

Modelling and Analyzing Unevenly Spaced Measurements

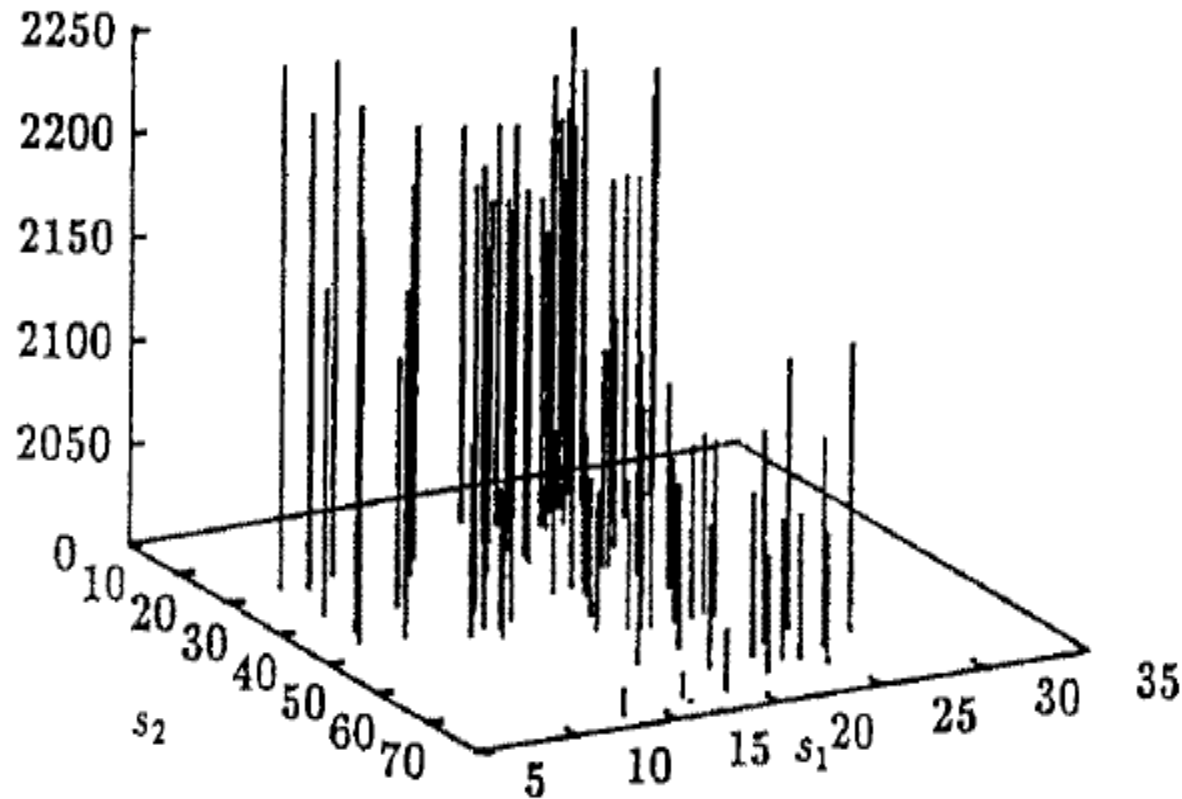
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$$2\pi \neq 1$$

Shumway analysis.



Water heights, Saratoga Valley, Wyoming

Yucel and Shumway (1996) *Stochastic Hydrology & Hydrolics*

Underlying surface, $y(\mathbf{s})$, \mathbf{s} in \mathbb{R}^2

Spatial marked point process data, $\{\mathbf{s}_j, M_j\}$, $M_j = y(\mathbf{s}_j)$

$y(\mathbf{s}) = \mathbf{v}(\mathbf{s})'\boldsymbol{\beta} + x(\mathbf{s}) + e(\mathbf{s})$, \mathbf{s} : location

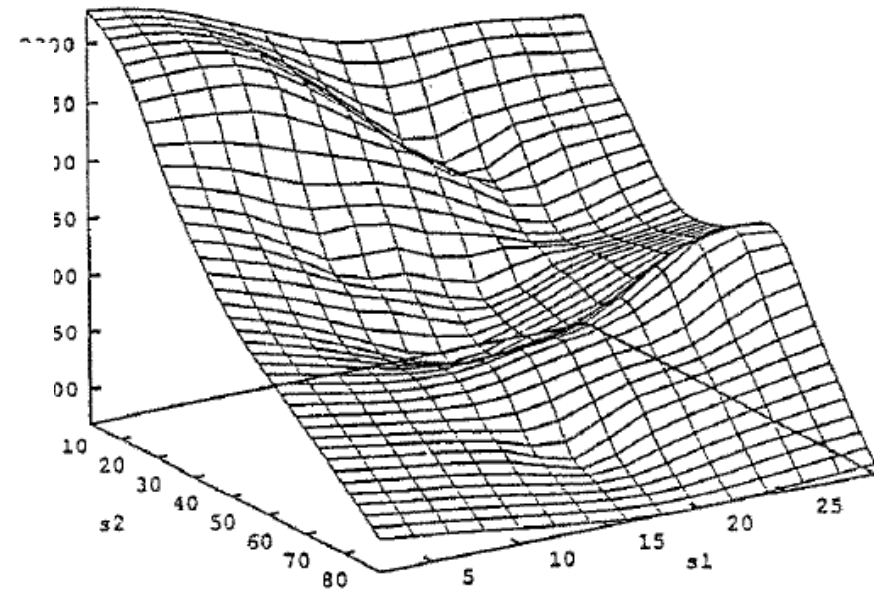
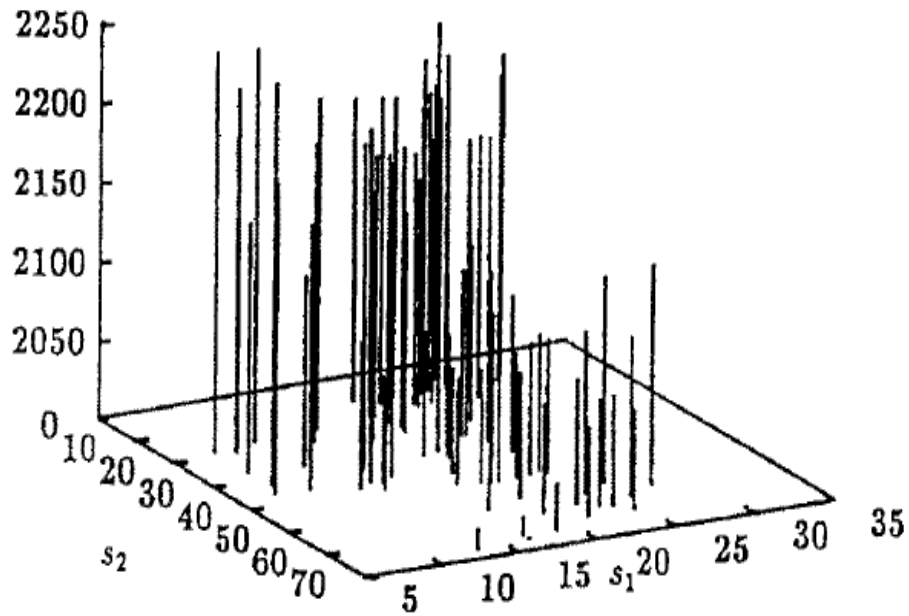
\mathbf{v} : known regressor function

x : signal, stationary Gaussian, autocov $C(t|\phi)$, spectrum $f(\Lambda|\psi)$

e : stationary Gaussian white noise

mle obtained via ECM after approximating C

Estimated $\mathbf{v}(\mathbf{s})'\beta + \mathbf{x}(\mathbf{s})$



Method and program checked by simulation

Other Shumway work.

