

# Outline

11/2/2023

- 1) Convergence in Probability and Distribution
- 2) Continuous Mapping, Slutsky's Theorem
- 3) Delta method

## Asymptotics

[ So far, everything has been finite-sample, often using special properties of model  $\mathcal{P}$  (e.g. exp. fam.) to do exact calculations. ]

[ For "generic" models, exact calculations may be intractable or impossible. But we may be able to approximate our problem with a simpler problem in which calculations are easy ]

[ Typically approximate by Gaussian, by taking limit as # observations  $\rightarrow \infty$ . But this is only interesting if approx. is good for "reasonable" sample size. ]

# Convergence

Let  $X_1, X_2, \dots \in \mathbb{R}^d$  sequence of random vectors

We care about 2 kinds of convergence:

1) cvg. in probability ( $X_n \approx \text{constant}$ )

2) cvg. in distribution ( $X_n \approx N_d(0, I_d)$ , usually)

We say the sequence converges in probability to  $c \in \mathbb{R}^d$  ( $X_n \xrightarrow{P} c$ ) if

$$\mathbb{P}(\|X_n - c\| > \varepsilon) \rightarrow 0, \quad \forall \varepsilon > 0$$

(could really be any distance on any  $\mathcal{X}$ )

[Can converge to a r.v.  $X$  too, but we don't need this]

We say the sequence converges in distribution to random variable  $X$  ( $X_n \Rightarrow X, X_n \xrightarrow{d} X$ ) if

$$\mathbb{E} f(X_n) \rightarrow \mathbb{E} f(X) \quad \text{for all bdd, cts } f: \mathcal{X} \rightarrow \mathbb{R}$$

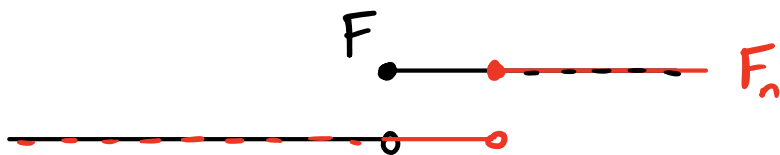
Thm  $X_1, X_2, \dots \in \mathbb{R}, F_n(x) = \mathbb{P}(X_n \leq x), F(x) = \mathbb{P}(X \leq x)$

Then  $X_n \Rightarrow X$  iff  $F_n(x) \rightarrow F(x) \quad \forall x: F \text{ cts at } x$

Also known as weak convergence

Ex: If  $X_n \sim \delta_{\frac{1}{n}}$ ,  $X \sim \delta_0$ , then  $X_n \Rightarrow X$

$$F_n(x) = 1\{\frac{1}{n} \leq x\} \rightarrow 1\{0 \leq x\} \quad \text{except } x=0$$



Prop  $X_n \xrightarrow{P} c$  iff  $X_n \Rightarrow \delta_c$

Proof ( $\Leftarrow$ ) Let  $f_\varepsilon(x) = \max(1, \|x-c\|/\varepsilon) \geq 1\{\|x-c\| > \varepsilon\}$

$$P(\|X_n - c\| > \varepsilon) \leq \mathbb{E} f_\varepsilon(X_n) \rightarrow 0$$

( $\Rightarrow$ )  $f$  bdd, cts, note  $\mathbb{E} f(X) = f(c)$

$$\forall \varepsilon > 0, \exists d(\varepsilon) > 0 \text{ s.t. } \|x - c\| \leq d(\varepsilon) \Rightarrow |f(x) - f(c)| \leq \varepsilon$$

$$\mathbb{E} f(X_n) - f(c) \leq \mathbb{E} \left[ |f(X_n) - f(c)| \cdot (1\{\|X_n - c\| \leq d(\varepsilon)\} + 1\{\|X_n - c\| > d(\varepsilon)\}) \right]$$

$$\leq \varepsilon + P(\|X_n - c\| > d(\varepsilon)) \cdot \sup_x |f(x) - f(c)|$$

$$\leq 2\varepsilon \cdot \sup |f| \quad \text{for suff. large } n \quad \square$$

