

formula (6.12) implies

$$\begin{aligned} E(S) &= \int_U \int \frac{y}{\|x\|^2} \lambda(x) p(y) dy dx \\ &= \int y p(y) dy \int_U \frac{1}{\|x\|^2} \lambda(x) dx. \end{aligned}$$

Given this naive physical model and a constant $\lambda(x)$, passage to spherical coordinates shows that it is possible for the three-dimensional integral $\int_U \|x\|^{-2} \lambda(x) dx$ to diverge on an unbounded set such as $U = \mathbb{R}^3$. The fact that we are not blinded by light on a starlit night suggests that U is bounded. ■

6.10 Problems

1. Consider a Poisson process in the plane with constant intensity λ . Find the distribution and density function of the distance from the origin of the plane to the nearest random point. What is the mean of this distribution? (Hint: Using Example 6.7.1, you should calculate the distribution function $1 - e^{-\lambda \pi r^2}$.)
2. In the context of Example 6.4.2, show that the j th order statistic $X_{(j)}$ had mean and variance

$$\begin{aligned} E(X_{(j)}) &= \sum_{k=1}^j \frac{1}{\lambda(n-k+1)} \\ \text{Var}(X_{(j)}) &= \sum_{k=1}^j \frac{1}{\lambda^2(n-k+1)^2}. \end{aligned}$$

3. Continuing Problem 2, prove that $X_{(j)}$ has distribution and density functions

$$\begin{aligned} F_{(j)}(x) &= \sum_{k=j}^n \binom{n}{k} (1 - e^{-\lambda x})^k e^{-(n-k)\lambda x} \\ f_{(j)}(x) &= n \binom{n-1}{j-1} (1 - e^{-\lambda x})^{j-1} e^{-(n-j)\lambda x} \lambda e^{-\lambda x}. \end{aligned}$$

(Hint: Ignore the representation of Example 6.4.2 and reason directly.)

4. In the context of Example 6.4.2, suppose you observe $X_{(1)}, \dots, X_{(r)}$ and wish to estimate λ^{-1} by a linear combination $S = \sum_{i=1}^r \alpha_i X_{(i)}$. Demonstrate that $\text{Var}(S)$ is minimized subject to $E(S) = \lambda^{-1}$ by taking $\alpha_j = 1/r$ for $1 < j < r$ and $\alpha_r = (n-r+1)/r$ [49]

by counting all possible successful sequences of births that lead to either the daughter quota or the son quota being fulfilled first. Combining equations (6.16) and (6.17) permits us to write

$$E(N_{sd}) = \frac{s}{p} + \frac{d}{q} - d \sum_{k=0}^{s-1} \binom{d+k}{k} p^k q^d - s \sum_{l=0}^{d-1} \binom{s+l}{l} p^s q^l,$$

replacing a double sum with two single sums.

12. Suppose you randomly drop n balls into m boxes. Assume that a ball is equally likely to land in any box. Use Schrödinger's method to prove that each box receives an even number of balls with probability

$$e_n = \frac{1}{2^m} \sum_{j=0}^m \binom{m}{j} \left(1 - \frac{2j}{m}\right)^n$$

and an odd number of balls with probability

$$o_n = \frac{1}{2^m} \sum_{j=0}^m \binom{m}{j} (-1)^j \left(1 - \frac{2j}{m}\right)^n.$$

(Hint: The even terms of e^t sum to $\frac{1}{2}(e^t + e^{-t})$ and the odd terms to $\frac{1}{2}(e^t - e^{-t})$.)

13. Continuing Problem 12, show that the probability that exactly j boxes are empty is

$$\binom{m}{j} \sum_{k=1}^{m-j} \binom{m-j}{k} (-1)^{m-j-k} \left(\frac{k}{m}\right)^n.$$

14. A one-way highway extends from 0 to ∞ . Cars enter at position 0 at times s determined by a Poisson process on $[0, t]$ with constant intensity λ . Each car is independently assigned a velocity v from a density $g(v)$ on $[0, \infty)$. Demonstrate that the number of cars located in the interval (a, b) at time t has a Poisson distribution with mean $\lambda \int_0^t [G(\frac{b}{t-s}) - G(\frac{a}{t-s})] ds$, where $G(v)$ is the distribution function of $g(v)$ [130].

15. Suppose we generate random circles in the plane by taking their centers (x, y) to be the random points of a Poisson process of constant intensity λ . Each center we independently mark with a radius r sampled from a probability density $g(r)$ on $[0, \infty)$. If we map each random triple (X, Y, R) to the point $U = \sqrt{X^2 + Y^2} - R$, then show that the random points so generated constitute a Poisson process with intensity

$$\eta(u) = 2\pi\lambda \int_0^\infty (r+u)_+ g(r) dr.$$

15. Conclude from this analysis that the number of random circles that overlap the origin is Poisson with mean $\lambda\pi \int_0^\infty r^2 g(r) dr$ [143].

16. Continuing Problem 15, perform the same analysis in three dimensions for spheres. Conclude that the number of random spheres that overlap the origin is Poisson with mean $\frac{4\lambda\pi}{3} \int_0^\infty r^3 g(r) dr$ [143].
17. If $f(x)$ be a simple function and Π is a Poisson process with intensity function $\lambda(x)$, then demonstrate the formulas in equation (6.13) for the characteristic function and generating function of the random sum S .
18. A train departs at time $t > 0$. During the interval $[0, t]$, passengers arrive at the depot at times T determined by a Poisson process with constant intensity λ . The total waiting time passengers spend at the depot is $W = \sum_T (t - T)$. Show that W has mean $E(W) = \frac{\lambda t^2}{2}$ and variance $\text{Var}(W) = \frac{\lambda t^3}{3}$ by invoking Campbell's formulas (6.12) and (6.15) [130].
19. Claims arrive at an insurance company at the times T of a Poisson process with constant intensity λ on $[0, \infty)$. Each time a claim arrives, the company pays S dollars, where S is independently drawn from a probability density $g(s)$ on $[0, \infty)$. Because of inflation and the ability of the company to invest premiums, the longer a claim is delayed, the less it costs the company. If a claim is discounted at rate β , then show that the company's ultimate liability $L = \sum_T S e^{-\beta T}$ has mean and variance

$$\begin{aligned} E(L) &= \frac{\lambda}{\beta} \int_0^\infty s g(s) ds \\ \text{Var}(L) &= \frac{\lambda}{2\beta} \int_0^\infty s^2 g(s) ds. \end{aligned}$$

(Hints: The random pairs (T, S) constitute a marked Poisson process. Use Campbell's formulas (6.12) and (6.15).)