"Some applications of the EM algorithm to analyzing incomplete time series data." Pp 290-324 in *Time Series Analysis of Irregularly Observed Data*, ed. E. Parzen (1984)

"Dynamic mixed models for irregularly observed time series." *Resenhas* 4, 433-456 (2000).

missing values in equispaced time series

state space models

Many important random process methods and programs expect equi-spaced data values.

Variants needed and effects to be understood.

Causes: missing values, irregular sampling, explanatory differ, measurement mechanism details, triggered observation, enforced gaps, ...

Reasons may include speed of computation.

Uneven/ irregular measurements described via point process of times, locations, ...

Process $\{Y(t)\}$ defined for all t , but data $(\tau_k, Y(\tau_k))$ available. (Hybrid process)

Suppose stationary, mean $E(Y(t)) = c_Y$

Given the sequence of times τ_1, \dots, τ_N for the interval $0 < t \leq T$, properties of average

$$[Y(\tau_1) + \dots + Y(\tau_N)]/N \quad (*)$$

as estimate of c_Y ?

Take τ 's as fixed or random?

Formal setup.

$N(t)$: number of τ 's in the interval $(0,t]$

$dN(t)$: number in interval $(t,t+dt]$ 0 or 1

Sampled process, $\{Y(\tau_k)\}$,

$$dX(t) = Y(t)dN(t)$$

estimate (*) of c_N

$$\int_0^T Y(t) dN(t)/N(T) \quad (*)$$

Dirac comb

$$dN(t) = \sum_k \delta(t - \tau_k) dt$$

$$dX(t) = \sum_k Y(\tau_k) \delta(t - \tau_k) dt$$

$\delta(x)$: density of r. v. $X = 0$ with prob 1

Stochastic stationary case.

Suppose there exist C's of bounded variation with

$$\text{cum}\{dN(t+u_1), \dots, dN(t+u_k), dN(t)\} = dC_{N\dots N}(u_1, \dots, u_k)dt$$

"cum": moment measure of joint dependence

C's define parameters and provides mixing conditions

One sees that

$$E\{N(T)\} = TC_N$$

$$E\{T^{-1} \int_0^T [N(t+u)-N(t)]dN(t)\} = uC_N^2 + C_{NN}(u)$$

Back to estimating c_Y . Suppose given $Y(\tau_k)$ for $(0, T]$.

If assume $\{Y(t)\}$ and $\{N(t)\}$ independent, then average (*) asymptotically normal with mean c_Y and large sample variance

$$T^{-1}[\int c_{YY}(u)du + \int c_{YY}(u) dC_{NN}(u)/C_N^2]$$

In homogenous Poisson case variance is

$$T^{-1} \int c_{YY}(u)du$$

One can look for $\{N(t)\}$ to obtain small variance.

Irregular sampling can be an advantage with good choice.

View τ_k as fixed. cp. GHA

Assumption:

τ_1, τ_2, \dots increasing positive numbers

$$|N(s) - N(t)| \leq A + B|s - t|, \quad A, B \text{ finite}$$

$$\lim N(T)/T = C_N$$

$$\lim T^{-1} \int_0^T [N(t+u) - N(t)] dN(t) = M_{NN}(u) \quad \text{as } T \rightarrow \infty$$

Large sample variance now

$$T^{-1} \int c_{YY}(u) dM_{NN}(u) / C_N^2$$

In practice $Y(\tau_k)$ often interpolated to obtain equispaced values

Irregular sampling can be an advantage

Can reduce aliasing problem

Estimate at a one frequency also estimate at second

Occurs if signal not sampled often

Suppose the series $\{Y(t)\}$ and the point process $\{N(t)\}$ are independent and stationary with pertinent moments existing.

Define the process with stationary increments

$$dX(t) = Y(t)dN(t)$$

When Y has mean 0

$$f_{XX}(\lambda) = c_N^2 f_{YY}(\lambda) + \int f_{NN}(\lambda - \alpha) f_{YY}(\alpha) d\alpha \quad [\text{CHECK}]$$

Can solve for ...

No aliasing when ...

Works Poisson

No interpolation is needed.

One can use a periodogram to assess if a process is bandlimited white noise when only irregularly sampled values are available and for residual analysis. (Missing seasonal?)

Having fit a model to sampled data one can examine the periodogram of the residuals for whiteness

Amazonas river flow.

Solimoes flows from Peru

Divides into, Parana do Careiro and Amazon near Manaus

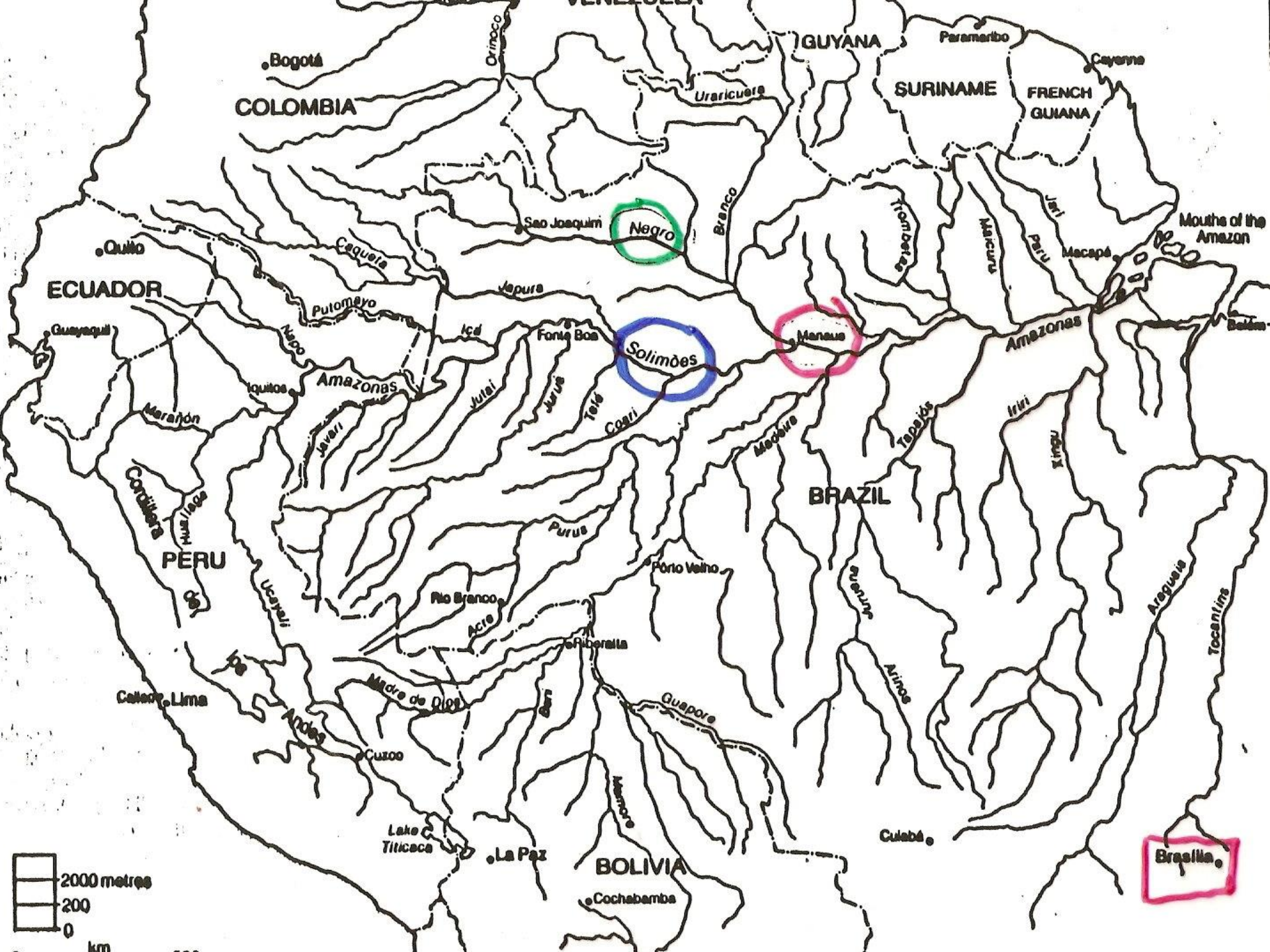
Change of flow in Parana could damage habitat

