

3.7 Exercises

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1. For $\mathcal{F} = \{\phi, \Omega\}$, show that $E[X|\mathcal{F}] = E[X]$.
2. Give the proof of Proposition 3.2 when X and Y are jointly continuous.
3. If $E[|X_i|] < \infty$, $i = 1, \dots, n$, show that

$$E\left[\sum_{i=1}^n X_i | \mathcal{F}\right] = \sum_{i=1}^n E[X_i | \mathcal{F}]$$

4. Prove that if f is a convex function, then

$$E[f(X)|\mathcal{F}] \geq f(E[X|\mathcal{F}])$$

provided the expectations exist.

5. Let X_1, X_2, \dots , be independent random variables with mean 1. Show that $Z_n = \prod_{i=1}^n X_i$, $n \geq 1$, is a martingale.

6. If $E[X_{n+1}|X_1, \dots, X_n] = a_n X_n + b_n$ for constants $a_n, b_n, n \geq 0$, find constants A_n, B_n so that $Z_n = A_n X_n + B_n, n \geq 0$, is a martingale with respect to the filtration $\sigma(X_0, \dots, X_n)$.

7. Consider a population of individuals as it evolves over time, and suppose that, independent of what occurred in prior generations, each individual in generation n independently has j offspring with probability $p_j, j \geq 0$. The offspring of individuals of generation n then make up generation $n+1$. Assume that $m = \sum_j j p_j < \infty$. Let X_n denote the number of individuals in generation n , and define a martingale related to $X_n, n \geq 0$. The process $X_n, n \geq 0$ is called a branching process.

8. Suppose X_1, X_2, \dots , are independent and identically distributed random variables with mean zero and finite variance σ^2 . If T is a stopping time with finite mean, show that

$$\text{Var}\left(\sum_{i=1}^T X_i\right) = \sigma^2 E(T).$$

9. Suppose X_1, X_2, \dots , are independent and identically distributed mean 0 random variables which each take value +1 with probability $1/2$ and take value -1 with probability $1/2$. Let $S_n = \sum_{i=1}^n X_i$. Which of the following are stopping times? Compute the mean for the ones that are stopping times.
- (a) $T_1 = \min\{i \geq 5 : S_i = S_{i-5} + 5\}$
 (b) $T_2 = T_1 - 5$
 (c) $T_3 = T_2 + 10$.
10. Consider a sequence of independent flips of a coin, and let P_h denote the probability of a head on any toss. Let A be the hypothesis that $P_h = a$ and let B be the hypothesis that $P_h = b$, for given values $0 < a, b < 1$. Let X_i be the outcome of flip i , and set

$$Z_n = \frac{P(X_1, \dots, X_n | A)}{P(X_1, \dots, X_n | B)}$$

If $P_h = b$, show that $Z_n, n \geq 1$, is a martingale having mean 1.

11. Let $Z_n, n \geq 0$ be a martingale with $Z_0 = 0$. Show that

$$E[Z_n^2] = \sum_{i=1}^n E[(Z_i - Z_{i-1})^2]$$

12. Consider an individual who at each stage, independently of past movements, moves to the right with probability p or to the left with probability $1 - p$. Assuming that $p > 1/2$ find the expected number of stages it takes the person to move i positions to the right from where she started.
13. In Example 3.19 obtain bounds on p when $\theta < 0$.
14. Use Wald's equation to approximate the expected time it takes a random walk to either become as large as a or as small as $-b$, for positive a and b . Give the exact expression if a and b are integers, and at each stage the random walk either moves up 1 with probability p or moves down 1 with probability $1 - p$.
15. Consider a branching process that starts with a single individual. Let π denote the probability this process eventually dies out. With X_n denoting the number of individuals in generation n , argue that $\pi^{X_n}, n \geq 0$, is a martingale.

