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The δ -method

If $\mu = (\mu_1, \dots, \mu_d)$

$$\hat{\mu} = (\hat{\mu}_1, \dots, \hat{\mu}_d)$$

$$\hat{\mu} \sim N_d\left(\mu, \text{diag}\left(\frac{\sigma_1^2}{n_1}, \dots, \frac{\sigma_d^2}{n_d}\right)\right)$$

for large n_1, \dots, n_d

and

$$\sigma = h(\mu), \quad h \text{ differentiable}$$

then

$$\nabla h \neq 0$$

$$\hat{\sigma} \sim N(\sigma, AV(\hat{\sigma}))$$

where

$$AV(\hat{\sigma}) = \sum_k \left(\frac{\partial h}{\partial \mu_k}(\mu) \right)^2 \frac{\sigma_k^2}{n_k}$$

Proof. Via Taylor series and convergence in probability. $(\hat{\mu}_k \xrightarrow{p} \mu_k)$

$$ASD = \sqrt{AV}$$

$$SE(\hat{\sigma}) = \sqrt{\sum_k \left(\frac{\partial h}{\partial \mu_k}(\hat{\mu}) \right)^2 \frac{\hat{\sigma}_k^2}{n_k}}$$

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Two sample parameters

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$$B(n_1, \pi_1) \parallel B(n_2, \pi_2)$$

Difference: $\pi_2 - \pi_1 = \tau_d$

Relative risk: $\pi_2 / \pi_1 = \tau_{rr}$

Odds ratio: $\frac{\pi_2 / (1 - \pi_2)}{\pi_1 / (1 - \pi_1)} = \tau_{or}$

$$SE(\hat{\tau}_d) = \sqrt{\frac{\hat{\pi}_1(1-\hat{\pi}_1)}{n_1} + \frac{\hat{\pi}_2(1-\hat{\pi}_2)}{n_2}}$$

$$SE(\hat{\tau}_{rr}) = \hat{\tau}_{rr} \sqrt{\frac{1-\hat{\pi}_1}{n_1 \hat{\pi}_1} + \frac{1-\hat{\pi}_2}{n_2 \hat{\pi}_2}}$$

$$SE(\hat{\tau}_{or}) = \hat{\tau}_{or} \sqrt{\frac{1}{n_1 \hat{\pi}_1 (1-\hat{\pi}_1)} + \frac{1}{n_2 \hat{\pi}_2 (1-\hat{\pi}_2)}}$$

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Abortion attitudes

Male vs. female attitudes

M support legal $y_1 = 319$, $n_1 = 600$ F " " $y_2 = 309$, $n_2 = 500$

$$H_0: \pi_1 = \pi_2 = 0 = \pi_0$$

$$\hat{\pi}_1 = .532 \quad \hat{\pi}_2 = .618 \quad \hat{\pi} = -.086$$

$$SE(\hat{\pi}) = \sqrt{\frac{\hat{\pi}_1(1-\hat{\pi}_1)}{n_1} + \frac{\hat{\pi}_2(1-\hat{\pi}_2)}{n_2}}$$

$$= .030$$

$$\text{Wald statistic } W = \frac{(\hat{\pi} - \pi_0)}{SE(\hat{\pi})} = -2.87$$

$$P\text{-value } 2[1 - \Phi(|W|)] = .0042$$

$$H_A: \pi \neq 0$$

Infer women more likely to support

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Relative risk: π_2 / π_1

$$H_0: \gamma = \pi_2 / \pi_1 = 1 = \gamma_0$$

$$\hat{\gamma} = \hat{\pi}_2 / \hat{\pi}_1 = 1.162$$

$$SE(\hat{\gamma}) = \hat{\gamma} \sqrt{\frac{1 - \hat{\pi}_1}{n_1 \hat{\pi}_1} + \frac{1 - \hat{\pi}_2}{n_2 \hat{\pi}_2}}$$