Does proportion of discharge entering Parana depend on Solimoes discharge level?

Solimoes flow increasing?

Careiro and Solimoes relationship changing?

Difficulty: flows measured at different irregular times





Rio Negro - Solimoes confluence

Single ship - measurements at different times and locations

Flow estimates via

discharge = area of section * velocity

Measurement times $\{\sigma_j\}$ and $\{\tau_k\}$

X(t): Solimoes $\{X(\sigma_j)\}$ explanatory

Y(t): Careiro $\{Y(\tau_k)\}$ response

Models/questions.

$Y(t) = g(X(t)) + \text{noise}$ $g(\cdot)$ linear?

$Y(t)/X(t) = h(X(t)) + \text{noise}$ $h(x)$ constant? $h(x,t)$?

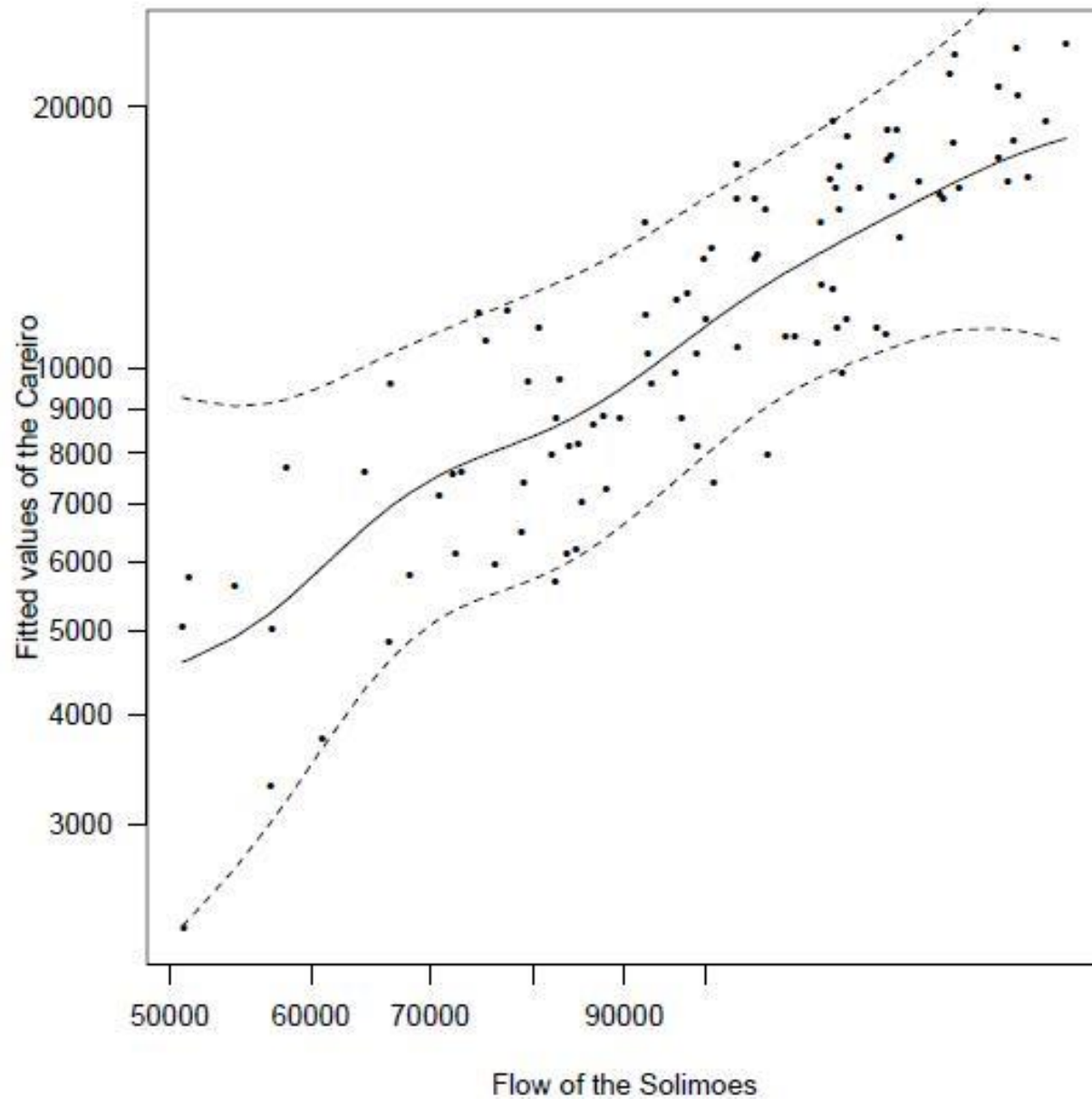
Time lag, δ , to handle travel time Manacapuru to Careiro

