## Lecture 18 : Stopping Times and Martingales

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## **18.1** Stopping Times

Assume that  $\{\mathcal{F}_n\}$  is an increasing set of  $\sigma$ -fields. Recall that a stopping time T is a random variable  $T: \Omega \leftarrow \mathbb{Z}^+$  such that  $\forall n < \infty$ , one of the following equivalent conditions holds -

- 1.  $\{T=n\} \in \mathcal{F}_n$
- 2.  $\{T \leq n\} \in \mathcal{F}_n$
- 3.  $\{T \ge n\} \in \mathcal{F}_n$

Walds Identity - Let  $X_1, X_2, X_3, ...$  be i.i.d. random variables with  $E|X_i| < \infty$  and T be a stopping time of  $\{\mathcal{F}_n\}$  where  $\mathcal{F}_n = \sigma(X_1, X_2, ..., X_n)$  with  $E(T) < \infty$ . Let  $S_n = X_1 + X_2 + ... + X_n$ . Then  $ES_T = EX_1ET$ .

Proof:

$$ES_T = E(\sum_{1}^{T} X_i) = \sum_{1}^{\infty} E(X_i 1_{T \ge i}) = \sum_{1}^{\infty} E(X_i 1_{T > i-1}) = \sum_{1}^{\infty} EX_i E 1_{T > i-1} = EX_1 ET$$

The final step is because  $X_i$  and  $1_{T>i-1}$  are independent  $\Box$ .

While it seems as if common means for the  $X_i s$  suffices instead of identical distributions. However this is not true. The reason for this is that we need  $E|X_i| < \infty$  to hold without which the summation and integral (in the calculation of expected value of  $S_T$ ) cannot be exchanged. The following example illustrates this.

Define  $X_i$  as  $P(X_i = \pm 2^i) = \frac{1}{2}$ . Let  $T = \{\inf n : S_n \ge 1\}$ . Clearly,  $P(T = n) = \frac{1}{2}^n$ ,  $ET = 2 < \infty$  and  $ES_T \ge 1$ . However,  $EX_i = 0$  which clearly violates Walds identity.

We now derive the classic **Gamblers Ruin** formula using Walds identity. The problem is that of a random walk that starts at  $X_0 = a > 0$  and  $X_i$  is a symmetric simple random walk i.e., a probability of  $\frac{1}{2}$  for both 1 and -1. Define  $T = \{\inf n : a + S_n = 0 \text{ or } a + S_n = b\}$ . A practical explanation of this problem is that of a gambler starting with a capital of a. We are interested in the probability that the gambler wins b before going broke. Formally, we want to calculate  $P(a + S_T = b)$ . Now,  $E(a + S_T) = bP(a + S_T = b)$  is also given by  $a + ES_T$  which is a by Walds identity. Hence, the gambler earns b before getting broke with a probability of  $\frac{a}{b}$ .

## 18.2 Martingales

Martingales are defined for a filtration i.e., an increasing sequence of  $\sigma$ -fields,  $\mathcal{F}_n$  (n = 1, 2...). A sequence of random variables  $M_n$  is adapted to this filtration of  $M_n \in \mathcal{F}_n$  (n = 1, 2...). Such a

filtration is a martingale (MG) (w.r.t.  $\mathcal{F}_n$ ) if

- $E|M_n| < \infty$  and
- $E(M_{n+1}|\mathcal{F}_n) = M_n \forall n.$

We define  $M_0 = 0$  for convenience and use this definition unless explicitly mentioning otherwise in the rest of the course. The filtration with finite means is a sub-martingale if  $E(M_{n+1}|\mathcal{F}_n) \geq M_n$ and a super-martingale if  $E(M_{n+1}|\mathcal{F}_n) \leq M_n \forall n$ . Note that in the case of martingales, the second condition implies that  $M_n$  is a filtration whereas this is not true in the case of sub-martingales and super-martingales. Another definition that will be used later is that of a predictable sequence. A predictable sequence of random variables  $M_n$  such that  $M_n \in \mathcal{F}_{n-1}$ .

