Abstract

Relative to the large literature on upper bounds on complexity of convex optimization, lesser attention has been paid to the fundamental hardness of these problems. Given the extensive use of convex optimization in machine learning and statistics, gaining an understanding of these complexity-theoretic issues is important. In this paper, we study the complexity of stochastic convex optimization in an oracle model of computation. We improve upon known results and obtain tight minimax complexity estimates for various function classes.

1 Introduction

Convex optimization forms the backbone of many algorithms for statistical learning and estimation. Given that many statistical estimation problems are large-scale in nature—with the problem dimension and/or sample size being large—it is essential to use bounded computational resources as efficiently as possible. Understanding the computational complexity of convex optimization is thus a key issue for large-scale learning. A large body of literature is devoted to obtaining rates of convergence of specific procedures for various classes of convex optimization problems. A typical outcome of such analysis is an upper bound on the error—for instance, gap to the optimal cost—as a function of the number of iterations. Such analyses have been performed for many standard optimization algorithms, among them gradient descent, mirror descent, interior point programming, and stochastic gradient descent, to name a few. We refer the reader to various standard texts on optimization (e.g., [1, 2, 3]) for further details on such results.

On the other hand, there has been relatively little study of the inherent complexity of convex optimization problems. To the best of our knowledge, the first formal study in this area was undertaken in the seminal work of Nemirovski and Yudin [4], hereafter referred to as NY. One obstacle to a classical complexity-theoretic analysis, as these authors observed, is that of casting convex optimization problems in a Turing Machine model. They avoided this problem by instead considering a natural oracle model of complexity, in which at every round the optimization procedure queries an oracle for certain information on the function being optimized. Working within this framework, the authors obtained a series of lower bounds on the computational complexity of convex optimization problems. In addition to the original text NY [4], we refer the interested reader to the book by Nesterov [5], and the lecture notes by Nemirovski [6] for further background.

In this paper, we consider the computational complexity of stochastic convex optimization within the oracle model. Our results lead to a characterization of the inherent difficulty of learning
and estimation problems when computational resources are constrained. In particular, we improve upon the work of NY [4] in two ways. First, our lower bounds have an improved dependence on the dimension of the space. In the context of statistical estimation, these bounds show how the difficulty of the estimation problem increases with the number of parameters. Second, our techniques naturally extend to give sharper results for optimization over simpler function classes. We show that the complexity of optimization for strongly convex losses is larger than that for convex, Lipschitz losses. Third, we show that for a fixed function class, if the set of optimizers is assumed to have special structure such as sparsity, then the fundamental complexity of optimization can be significantly smaller. All of our proofs exploit a new notion of the discrepancy between two functions that appears to be natural for optimization problems. They involve a reduction from a statistical parameter estimation problem to the stochastic optimization problem, and an application of information-theoretic lower bounds for the estimation problem. We note that special cases of the first two results in this paper appeared in the extended abstract [6], and that a similar study was independently undertaken by Raginsky and Rakhlin [7].

The remainder of this paper is organized as follows. We begin in Section 2 with background on oracle complexity, and a precise formulation of the problems addressed in this paper. Section 3 is devoted to the statement of our main results, and discussion of their consequences. In Section 4, we provide the proofs of our main results, which all exploit a common framework of four steps. More technical aspects of these proofs are deferred to the appendices.

Notation: For the convenience of the reader, we collect here some notation used throughout the paper. For $p \in [1, \infty]$, we use $\|x\|_p$ to denote the $\ell_p$-norm of a vector $x \in \mathbb{R}^p$, and we let $q$ denote the conjugate exponent, satisfying $\frac{1}{p} + \frac{1}{q} = 1$. For two distributions $P$ and $Q$, we use $D(P \| Q)$ to denote the Kullback-Leibler divergence between the distributions. The notation $I(A)$ refers to the 0-1 valued indicator random variable of the set $A$. For two vectors $\alpha, \beta \in \{-1, +1\}^d$, we define the Hamming distance $\Delta_H(\alpha, \beta) := \sum_{i=1}^d I[\alpha_i \neq \beta_i]$.