$$Y(t) = g(X(t-\delta)) + \text{noise}$$

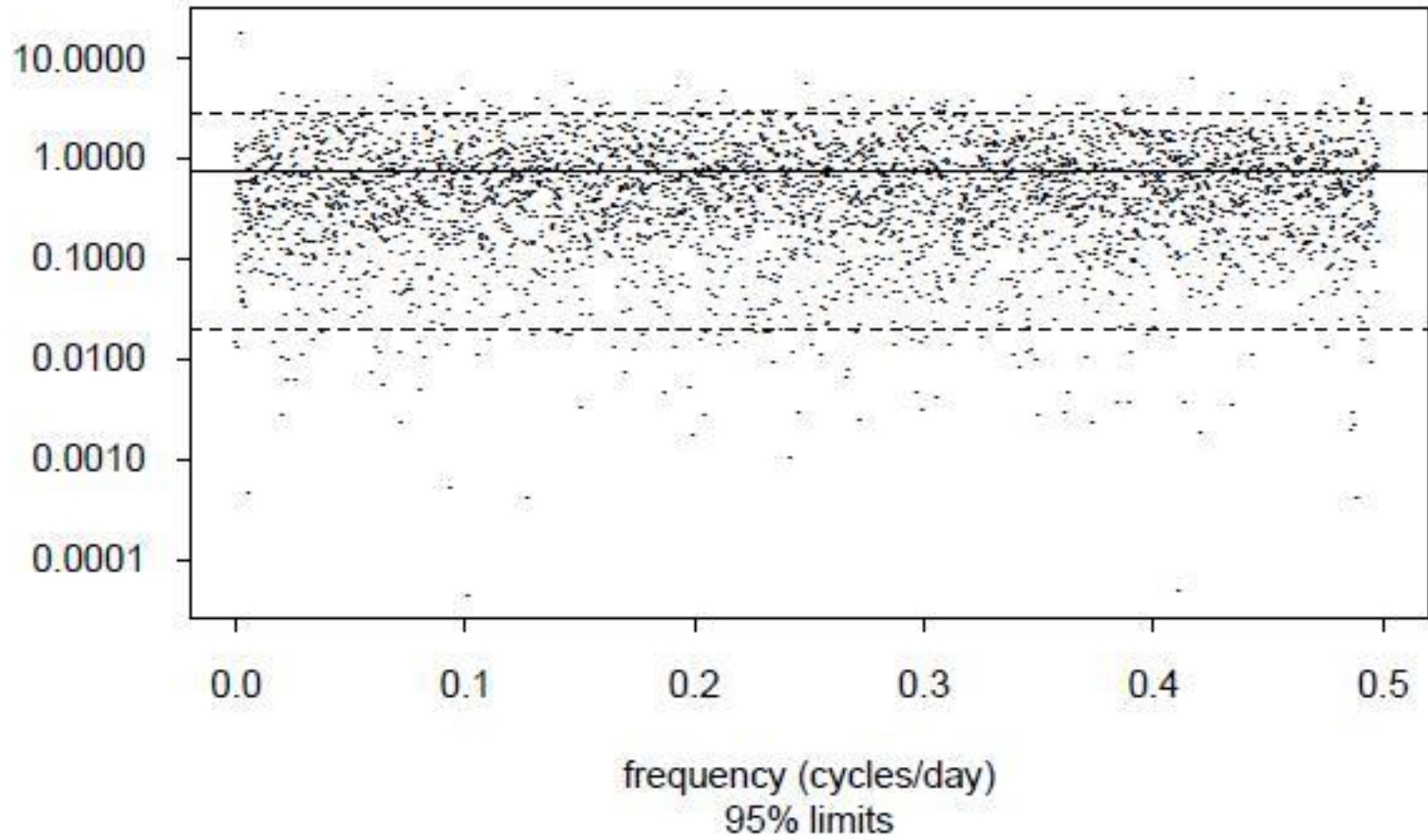
Interpolation. Smoothing to estimate $X(\tau_k - \delta)$

Smoothing to estimate $g(X)$

Fitted values of the Careiro vs. the Solimoes



Periodogram of the residuals



Error of interpolation.

$$n=M(T) \quad F_n(\sigma) = n^{-1}\#\{\sigma_j \leq \sigma\} \quad W_n(\sigma) = b_n^{-1}W(b_n^{-1}\sigma)$$

$$d_n = \sup |F_n(\sigma) - F(\sigma)|$$

Interpolant.

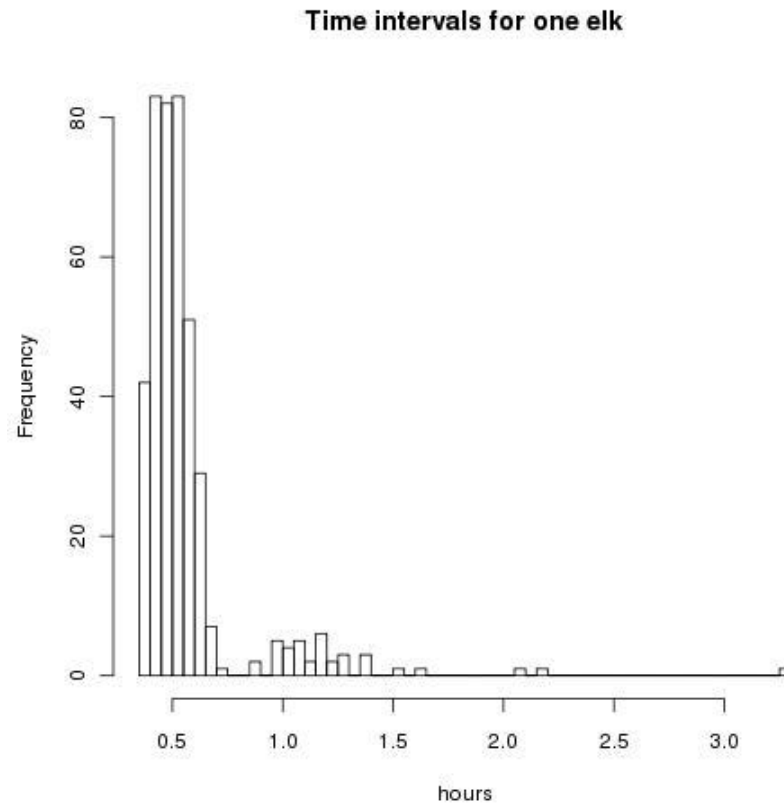
$$\begin{aligned} X(\sigma) &\sim \sum W_n(\sigma - \sigma_j) X(\sigma_j) / \sum W_n(\sigma - \sigma_j) \\ &= \int W_n(\sigma - s) X(s) dF_n(s) / \int W_n(\sigma - s) dF_n(s) \\ &\sim X(\sigma)f(\sigma)/f(\sigma) + O(b_n^{-2}) + O(b_n^{-1}d_n) \quad \text{uniformly} \end{aligned}$$

Work with σ_j/T as necessary

Animal tracks. 94 elk 1991 - 1993 LORAN

Starkey Experimental Forest and Range, Eastern Oregon

time differences: irregular, and varied with animal



Vector field model. Location $\mathbf{r} = (x,y)$

$$(\mathbf{r}(t_{i+1}) - \mathbf{r}(t_i)) / (t_{i+1} - t_i) = \boldsymbol{\mu}(\mathbf{r}(t_i), \langle t_i \rangle) + \boldsymbol{\sigma} \mathbf{Z}_{i+1} / \sqrt{(t_{i+1} - t_i)}$$