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16. Given X_1, X_2, \dots , Suppose for all i, j , $E[X_i X_j] = 0$ if $i \neq j$.
17. Suppose n random variables X_1, \dots, X_n are each equal to 1 or 0 with probability $1/2$. Let $S_n = \sum_{i=1}^n X_i$. Compute the probability that $S_n = k$ for $k = 0, 1, \dots, n$.
18. Let X_1, X_2, \dots be independent discrete random variables with bounds on the support: $X_1 \in [0, 1], X_2 \in [0, 2], \dots, X_n \in [0, n]$. Let $S_n = \sum_{i=1}^n X_i$. Compute the probability that $S_n = k$ for $k = 0, 1, \dots, n$.
19. Repeat Example 3.19 for a sequence of independent but not identically distributed random variables.
20. Let $Z_n, n \geq 0$ be a martingale with $Z_0 = 0$. Let $v_j, j \geq 0$, be a sequence of nonnegative numbers. Prove the Kolmogorov inequality: $P(\max_{0 \leq k \leq n} |Z_k| \leq v_j) \geq \frac{v_j^2}{E[Z_n^2]}$.
21. Consider a gambler who plays a game at a casino. The gambler starts with 1 unit of money. At each stage, the gambler either wins 1 unit with probability p or loses 1 unit with probability $1 - p$. The gambler stops playing when he either reaches a units or loses all his money. Compute the probability that the gambler reaches a units.
22. What is the probability that a gambler who starts with 1 unit of money and plays a game with $p > 1/2$ will reach a units before reaching $-b$ units?
23. Three gamblers play a game. In each round, each gambler either wins 1 unit with probability p or loses 1 unit with probability $1 - p$. The game ends when all three gamblers have lost all their money. Let X_n, Y_n, Z_n be the number of units each gambler has at time n . Let $S_n = X_n + Y_n + Z_n$. Compute the probability that the game ends at time n .

Problem

a subtle bound on the traveling salesman problem in 10.5.1, we take Z to be 5.2. Consider the integrand $v_j - X_{i-1}$. If S denotes the sum as in Example 5.7.2

then the sum $\sum_{i=1}^n c_i^2$ figuring in Proposition 10.5.1 can be bounded by

$$\sum_{i=1}^n c_i^2 \leq (2\sqrt{2})^2 + 4^2 \sum_{i=1}^{n-1} \frac{1}{n-i} \leq 8 + 16(\ln n + 1).$$

This in turn translates into the Azuma-Hoeffding bound

$$\Pr[|D_n - E(D_n)| \geq \lambda] \leq 2e^{-\frac{\lambda^2}{48+32 \ln n}}.$$

$$-2 \min_{j>i} \|v_j - y_i\|$$

$$-2 \min_{j>i} \|v_j - v_i\|$$

$\{v_i\}$. It follows that

$$2 \min_{j>i} \|v_j - v_i\|$$

$$E(\min_{j>i} \|Y_j - Y_i\|).$$

$(\min_{j>i} \|Y_j - y\| \geq r)$ for the case at a distance of r or less extreme case occurs when y result,

$$\leq e^{-\frac{(n-i)\pi r^2}{4}}.$$

$$\frac{(n-i)\pi r^2}{4} dr$$

$$-\frac{(n-i)\pi r^2}{4} dr$$

$$\frac{1}{i}.$$

we use the crude inequality

$$\leq 2\sqrt{2},$$

10.6 Problems

1. Define the random variables Y_n inductively taking by $Y_0 = 1$ and Y_{n+1} to be uniformly distributed on the interval $(0, Y_n)$. Show that the sequence $X_n = 2^n Y_n$ is a martingale.
2. An urn contains b black balls and w white balls. Each time we randomly withdraw a ball, we replace it by $c + 1$ balls of the same color. Let X_n be the fraction of white balls after n draws. Demonstrate that X_n is a martingale.
3. Let Y_1, Y_2, \dots be a sequence of independent random variables with zero means and common variance σ^2 . If $X_n = Y_1 + \dots + Y_n$, then show that $X_n^2 - n\sigma^2$ is a martingale.
4. Let Y_1, Y_2, \dots be a sequence of i.i.d. random variables with common moment generating function $M(t) = E(e^{tY_1})$. Prove that

$$X_n = M(t)^{-n} e^{t(Y_1 + \dots + Y_n)}$$

is a martingale whenever $M(t) < \infty$.