2 Background and problem formulation

We begin by introducing background on the oracle model of convex optimization, and precisely defining the problem to be studied.

2.1 Convex optimization in the oracle model

Convex optimization is the task of minimizing a convex function $f$ over a convex set $S \subseteq \mathbb{R}^d$. Assuming that the minimum is achieved, it corresponds to computing an element $x^*_f$ that achieves the minimum—that is, $x^*_f \in \arg\min_{x \in S} f(x)$. An optimization method is any procedure that solves this task, typically by repeatedly selecting values from $S$. Our primary focus in this paper is the following question: given any class of convex functions $F$, what is the minimum computational labor any such optimization method would expend for any function in $F$?

In order to address this question, we follow the approach of Nemirovski and Yudin [4], and measure computational labor based on the oracle model of optimization. The main components of this model are an oracle and an information set. An oracle is a (possibly random) function $\phi : S \mapsto I$ that answers any query $x \in S$ by returning an element $\phi(x)$ in an information set $I$. The information set varies depending on the oracle; for instance, for an exact oracle of $m^{th}$ order, the answer to a query $x_t$ consists of $f(x_t)$ and the first $m$ derivatives of $f$ at $x_t$. For the case
of stochastic oracles studied in this paper, these values are corrupted with zero-mean noise with bounded variance. We then measure the computational labor of any optimization method as the number of queries it poses to the oracle.

In particular, given a positive integer $T$ corresponding to the number of iterations, an optimization method $M$ designed to approximately minimize the convex function $f$ over the convex set $S$ proceeds as follows. At any given iteration $t = 1, \ldots, T$, the method $M$ queries at $x_t \in S$, and the oracle reveals the information $\phi(x_t, f)$. The method then uses this information to decide at which point $x_{t+1}$ the next query should be made. For a given oracle function $\phi$, let $M_T$ denote the class of all optimization methods $M$ that make $T$ queries according to the procedure outlined above. For any method $M \in M_T$, we define its error on function $f$ after $T$ steps as

$$
\epsilon_T(M, f, S, \phi) := f(x_T) - \min_{x \in S} f(x) = f(x_T) - f(x^*_T),
$$

where $x_T$ is the method’s query at time $T$. Note that by definition of $x^*_T$ as a minimizing argument, this error is a non-negative quantity.

When the oracle is stochastic, the method’s query $x_T$ at time $T$ is itself random, since it depends on the random answers provided by the oracle. In this case, the optimization error $\epsilon_T(M, f, S, \phi)$ is also a random variable. Accordingly, for the case of stochastic oracles, we measure the accuracy in terms of the expected value $\mathbb{E}_{\phi}[\epsilon_T(M, f, S, \phi)]$, where the expectation is taken over the oracle randomness. Given a class of functions $\mathcal{F}$ defined over a convex set $S$ and a class $M_T$ of all optimization methods based on $T$ oracle queries, we define the minimax error

$$
\epsilon^*_T(\mathcal{F}, S; \phi) := \inf_{M \in M_T} \sup_{f \in \mathcal{F}} \mathbb{E}_{\phi}[\epsilon_T(M, f, S, \phi)].
$$

In the sequel, we provide results for particular classes of oracles. So as to ease the notation, when the oracle $\phi$ is clear from the context, we simply write $\epsilon^*_T(\mathcal{F}, S)$.

### 2.2 Stochastic first-order oracles

In this paper, we study stochastic oracles for which the information set $\mathcal{I} \subset \mathbb{R} \times \mathbb{R}^d$ consists of pairs of noisy function and subgradient evaluations. More precisely, we have:

**Definition 1.** For a given set $S$ and function class $\mathcal{F}$, the class of first-order stochastic oracles consists of mappings $\phi : S \times \mathcal{F} \to \mathcal{I}$ of the form $\phi(x, f) = (\hat{f}(x), \hat{z}(x))$ such that

$$
\mathbb{E}[\hat{f}(x)] = f(x), \quad \mathbb{E}[\hat{z}(x)] \in \partial f(x), \quad \text{and} \quad \mathbb{E}[\|\hat{z}(x)\|_p^2] \leq \sigma^2.
$$

We use $O_{p,\sigma}$ to denote the class of all stochastic first-order oracles with parameters $(p, \sigma)$. Note that the first two conditions imply that $\hat{f}(x)$ is an unbiased estimate of the function value $f(x)$, and that $\hat{z}(x)$ is an unbiased estimate of a subgradient $z \in \partial f(x)$. When $f$ is actually differentiable, then $\hat{z}(x)$ is an unbiased estimate of the gradient $\nabla f(x)$. The third condition in equation (3) controls the “noisiness” of the subgradient estimates in terms of the $\ell_p$-norm.