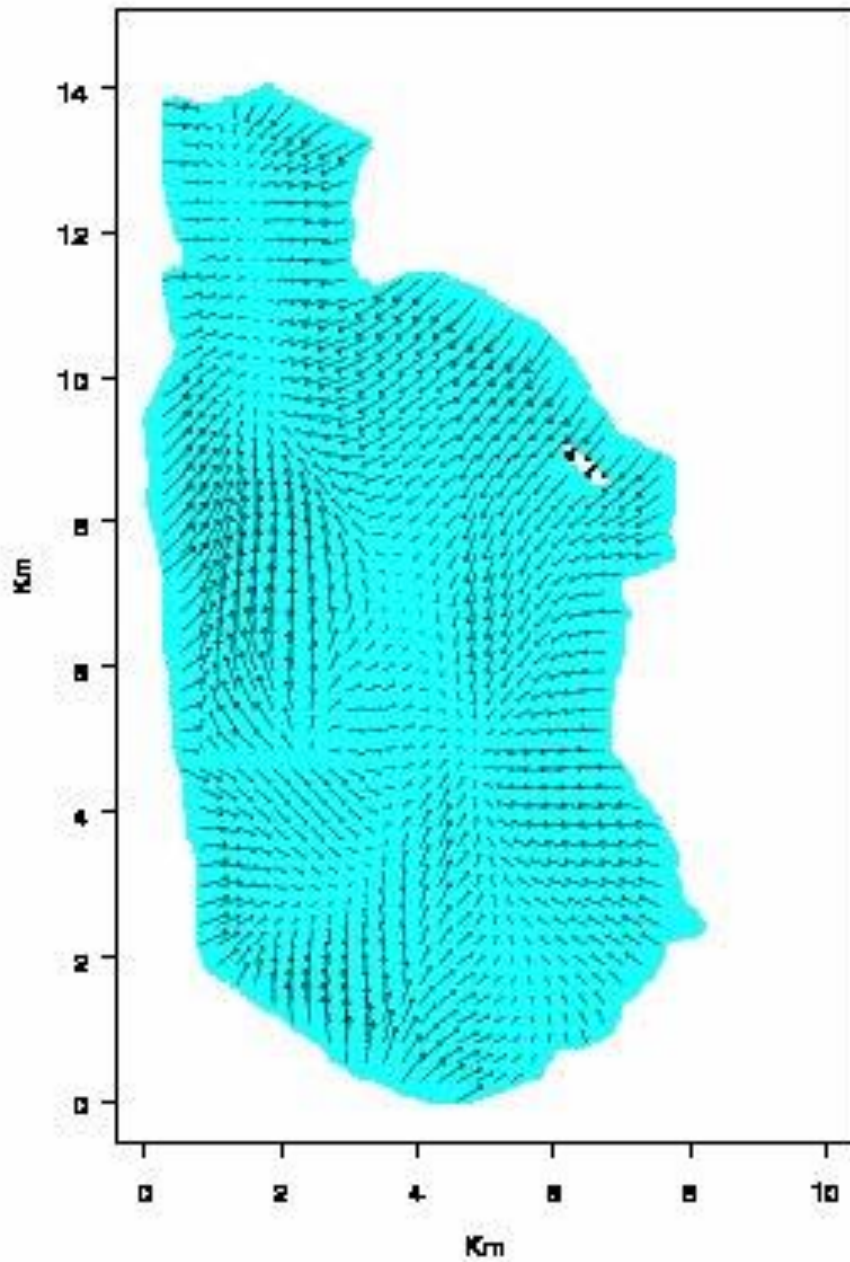
LHS: velocity, $\langle t \rangle$: time of day

Motivated by SDE

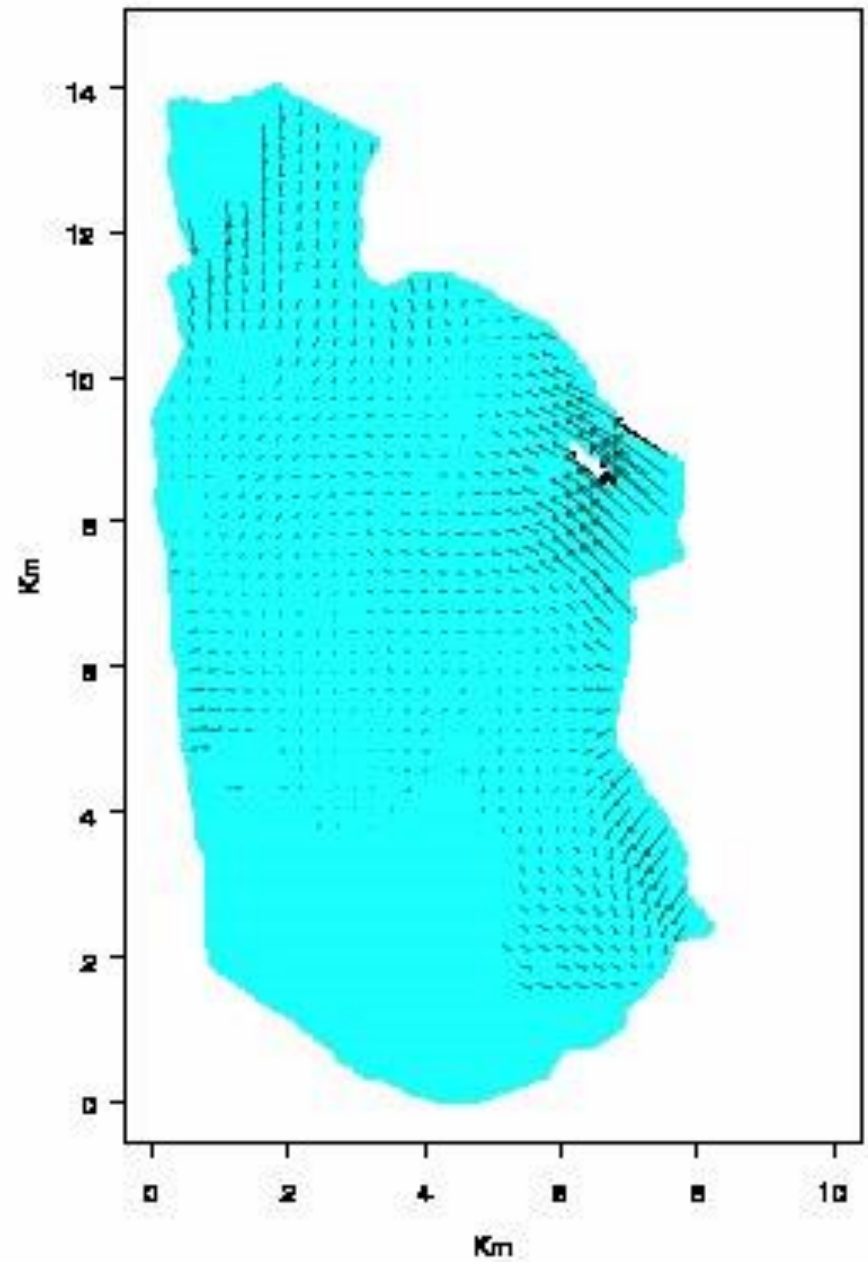
$$d\mathbf{r}(t) = \boldsymbol{\mu}(\mathbf{r}(t), \langle t \rangle) dt + \boldsymbol{\sigma} d\mathbf{B}(t)$$