In a sequence of statistical models  $\mathcal{P}_n = \{P_{n,\theta} : \theta \in \Theta\}$

with  $X_n \sim P_{n,\theta}$ , we say  $\delta_n(X_n)$  is consistent

for  $g(\theta)$  if  $\delta_n(X_n) \xrightarrow{P_\theta} g(\theta)$ , meaning

$$P_\theta(\|\delta_n(X_n) - g(\theta)\| > \varepsilon) \rightarrow 0$$

Usually we omit the index  $n$ ; sequence is implicit.

# Limit Theorems

Let  $X_1, X_2, \dots$  iid random vectors

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$

## Law of large numbers (LLN)

If  $\mathbb{E}|X_i| < \infty$ ,  $\mathbb{E}X_i = \mu$ , then  $\bar{X}_n \xrightarrow{P} \mu$  ( $\bar{X}_n \xrightarrow{a.s.} \mu$ )

## Central limit theorem (CLT)

If  $\mathbb{E}X = \mu \in \mathbb{R}^d$ ,  $\text{Var}(X_n) = \Sigma$  (finite)

Then  $\sqrt{n}(\bar{X}_n - \mu) \Rightarrow N(0, \Sigma)$

[ There are stronger versions of both the LLN & CLT, but this will generally be enough for us ]

# Continuous Mapping

Theorem (Cts Mapping)  $g$  cts;  $X_1, X_2, \dots$  r.v.s

If  $X_n \Rightarrow X$  then  $g(X_n) \Rightarrow g(X)$

If  $X_n \xrightarrow{P} c$  then  $g(X_n) \xrightarrow{P} g(c)$

Proof  $f$  bdd, cts  $\Rightarrow f \circ g$  bdd, cts

If  $X_n \Rightarrow X$  then  $\mathbb{E} f(g(X_n)) \rightarrow \mathbb{E} f(g(X))$

$X_n \xrightarrow{P} c$  special case with  $X \sim \delta_c$   $\square$

Theorem (Slutsky) Assume  $X_n \Rightarrow X$ ,  $Y_n \xrightarrow{P} c$

Then:  $X_n + Y_n \Rightarrow X + c$

$X_n \cdot Y_n \Rightarrow cX$

$X_n / Y_n \Rightarrow X/c$  if  $c \neq 0$

Proof Show  $(X_n, Y_n) \Rightarrow (X, c)$ , apply cts mapping.

[ Wouldn't normally be true that  $X_n \Rightarrow X$ ,  $Y_n \Rightarrow Y$  implies  $(X_n, Y_n) \Rightarrow (X, Y)$  without specifying joint dist. ]

## Theorem (Delta Method)

$$\text{If } \sqrt{n}(X_n - \mu) \Rightarrow N(0, \sigma^2)$$

•  $f(x)$  differentiable at  $x = \mu$

$$\text{Then } \sqrt{n}(f(X_n) - f(\mu)) \Rightarrow N(0, \dot{f}(\mu)^2 \sigma^2)$$

Informal statement:

$$X_n \approx N(\mu, \sigma^2/n) \Rightarrow f(X_n) \approx N(f(\mu), \dot{f}(\mu)^2 \sigma^2/n)$$

Proof  $f(X_n) = f(\mu) + \dot{f}(\mu)(X_n - \mu) + o(X_n - \mu)$

$$\begin{aligned} \sqrt{n}(f(X_n) - f(\mu)) &= \dot{f}(\mu) \cdot \sqrt{n}(X_n - \mu) + \underbrace{\sqrt{n} \cdot o(X_n - \mu)}_{\xrightarrow{p} 0} \\ &= N(0, \dot{f}(\mu)^2 \sigma^2) \end{aligned}$$