Stochastic gradient methods are a widely used class of algorithms that can be understood as operating based on information provided by a stochastic first-order oracle. As a particular example, consider a function of the separable form $f(x) = \frac{1}{n} \sum_{i=1}^{n} h_i(x)$, where each $h_i$ is differentiable. Problems of this form arise very frequently in statistical problems, where each term $i$ corresponds to a different sample and the overall cost function is some type of statistical loss (e.g., maximum likelihood, support vector machines, boosting etc.) The natural stochastic gradient method for this
The problem is to choose an index \( i \in \{1, 2, \ldots, n\} \) uniformly at random, and then to return the pair \((h_i(x), \nabla h_i(x))\). Taking averages over the randomly chosen index \( i \) yields \( \frac{1}{n} \sum_{i=1}^{n} h_i(x) = f(x) \), so that \( h_i(x) \) is an unbiased estimate of \( f(x) \), with an analogous unbiased property holding for the gradient of \( h_i(x) \).

2.3 Function classes of interest

We now turn to the classes \( \mathcal{F} \) of convex functions for which we study oracle complexity. In all cases, we consider real-valued convex functions defined over some convex set \( S \). We assume without loss of generality that \( S \) contains an open set around 0, and many of our lower bounds involve the maximum radius \( r = r(S) > 0 \) such that

\[
S \supseteq B_{\infty}(r) := \{ x \in \mathbb{R}^d \mid \|x\|_\infty \leq r \}. \tag{4}
\]

Our first class consists of convex Lipschitz functions:

**Definition 2.** For a given convex set \( S \subseteq \mathbb{R}^d \) and parameter \( p \in [1, \infty] \), the class \( \mathcal{F}_{cv}(S, L, p) \) consists of all convex functions \( f : S \to \mathbb{R} \) such that

\[
|f(x) - f(y)| \leq L \|x - y\|_q \quad \text{for all } x, y \in S, \tag{5}
\]

where \( \frac{1}{q} = 1 - \frac{1}{p} \).

We have defined the Lipschitz condition \( (5) \) in terms of the conjugate exponent \( q \in [1, \infty] \), defined by the relation \( \frac{1}{q} = 1 - \frac{1}{p} \). To be clear, our motivation in doing so is to maintain consistency with our definition of the stochastic first-order oracle, in which we assumed that \( E[\|\hat{z}(x)\|_p^2] \leq \sigma^2 \).

We note that the Lipschitz condition \( (5) \) is equivalent to the condition

\[
\|z\|_p \leq L \quad \forall z \in \partial f(x), \quad \text{and for all } x \in \text{int}(S).
\]

If we consider the case of a differentiable function \( f \), the unbiasedness condition in Definition 1 implies that

\[
\|\nabla f(x)\|_p = \|E[\hat{z}(x)]\|_p \overset{(a)}{\leq} E[\|\hat{z}(x)\|_p] \overset{(b)}{\leq} \sqrt{E[\|\hat{z}(x)\|_p^2]} \leq \sigma,
\]

where inequality (a) follows the convexity of the \( \ell_p \)-norm and Jensen’s inequality, and inequality (b) is a result of Jensen’s inequality applied to the concave function \( \sqrt{x} \). This bound implies that \( f \) must be Lipschitz with constant at most \( \sigma \) with respect to the dual \( \ell_q \)-norm. Therefore, we necessarily must have \( L \leq \sigma \), in order for the function class from Definition 2 to be consistent with the stochastic first-order oracle.

The second function class we consider is that of strongly convex functions.