$$= 0.0604$$

$$W = \frac{(\hat{\gamma} - \gamma_0)}{SE(\hat{\gamma})} = 2.677$$

$$P\text{-value: } 2 [1 - \Phi(|W|)] = 0.0074$$

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Grandchildren

$$n = 3$$

$y = 2$ girls

$$H_0: \pi = 1/2$$

$$\text{P-value} = \text{Prob}\{Y \geq 2\} =$$

$$\text{Prob}\{GGB, GGG\}$$

$$= \binom{3}{2} \left(\frac{1}{2}\right)^2 \left(\frac{1}{2}\right) + \left(\frac{1}{2}\right)^3$$

$$= \frac{4}{8}$$

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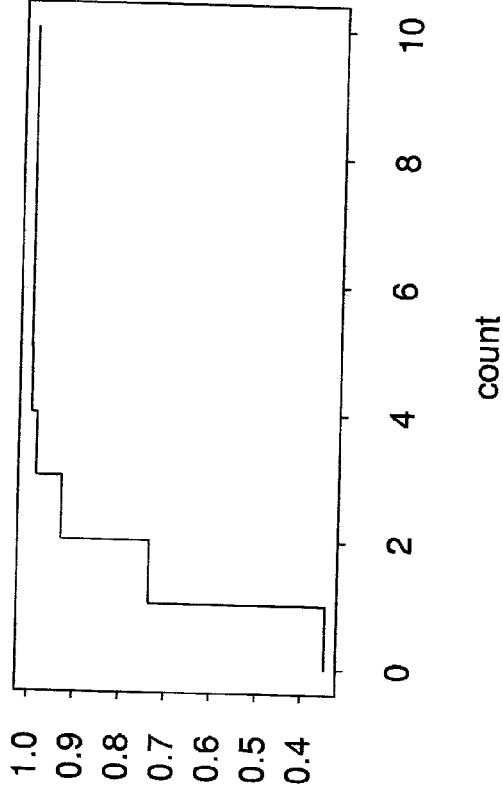
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Exact P-values.

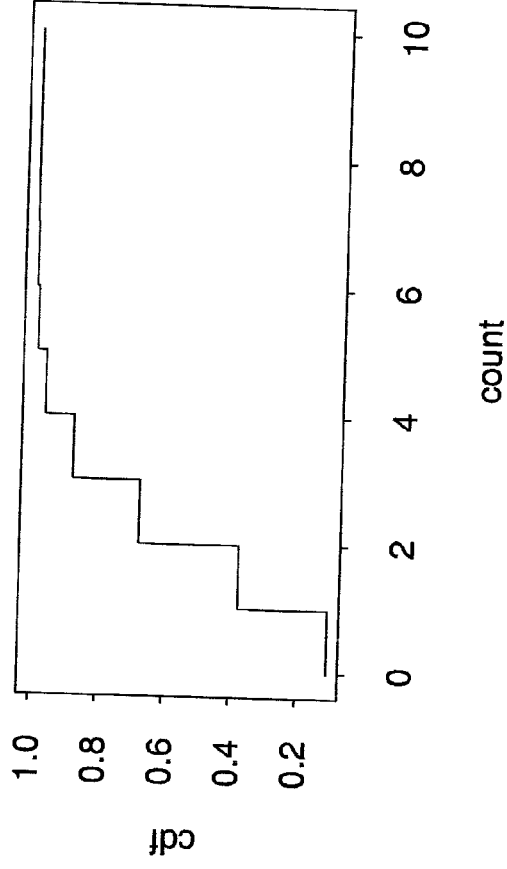
Binomial cdf

$$\begin{aligned} \text{Prob}\{Y \leq y\} &= F(y; n, \pi) \\ &= \sum_{z=0}^y f(z; n, \pi) \end{aligned}$$