Multivariate:  $\sqrt{n}(X_n - \mu) \Rightarrow N_d(0, \Sigma)$ ,  $f: \mathbb{R}^d \rightarrow \mathbb{R}^k$

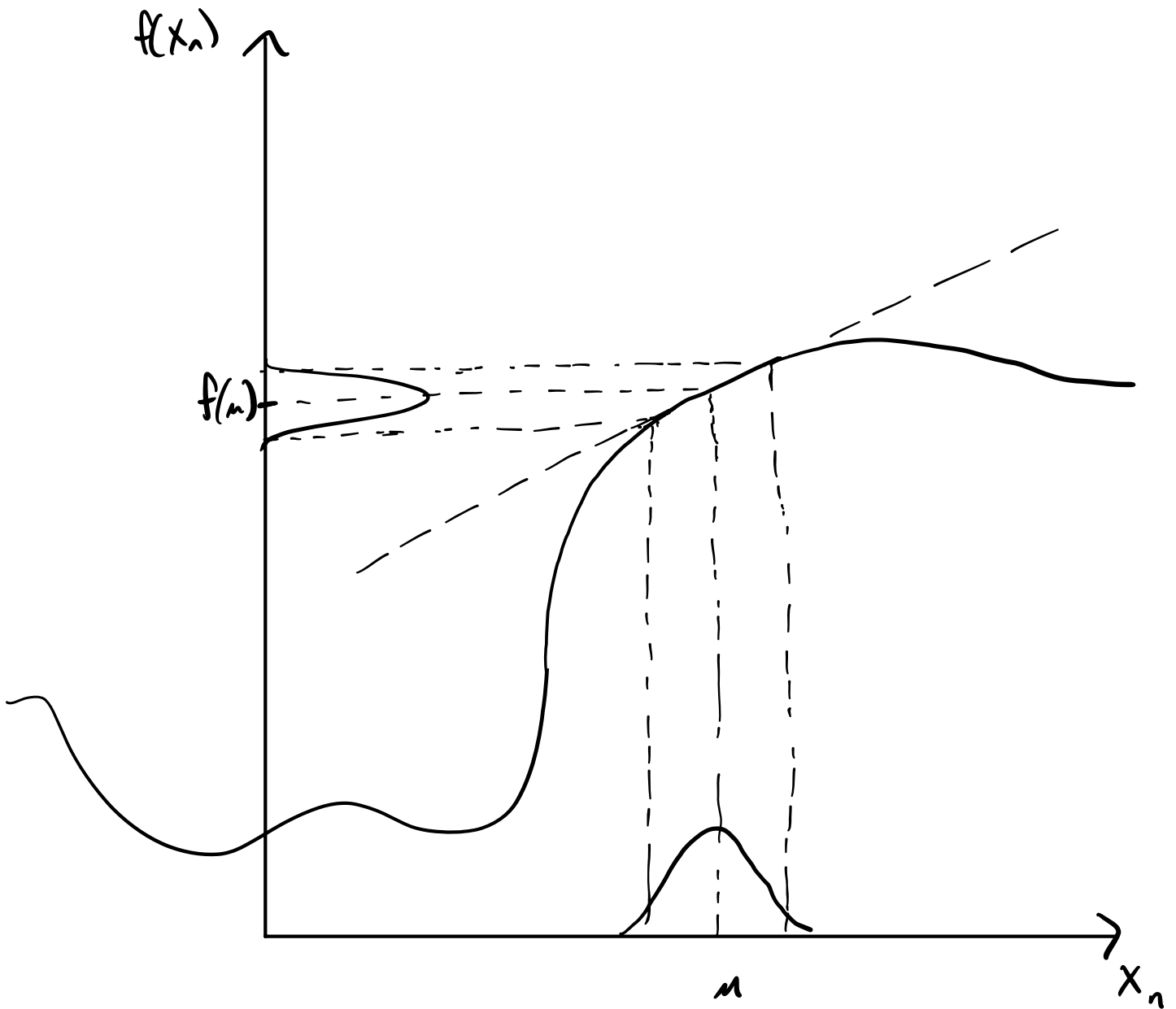
Derivative  $Df(x) = \begin{pmatrix} -\nabla f_1(x) & - \\ \vdots & \\ -\nabla f_k(x) & - \end{pmatrix}$  exists at  $\mu$