**Definition 3.** For a given convex set \( S \subseteq \mathbb{R}^d \) and parameter \( p \in [1, \infty] \), the class \( \mathcal{F}_{scv}(S, p; L, \gamma) \) consists of all convex functions \( f : S \to \mathbb{R} \) such that the Lipschitz condition \( (5) \) holds, and such that \( f \) satisfies the \( \ell_2 \)-strong convexity condition

\[
f(\alpha x + (1-\alpha)y) \geq \alpha f(x) + (1-\alpha)f(y) + \alpha(1-\alpha)\gamma^2/2\|x-y\|^2 \quad \text{for all } x, y \in S. \tag{6}
\]

4
In this paper, we restrict our attention to the case of strong convexity with respect to the \( \ell_2 \)-norm. (Similar results on the oracle complexity for strong convexity with respect to different norms can be obtained by straightforward modifications of the arguments given here). For future reference, it should be noted that the Lipschitz constant \( L \) and strong convexity constant \( \gamma \) interact with one another. In particular, whenever \( S \subset \mathbb{R}^d \) contains the \( \ell_\infty \)-ball of radius \( r \), the Lipschitz \( L \) and strong convexity \( \gamma \) constants must satisfy the inequality

\[
\frac{L}{\gamma^2} \geq \frac{r}{8} d^{1/p}.
\]  

(7)

In order to establish this inequality, we note that strong convexity condition with \( \alpha = 1/2 \) implies that

\[
\frac{\gamma^2}{8} \leq \frac{f(x) + f(y) - 2f\left(\frac{x+y}{2}\right)}{2\|x - y\|_2^2} \leq \frac{L\|x - y\|_q}{\|x - y\|_2^2}
\]

We now choose the pair \( x, y \in S \) such that \( \|x - y\|_\infty = r \) and \( \|x - y\|_2 = r\sqrt{d} \). Such a choice is possible whenever \( S \) contains the \( \ell_\infty \) ball of radius \( r \). Since we have \( \|x - y\|_q \leq d^{1/q}\|x - y\|_\infty \), this choice yields \( \frac{\gamma^2}{8} \leq \frac{Ld^{\frac{q}{2}}}{r} \), which establishes the claim (7).

As a third example, we study the oracle complexity of optimization over the class of convex functions that have sparse minimizers. This class of functions is well-motivated, since a large body of statistical work has studied the estimation of vectors, matrices and functions under various types of sparsity constraints. A common theme in this line of work is that the ambient dimension \( d \) enters only logarithmically, and so has a mild effect. Consequently, it is natural to investigate whether the complexity of optimization methods also enjoys such a mild dependence on ambient dimension under sparsity assumptions.

For a vector \( x \in \mathbb{R}^d \), we use \( \|x\|_0 \) to denote the number of non-zero elements in \( x \). Recalling the set \( \mathcal{F}_{cv}(S, L, p) \) from Definition 2, we now define a class of Lipschitz functions with sparse minimizers.

**Definition 4.** For a convex set \( S \subset \mathbb{R}^d \) and positive integer \( k \leq \lfloor d/2 \rfloor \), we define

\[
\mathcal{F}_{sp}(k; S, L) := \{ f \in \mathcal{F}_{cv}(S, L, \infty) \mid \exists \ x^* \in \text{arg min}_{x \in S} f(x) \text{ satisfying } \|x^*\|_0 \leq k. \}
\]

(8)

We frequently use the shorthand notation \( \mathcal{F}_{sp}(k) \) when the set \( S \) and parameter \( L \) are clear from context. In words, the set \( \mathcal{F}_{sp}(k) \) consists of all convex functions that are \( L \)-Lipschitz in the \( \ell_\infty \)-norm, and have at least one \( k \)-sparse optimizer.

## 3 Main results and their consequences

With the setup of stochastic convex optimization in place, we are now in a position to state the main results of this paper, and to discuss some of their consequences. As previously mentioned, a subset of our results assume that the set \( S \) contains an \( \ell_\infty \) ball of radius \( r = r(S) \). Our bounds scale with \( r \), thereby reflecting the natural dependence on the size of the set \( S \). Also, we set the oracle variance bound \( \sigma \) to be the same as the Lipschitz constant \( L \) in our results.
3.1 Oracle complexity for convex Lipschitz functions

We begin by analyzing the minimax oracle complexity of optimization for the class of bounded and convex Lipschitz functions $F_{cv}$ from Definition 2.