Robust fit via `gam()`

Elk 0500 hours Spring



Deer 0500 hours Spring



Residuals provided evidence for 24 hour period
included in model

spectra of hybrid estimated by averaging periodograms

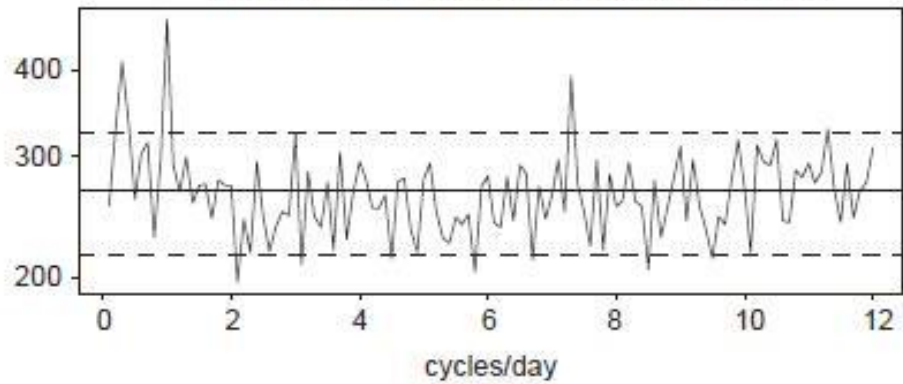
$$(2\pi T)^{-1} |\sum \exp\{-i\lambda\tau_k\} r_{\text{residual}}(\tau_k)|^2$$

Fundamental relation

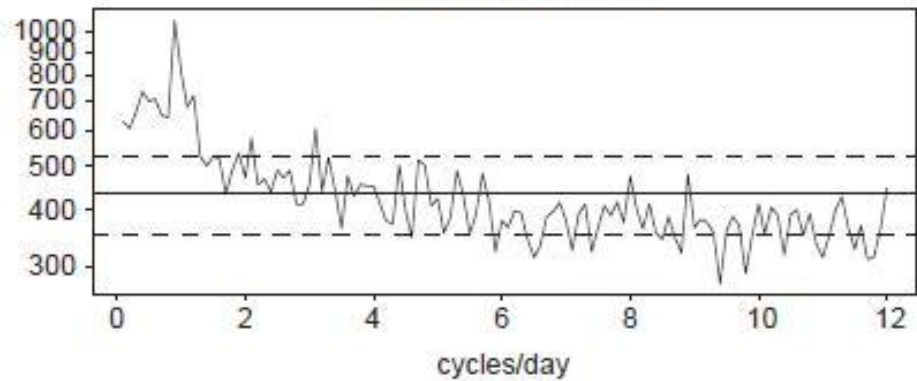
$$c_N^2 f_{\text{noise}}(\lambda) + \int f_{\text{NN}}(\lambda - \alpha) f_{\text{noise}}(\alpha) d\alpha \quad (*)$$

flat if noise spectrum bandlimited white [CHECK]

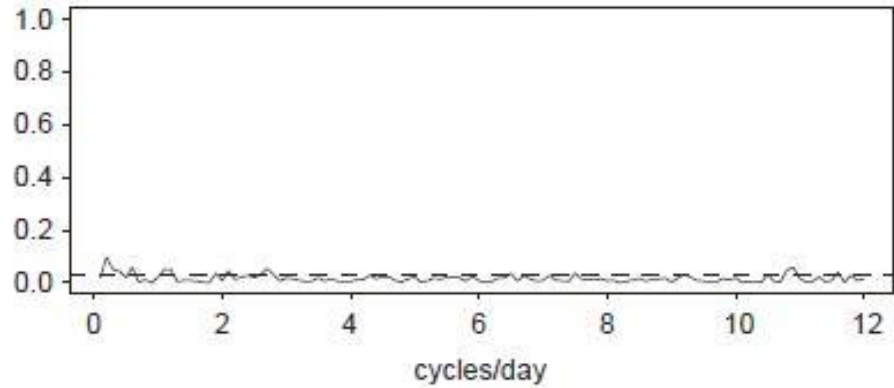
Elk standardized residuals: X-spectrum estimate



