### Example. Throw 2 dice

- A first die lands 1
- B second die shows larger number than first die
- C both dice show same number

$$P(B|A) = \frac{5}{6} \quad P(B) = ? = \frac{15}{36} \text{ by counting}$$

so A and B **dependent**.

$$P(C|A) = \frac{1}{6} \quad P(C) = \frac{6}{36} = \frac{1}{6}$$

so A and C **independent**.

Note 1: here  $B$  and  $C$  are mutually exclusive.

Note 2: writing  $B' =$  "second die shows smaller number than first die"  
we have

$$P(B') = P(B) \text{ by symmetry}$$

$$P(B \cup B') = P(C^c) = 1 - P(C) = \frac{5}{6}$$

giving a "non-counting" argument that  $P(B) = \frac{5}{12}$ .

### Example. Deal 1 card from deck

A card is Ace

S card is Spade

$$P(A) = \frac{4}{52} \quad P(S) = \frac{13}{52} \quad P(A \cap S) = \frac{1}{52}.$$

Here  $P(A \cap S) = P(A) \times P(S)$  so **independent**.

### Conceptual point.

(a) In a fully-specified math model, two events are either dependent or independent; can be checked by calculation.

(b) Often we use independence as an **assumption** in making a model.

For instance we **assume** that different die throws give independent results. Most probability models one encounters in engineering or science have some assumption of “bottom level” independence; but one needs to be careful about which other events within the model are independent.

**(silly) Example.**

**Throw 2 dice. If sum is at least 7 I show you the dice; if not, I don't.**

A: I show you first die lands 1

B: I show you second die lands 1

$$P(A) = \frac{1}{36}, \quad P(B) = \frac{1}{36}, \quad P(A \cap B) = 0$$

so A and B **dependent**.

**Conceptual point.** This illustrates a subtle point: being told by a truthful person that “A happened” is not (for probability/statistics purposes) exactly the same as “knowing A happened”.

[car accident example]

## Systems of components

Will show **logic diagrams**: system works if there is some path left-to-right which passes only through working components.

Assume components work/fail independently,

$$P(C_i \text{ works}) = p_i, \quad P(C_i \text{ fails}) = 1 - p_i.$$

Note in practice the independence assumption is usually unrealistic.

Math question: calculate  $P(\text{system works})$  in terms of the numbers  $p_i$  and the network structure.

**Example: “in series”.**

[picture on board]

$$P(\text{ system works } ) = p_1 p_2 p_3.$$

**Example: “in parallel”.**

[picture on board]

$$P(\text{ system fails } ) = (1 - p_1)(1 - p_2)(1 - p_3).$$

$$P(\text{ system works } ) = 1 - (1 - p_1)(1 - p_2)(1 - p_3).$$

## More complicated example:

[picture on board]

We could write out all 16 combinations; instead let's condition on whether or not  $C_1$  works.

$$\begin{aligned} P(\text{system works}) &= P(\text{system works} | C_1 \text{ works})P(C_1 \text{ works}) \\ &+ P(\text{system works} | C_1 \text{ fails})P(C_1 \text{ fails}) \end{aligned}$$

[continue on board]

**Example:** Deal 4 cards. What is chance we get exactly one Spade?

event	1st	2nd	3rd	4th
$F_1$	S	N	N	N
$F_2$	N	S	N	N
$F_3$				
$F_4$	N	N	N	S

[board: repeated conditioning]

$$P(F_1) = \frac{13}{52} \times \frac{39}{51} \times \frac{38}{50} \times \frac{37}{49}$$

$$P(F_1) = P(F_2) = P(F_3) = P(F_4)$$

$$\begin{aligned} P(\text{exactly one Spade}) &= P(F_1 \text{ or } F_2 \text{ or } F_3 \text{ or } F_4) \\ &= P(F_1) + P(F_2) + P(F_3) + P(F_4) = 4 \times P(F_1) \approx 44\%. \end{aligned}$$



**Example:** Deal 4 cards. What is chance we get one card of each suit?

event	1st	2nd	3rd	4th
$A_1$	C	D	H	S
$A_2$	C	D	S	H
.	.	.	.	.
.	.	.	.	.

$$P(A_1) = \frac{13}{52} \times \frac{13}{51} \times \frac{13}{50} \times \frac{13}{49}$$

$$P(A_1) = P(A_2) = \dots$$

Number of possible orders =  $4 \times 3 \times 2 \times 1 = 24 = 4!$

$$P(\text{one card of each suit}) = 24 \times P(A_1) \approx 10.5\%.$$

**Bayes rule:** updating probabilities as new information is acquired.

**(silly) Example** There are 2 coins:

one is fair:  $P(\text{Heads}) = 1/2$ ; one is biased:  $P(\text{Heads}) = 9/10$

Pick one coin at random. Toss 3 times. Suppose we get 3 Heads. What then is the chance that the coin we picked is the biased coin?

**Abstract set-up:** Partition  $(B_1, B_2, \dots)$  of “alternate possibilities”.

Know **prior** probabilities  $P(B_i)$ .

Then observe some event  $A$  happens (the “new information”) for which we know  $P(A|B_i)$ . We want to calculate the **posterior** probabilities  $P(B_i|A)$ .

**Bayes formula:**

$$P(B_i|A) = \frac{P(A|B_i)P(B_i)}{P(A|B_1)P(B_1) + P(A|B_2)P(B_2) + \dots}$$

[example above on board]