An example of a MG is  $S_n = \sum_{i=1}^n X_i - nEX_1$  where  $X_i$  is a sequence of i.i.d random variables. This is because

$$E(S_{n+1} - (n+1)EX_1 | \mathcal{F}_n) = E(S_n - (n+1)EX_1 | \mathcal{F}_n) + E(X_{n+1} | \mathcal{F}_n) = S_n - (n+1)EX_1 + EX_{n+1} = M_n + M_n$$

Also notice that if  $X_n = M_n - M_{n-1}$  then  $E(X_n | \mathcal{F}_{n-1}) = 0$ . Similar results can easily be derived for super-martingales and sub-martingales. We will be considering two kinds of results involving MGs. These are optional stopping theorems (maximal inequalities) and convergence theorems.

Martingales and predictable sequences can be used in a natural way in gambling systems. If  $X_n$  is the outcome of the  $n^{th}$  bet and  $H_n$  is the multiplier that the gambler places for this bet his/her earnings on this bet are  $X_n \cdot H_n$ . Since gamblers can place bets at time n based upon the outcomes at times 1...n - 1,  $H_n \in \mathcal{F}_{n-1}$  i.e.,  $H_n$  is predictable and  $S_n$  is a martingale. Denoting the gamblers earnings after n bets as a new variable  $Y_n$  we get, assuming  $X_0 = 0$ ,

$$Y_n = H_1 \cdot X_1 + \dots + H_n \cdot X_n = H_1(S_1 - S_0) + \dots + H_n(S_n - S_{n-1}) = (H \cdot S)_n$$
(18.1)

which is the MG-transform.  $Y_n$  is an  $\mathcal{F}_n$ -martingale if  $E(Y_n - Y_{n-1}|\mathcal{F}_n) = 0$ . But,

$$E(Y_n - Y_{n-1}|\mathcal{F}_n) = E(H_n X_n |\mathcal{F}_n) = H_n E(X_n |\mathcal{F}_n)$$
(18.2)

The last equality holds only if  $X_n$  and  $X_nH_n$  are integrable. Since we assume that  $X_n$  is integrable,  $X_nH_n$  is integrable if either  $H_n$  is bounded or  $X_n$ ,  $H_n$  are in  $\mathcal{L}^2$  (Cauchy Shwartz implies that  $X_nH_n$ is integrable in this case).

Martingales in conjunction with stopping times are a neat way of modeling gamblers strategies. A martingale  $M_n$  whose differences represent the outcomes at time n may be bet upon by a gambler until he stops at some time. This time is intuitively a stopping time T because (by definition of stopping times) T = n is measurable w.r.t  $\mathcal{F}_n$ . We can thus define a new process  $M_{n\wedge T}$ .

**Theorem 18.1** If  $M_n$  is an  $\mathcal{F}_n$ -martingale and T is a stopping time, then  $M_{n \wedge T}$  is also an  $\mathcal{F}_n$ -martingale.

**Proof:** Using  $H_n = 1_{T>n-1} \in F_{n-1}$  in the MG-transform formula the result is achieved (note that H is bounded).

Now, if  $M_n$  is a martingale and T is a stopping time bounded by b, then

$$E(M_T) = E(M_{T \wedge b}) = E(M_0)$$
(18.3)

In the case of an unbounded stopping time T, we have that  $M_{T \wedge n} \to M_T$  a.s. Hence, if expectations and limits can be swapped as in

$$E(lim M_{T \wedge n}) = lim E(M_{T \wedge n}) = E(M_0)$$
(18.4)

we can calculate the L.H.S. But, this is not possible always. For instance, in case of a random symmetric walk starting at  $S_0 = 1$  and  $T = (inf \ n : S_n = 0)$ , we have  $ES_T = 0$  because  $P(T < \infty) = 1$ . However,  $1 = ES_0 = ES_{T \wedge n}!$  As we will see later, uniform integrability is enough to justify the swapping of the expectations and limits.