Y-spectrum estimate



Coherence estimate



Evidence for low frequency departure for N-S component

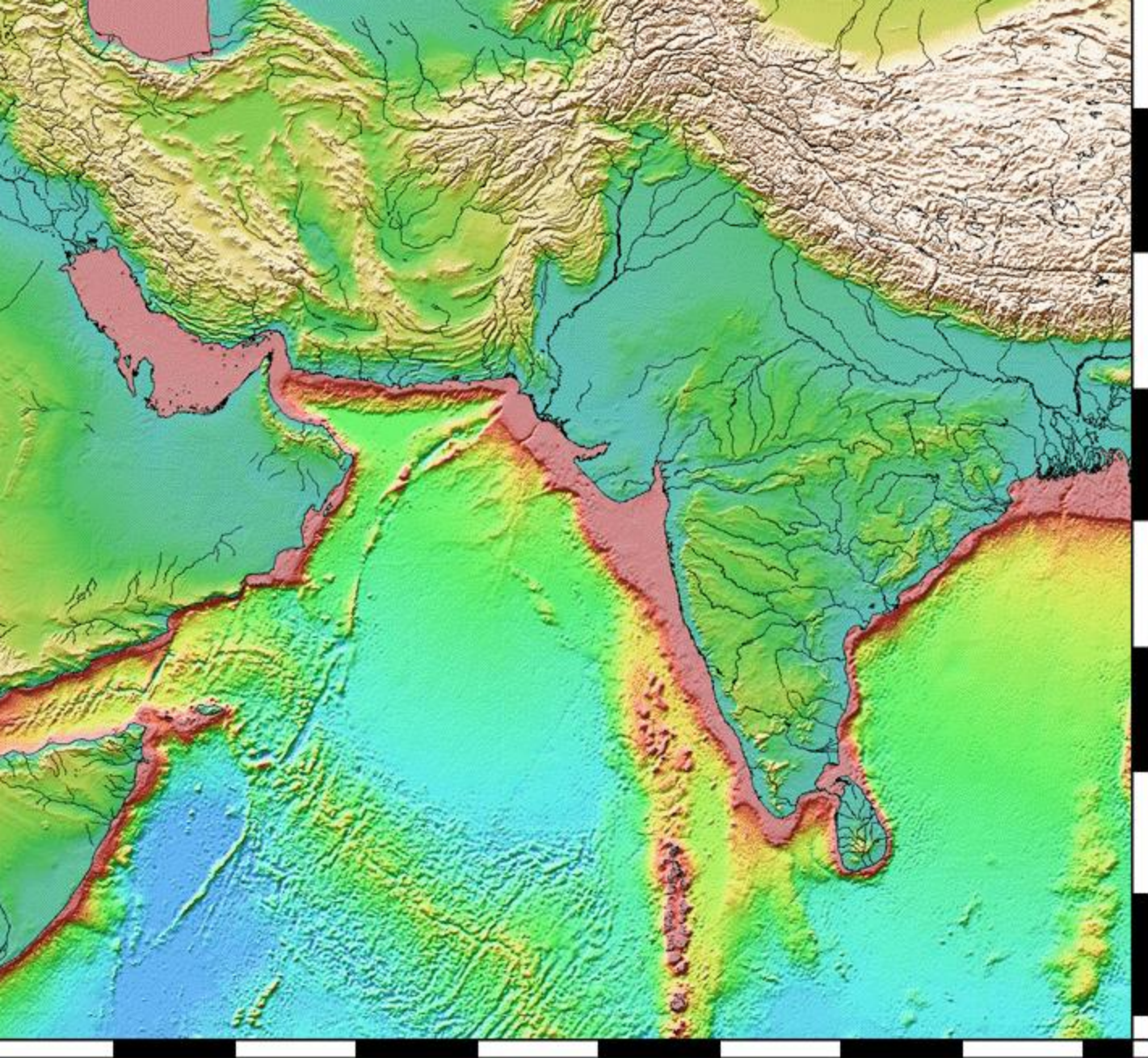
Whaleshark tagging study. Off Kenya

To study ecology, e.g. where they travelled and foraged, and when?

Goal species protection.



Brent



50°E

60°E

70°E

80°E

90°E

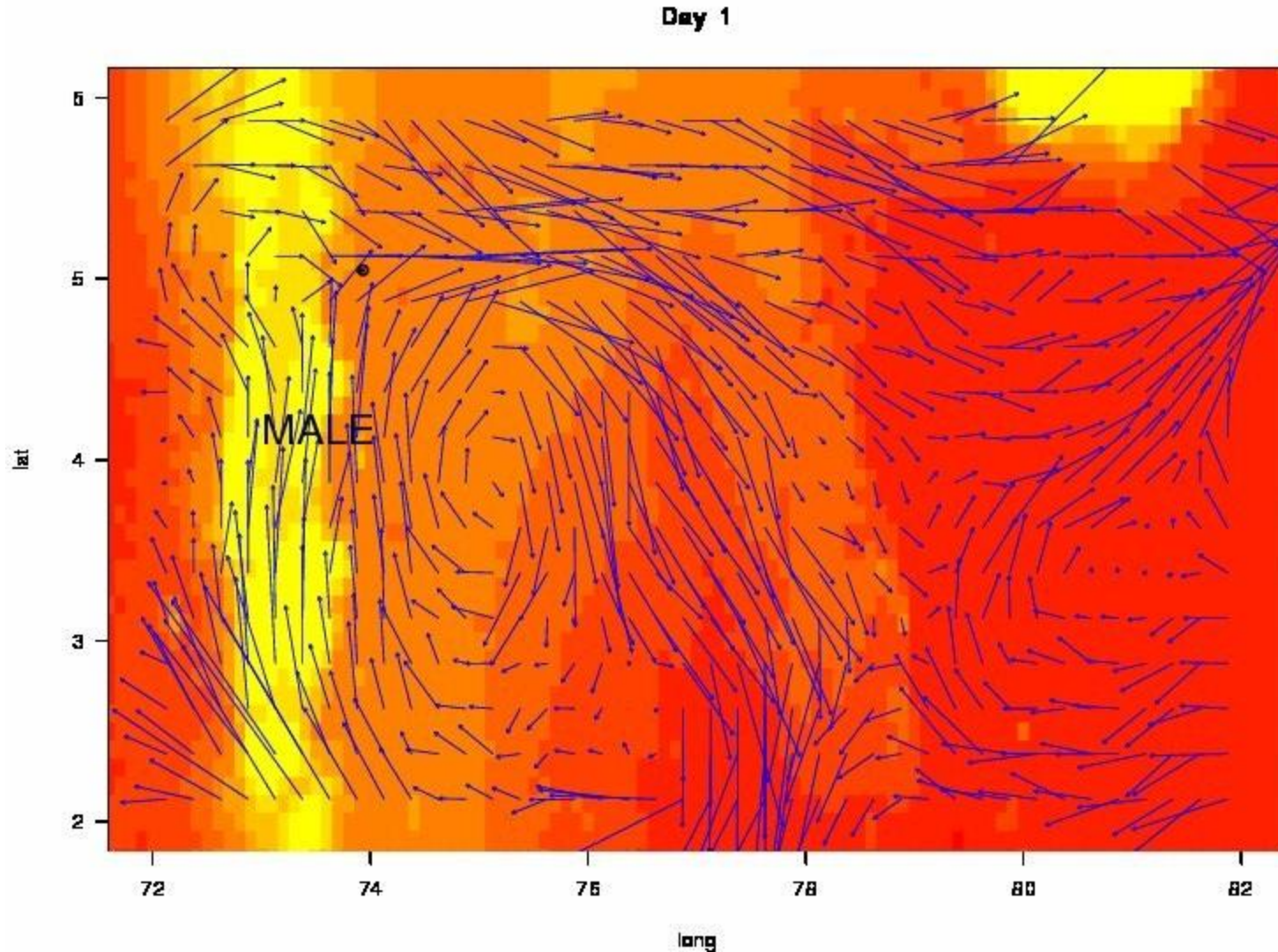
29 June - 19 July, 2008 Indian Ocean

One shark lost its tag, drifted til batteries expired

Brent realized could check NOAA's sea surface current model results

Unequally spaced location times, about 250 time points

Geostrophic currents for June 29, 2008 (ten day composite)
some days missing (clouds)



display of sequence of locations

Functional stochastic differential equation (FSDE)

$$d\mathbf{r}(t) = \boldsymbol{\mu}(\mathbf{H}(t), t)dt + \boldsymbol{\sigma}(\mathbf{H}(t), t)d\mathbf{B}(t)$$

$\mathbf{H}(t)$: a history based on the past, $\{\mathbf{r}(s), s \leq t\}$

Process Markov when $\mathbf{H}(t) = \{\mathbf{r}(t)\}$

Interpretation

$$\mathbf{r}(t) - \mathbf{r}(0) = \int_0^t \boldsymbol{\mu}(\mathbf{H}(s), s)ds + \int_0^t \boldsymbol{\sigma}(\mathbf{H}(s), s)d\mathbf{B}(s)$$

Approximation

$$\mathbf{r}(t_{i+1}) - \mathbf{r}(t_i) = \boldsymbol{\mu}(\mathbf{H}(t_i), t_i)(t_{i+1} - t_i) + \boldsymbol{\sigma}(\mathbf{H}(t_i), t_i)\sqrt{t_{i+1} - t_i}\mathbf{Z}_{i+1}$$

Analysis.

How was the irregular sampling handled?