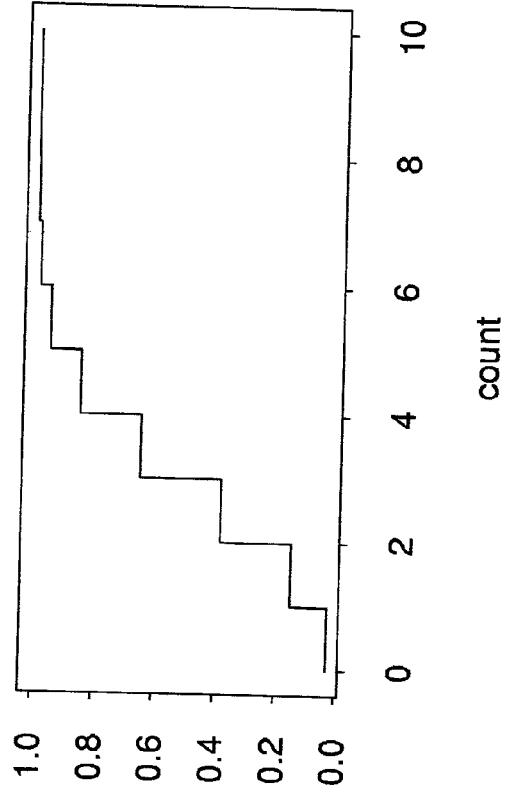
Binomial cdf, $n = 10$, param = 0.1



Binomial cdf, $n = 10$, param = 0.2



Binomial cdf, $n = 10$, param = 0.3



Binomial cdf, $n = 10$, param = 0.4

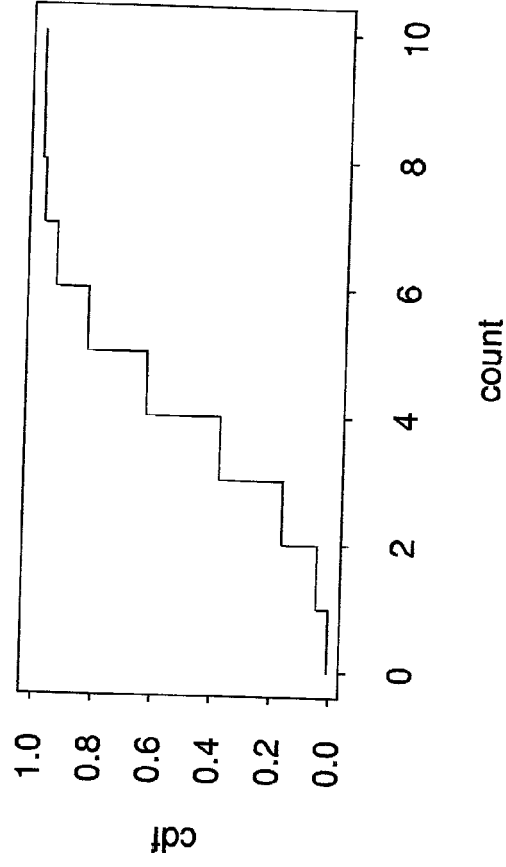
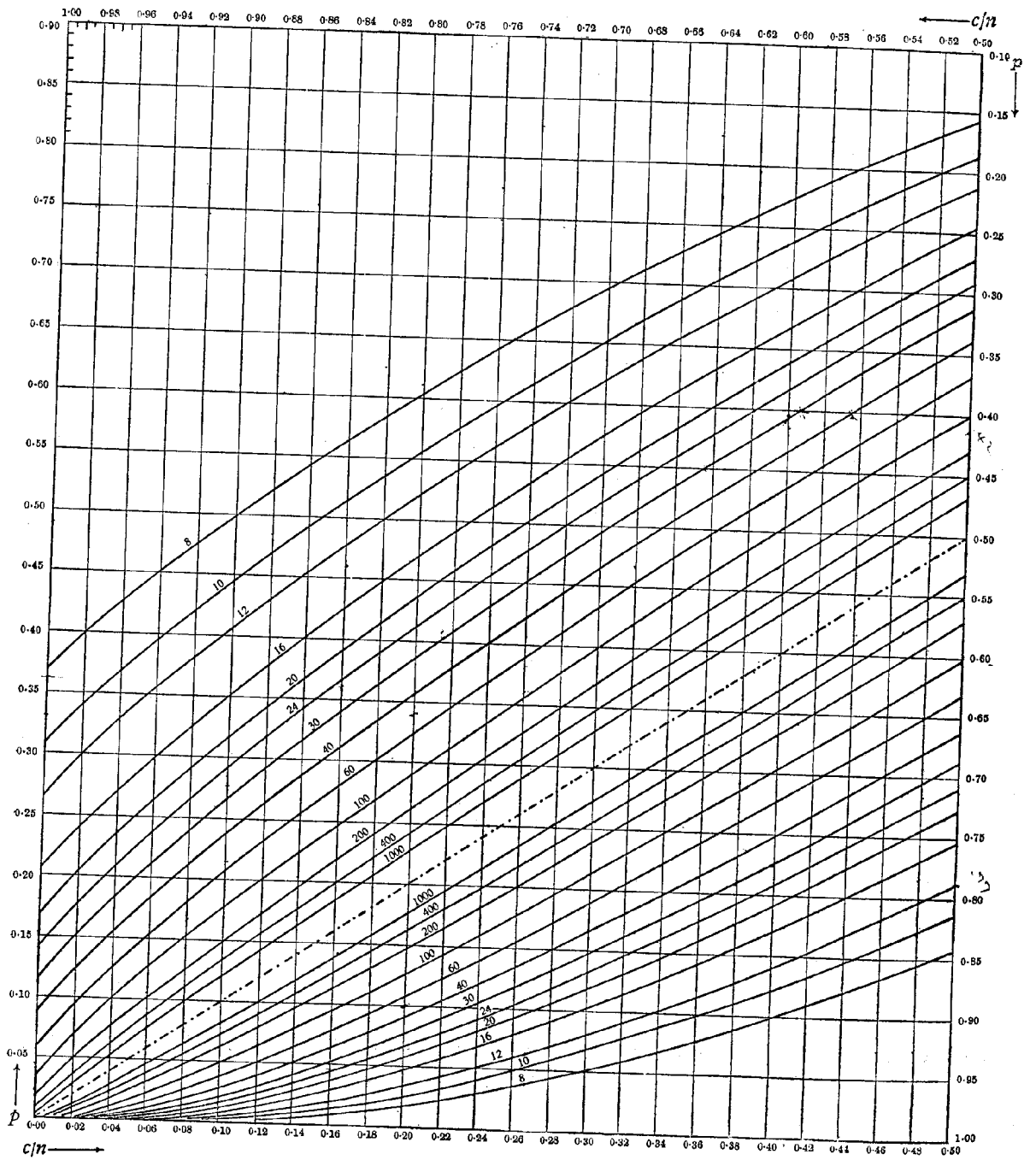