5. Let Y_n be a finite-state, discrete-time Markov chain with transition matrix $P = (p_{ij})$. If v is a column eigenvector for P with nonzero eigenvalue λ , then verify that $X_n = \lambda^{-n} v_{Y_n}$ is a martingale, where v_{Y_n} is coordinate Y_n of v .
6. Suppose Y_n is the number of particles at the n th generation of a branching process. If s_∞ is the extinction probability, prove that $X_n = s_\infty^{Y_n}$ is a martingale. (Hint: If $Q(s)$ is the progeny generating function, then $Q(s_\infty) = s_\infty$.)
7. In Example 10.3.2, show that $\text{Var}(X_\infty) = \frac{\sigma^2}{\mu(\mu-1)}$ by differentiating equation (10.12) twice. This result is consistent with the mean square convergence displayed in equation (10.9).

8. In Example 10.3.2, show that the fractional linear transformation

$$L_{\infty}(t) = \frac{pt - p + q}{qt - p + q}$$

solves equation (10.12) when $Q(s) = \frac{p}{1-qs}$ and $\mu = \frac{q}{p}$. Also verify equation (10.13).

9. In Example 10.2.2, suppose that each Y_n is equally likely to assume the values $\frac{1}{2}$ and $\frac{3}{2}$. Show that $\prod_{i=1}^{\infty} Y_i \equiv 0$, but $\prod_{i=1}^{\infty} E(Y_i) = 1$ [19]. (Hint: Apply the strong law of large numbers to the sequence $\ln Y_n$.)
10. Given $X_0 = \mu \in (0, 1)$, define X_n inductively by

$$X_{n+1} = \begin{cases} \alpha + \beta X_n, & \text{with probability } X_n \\ \beta X_n, & \text{with probability } 1 - X_n. \end{cases}$$

where $\alpha, \beta > 0$ and $\alpha + \beta = 1$. Prove that X_n is a martingale with (a) $X_n \in (0, 1)$, (b) $E(X_n) = \mu$, and (c) $\text{Var}(X_n) = [1 - (1 - \alpha^2)^n] \mu(1 - \mu)$. Also prove that Proposition 10.3.2 implies that $\lim_{n \rightarrow \infty} X_n = X_{\infty}$ exists with $E(X_{\infty}) = \mu$ and $\text{Var}(X_{\infty}) = \mu(1 - \mu)$. (Hint: Derive a recurrence relation for $\text{Var}(X_{n+1})$ by conditioning on X_n .)

11. Let $S_n = X_1 + \cdots + X_n$ be a symmetric random walk on the integers $\{-a, \dots, b\}$ starting at $S_0 = 0$. For the stopping time

$$T = \min\{n : S_n = -a \text{ or } S_n = b\},$$

prove that $\Pr(S_T = b) = a/(a + b)$ by considering the martingale S_n and that $E(T) = ab$ by considering the martingale $S_n^2 - n$. (Hints: Apply Proposition 10.4.1 and Problem 3.)

12. In the Wright-Fisher model of Example 10.2.6, show that

$$Z_n = \frac{X_n(1 - X_n)}{\left(1 - \frac{1}{2m}\right)^n}$$

is a martingale with values on $[0, 1]$. In view of Proposition 10.3.2, $\lim_{n \rightarrow \infty} Z_n = Z_{\infty}$ exists. Thus, $X_n(1 - X_n) \approx \left(1 - \frac{1}{2m}\right)^n Z_{\infty}$ for n large. In other words, X_n approaches either 0 or 1 at rate $1 - \frac{1}{2m}$.

13. In Proposition 10.4.1 suppose we can write either

$$X_n = B_n + I_n \quad \text{or} \quad X_n = B_n - I_n$$

for $n \leq T$, where $|B_n| \leq c$ and $0 \leq I_{n-1} \leq I_n$ when $n \leq T$. In other words for all times up to T , the B process is bounded and the I process is increasing. Show that $E(X_T) = \mu$ holds without making assumptions (b) and (c) of the proposition. (Hints: Show that $E(X_{T \wedge n}) = \mu$ for $T \wedge n = \min\{T, n\}$. Apply the bounded convergence theorem to $B_{T \wedge n}$ and the monotone convergence theorem to $I_{T \wedge n}$.)

14. Let Y_1, \dots, Y_n be independent random variables with success probability μ (10.22) to Chebyshev's

P

when $\mu = 1/2$. Which is more likely than the other, to favor Chebyshev's?

15. Suppose that v_1, Y_1, \dots, Y_n be independent random variables with the two-point set that