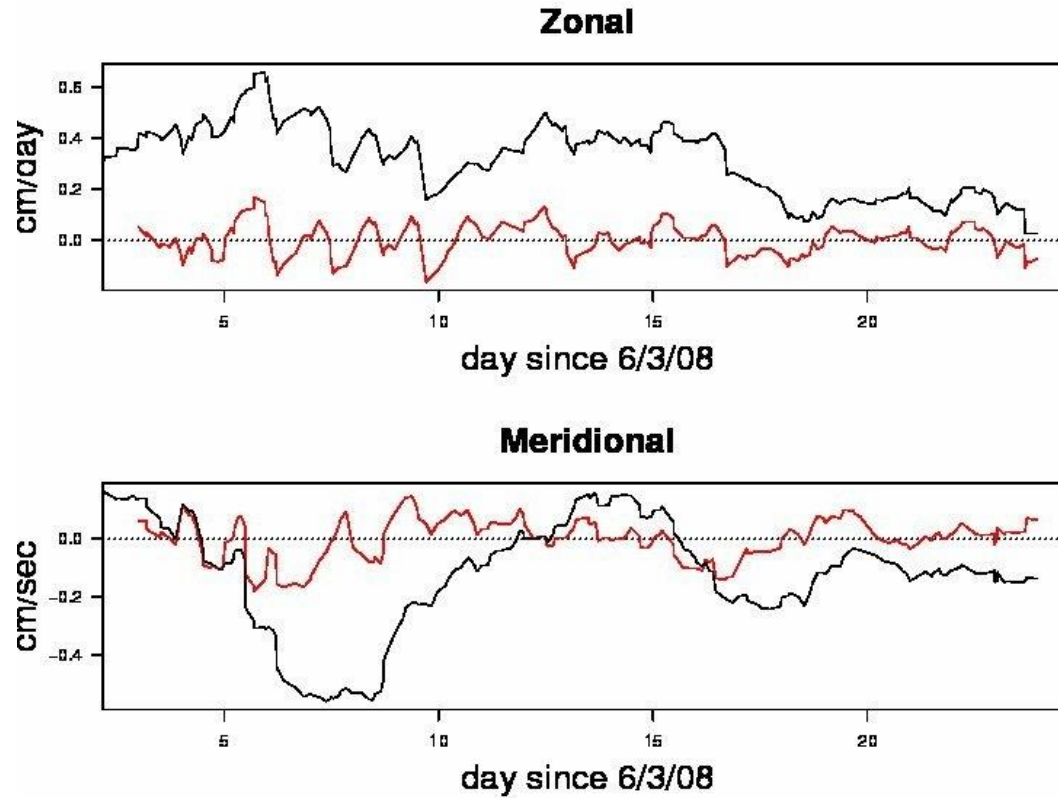
Interpolated

Reduced tag data to 46 contiguous 12 hour periods

median values

Estimated local zonal and meridional speeds from tag locations and ... [Get straight]

Tag currents and residuals from fits



Current work.

Moving particles.

Are they interacting?

Pairwise interaction model

$$d\mathbf{r}_i(t) = \Psi(\mathbf{r}_i(t))dt + \sum_{j \neq i} \Phi(\mathbf{r}_i(t) - \mathbf{r}_j(t))dt + \sigma d\mathbf{B}_i(t)$$

$i = 1, \dots, p$ for some functions Ψ, Φ

Different time spacings for different models

Other cases.

P.p. case: deletions/thinning

t in $0, \pm 1, \pm 2, \dots$

missings, gaps, ...

Discussion.

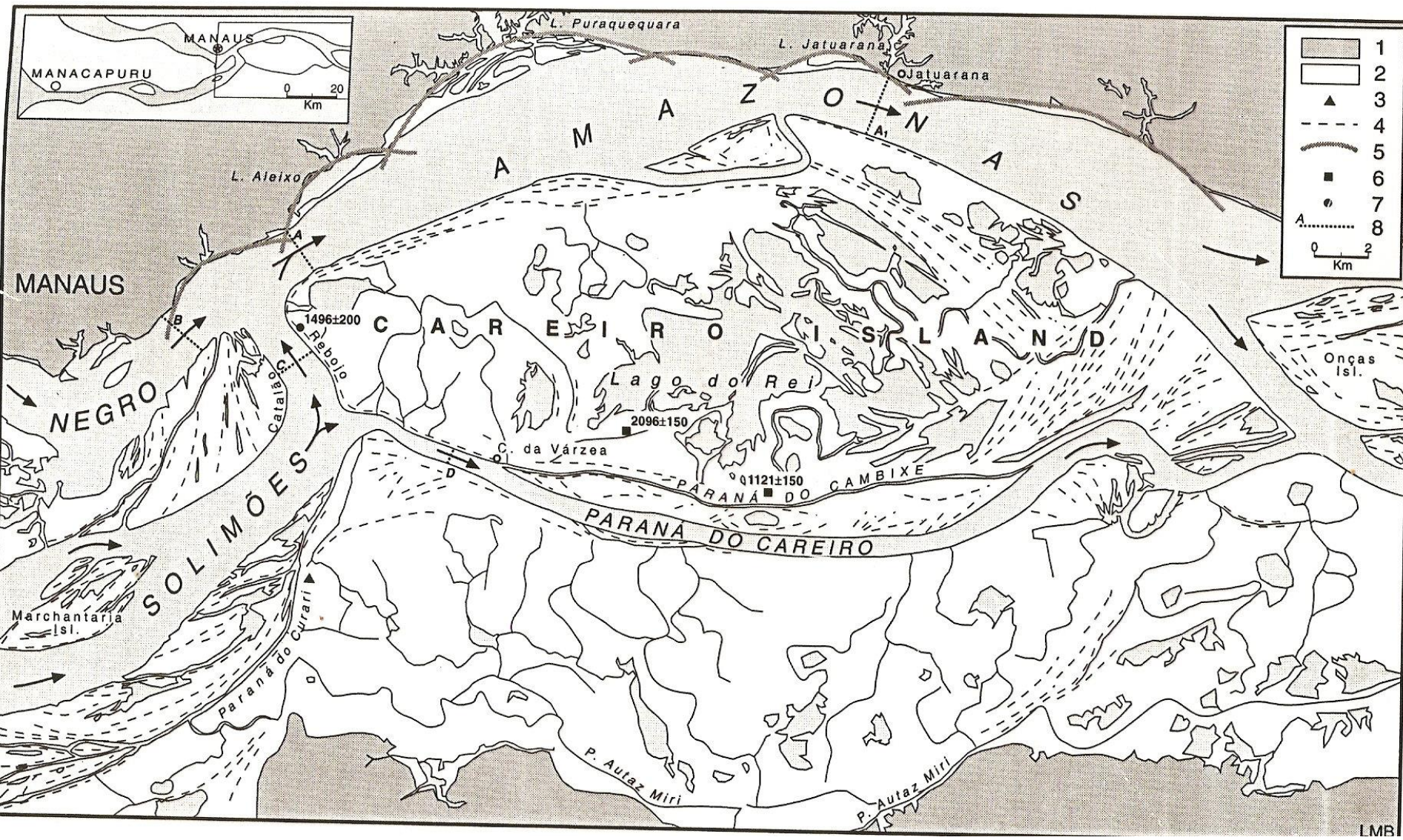
Sometimes can model the missingness

There are FFTs for the irregular case

Common features of examples

Acknowledgements.

B. Stewart, ...





Water Consumption in Edmonton During Olympic Gold Medal Hockey Game