$$\text{Then } \sqrt{n}(f(X_n) - f(\mu)) \approx \sqrt{n} Df(\mu) (X_n - \mu)$$

$$\approx N_k(0, Df(\mu) \Sigma Df(\mu)')$$

$$= N(0, \nabla f(\mu)' \Sigma \nabla f(\mu)) \text{ if } k=1$$

# Delta Method



[ Scaling factor doesn't need to be  $\sqrt{n}$ ,  
but need  $X_n - m \xrightarrow{P} 0$  ]



Ex  $X_1, \dots, X_n \stackrel{iid}{\sim} (\mu, \sigma^2)$   
 $Y_1, \dots, Y_n \stackrel{iid}{\sim} (\nu, \tau^2)$   $X, Y$  indep.

For large  $n$ , what is the distribution of  $(\bar{X} + \bar{Y})^2$ ?

1)  $\bar{X} \xrightarrow{P} \mu, \bar{Y} \xrightarrow{P} \nu$  as  $n \rightarrow \infty$

$\Rightarrow (\bar{X} + \bar{Y})^2 \xrightarrow{P} (\mu + \nu)^2 \quad \checkmark$

2)  $\sqrt{n}(\bar{X} - \mu) \Rightarrow N(0, \sigma^2) \quad \sqrt{n}(\bar{Y} - \nu) \Rightarrow N(0, \tau^2)$

Let  $f(x, y) = (x + y)^2$

$\frac{\partial f}{\partial x}(x, y) = \frac{\partial f}{\partial y}(x, y) = 2(x + y)$

$f(\bar{X}, \bar{Y}) \approx N(f(\mu, \nu), \nabla f' \begin{pmatrix} \sigma^2 & 0 \\ 0 & \tau^2 \end{pmatrix} \nabla f / n)$

$= N((\mu + \nu)^2, 4(\mu + \nu)^2(\sigma^2 + \tau^2) / n)$

More accurate:

$\sqrt{n}((\bar{X} + \bar{Y})^2 - (\mu + \nu)^2) \Rightarrow N(0, 4(\mu + \nu)^2(\sigma^2 + \tau^2))$

3) What if  $(\mu + \nu)^2 = 0$ ? Conclusion still holds:

$\sqrt{n}(\bar{X} + \bar{Y})^2 \xrightarrow{P} 0$

Note  $\sqrt{n}\bar{X} + \sqrt{n}\bar{Y} \Rightarrow N(0, \sigma^2 + \tau^2)$  (cts mapping)  
not Slutsky!!

So  $n(\bar{X} + \bar{Y})^2 \Rightarrow (\sigma^2 + \tau^2)\chi_1^2$  (cts mapping)  
 why not delta method?

In general, can do higher-order Taylor expansions for delta method if derivatives  $\neq 0$ :

$$f(X_n) \approx \underbrace{f(\mu)}_{O(1)} + \underbrace{\dot{f}(\mu)(X_n - \mu)}_{O_p(n^{-1/2})} + \underbrace{\frac{\ddot{f}(\mu)}{2}(X_n - \mu)^2}_{O_p(n^{-1})} + \dots$$

If  $\dot{f}(\mu) = 0$ , use second-order term:

$$\begin{aligned} n(f(X_n) - f(\mu)) &\approx \frac{\ddot{f}(\mu)}{2} (\sqrt{n}(X_n - \mu))^2 \\ &\approx \frac{\ddot{f}(\mu)\sigma^2}{2} \chi_1^2 \end{aligned}$$