Table 41. Chart providing confidence limits for p in binomial sampling, given a sample fraction c/n .
 Confidence coefficient, $1 - 2\alpha = 0.95$.



The numbers printed along the curves indicate the sample size n . If for a given value of the abscissa c/n , p_A and p_B are the ordinates read from (or interpolated between) the appropriate lower and upper curves, then

$$\Pr \{p_A \leq p \leq p_B\} \geq 1 - 2\alpha.$$

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The Poisson $P(\lambda)$

$$Y = 0, 1, 2, \dots$$

$$\text{Prob}\{Y=y\} = \frac{\lambda^y}{y!} e^{-\lambda}$$

$$E(Y) = \lambda, \quad \text{var}(Y) = \lambda$$

$$P(\lambda_1) + P(\lambda_2) + \dots + P(\lambda_n) \quad \text{independent}$$
$$\sim P(\lambda_1 + \lambda_2 + \dots + \lambda_n)$$

$$B\left(m, \frac{\lambda}{m}\right) \rightarrow P(\lambda) \quad m \text{ large}$$

rare events

$$P(\lambda) \sim N(\lambda, \sqrt{\lambda}) \quad \lambda \text{ large}$$

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Poisson d-sample model.

$$Y_1, \dots, Y_d \sim IP(\lambda_k)$$

Hypothesis of homogeneity

$$H_0: \lambda_1 = \lambda_2 = \dots = \lambda_d$$

Chapter 13

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Boy's born

1334 Swedish ministers

No 0 1 2 3 4 5

Freq 6 57 206 362 365 256

6 7

69 13

Boys born