**Theorem 1.** Let $S \subset \mathbb{R}^d$ be a convex set such that $S \supseteq B_\infty(r)$ for some $r > 0$. Then for a universal constant $c > 0$, the minimax oracle complexity over the class $F_{cv}(S, L, p)$ satisfies the following lower bounds:

(a) For $1 \leq p \leq 2$,

$$\sup_{\phi \in \mathcal{O}_{p, L}} \epsilon^*_T(F_{cv}, S; \phi) \geq \min \left\{ cLr \sqrt{dT}, Lr \right\}.$$  \hfill (9)

(b) For $p > 2$,

$$\sup_{\phi \in \mathcal{O}_{p, L}} \epsilon^*_T(F_{cv}, S; \phi) \geq \left\{ cLr \frac{d^{1-\frac{1}{p}}}{\sqrt{T}}, Ld^{1-\frac{1}{p}}r \right\}.$$  \hfill (10)

**Remarks:** Nemirovski and Yudin [4] proved the lower bound $\Omega\left(\frac{1}{\sqrt{T}}\right)$ for the class $F_{cv}$, where the set $S$ is the unit ball of a norm, and the functions are Lipschitz in the dual norm. By appealing to a matching upper bound for the method of mirror descent, they show minimax optimality of their dimension-independent result in this primal-dual norm setting when $p \geq 2$. In contrast, here we impose an $\ell_\infty$ constraint on the set since optimization under upper and lower bound constraints on variables is quite common and not sufficiently addressed by prior work. In addition, by looking at the relative geometry of an arbitrary set with respect to the $\ell_\infty$ ball, we can leverage the results to arbitrary sets as demonstrated by the theorem obtaining dimension-dependent lower bounds. Obtaining this correct dependence highlights the role of the geometry of the set $S$ in determining the oracle complexity.

In general, our lower bounds cannot be improved, and hence specify the optimal minimax oracle complexity. We consider here some examples to illustrate their sharpness. Throughout we assume that $T$ is large enough to ensure that the $1/\sqrt{T}$ term attains the lower bound and not the $L/144$ term, which is reasonable since we want to understand the rate of the optimization procedure and not in the transient behavior over the first few iterations.

(a) Let us begin by considering the special case $S = B_\infty(1)$. For this choice, we have $r(S) = 1$, and hence Theorem 1 implies that for all $p \in [1, 2]$, the minimax oracle complexity is lower bounded as

$$\sup_{\phi \in \mathcal{O}_{p, L}} \epsilon^*_T(F_{cv}, B_\infty(1); \phi) = \Omega\left( L \sqrt{\frac{1}{T}} \right).$$

Up to constant factors, this lower bound is sharp for all $p \in [1, 2]$. Indeed, for any convex set $S$, stochastic gradient descent achieves a matching upper bound (see Section 5.2.4, p. 196 of NY [4], as well as Appendix C in this paper for further discussion).

(b) As another example, suppose that $S = B_2(1)$. Observe that this $\ell_2$-norm unit ball satisfies the relation $B_2(1) \supseteq \frac{1}{\sqrt{d}} B_\infty(1)$, so that we have $r(B_2(1)) = 1/\sqrt{d}$. Consequently, for this choice,
the lower bound $\Omega(2)$ takes the form

$$
\sup_{\phi \in \mathcal{O}_{p,L}} \epsilon^*_T(\mathcal{F}_{cv}, \mathbb{B}_2(1); \phi) = \Omega\left(L \frac{1}{\sqrt{T}}\right),
$$

which is a dimension-independent lower bound. This lower bound for $\mathbb{B}_2(1)$ is indeed tight for $p \in [1, 2]$, and as before, this rate is achieved by stochastic gradient descent [4].

(c) Turning to the case of $p > 2$, the lower bound $\Omega(10)$ is matched up to constants by the method of mirror descent using the dual norm $\| \cdot \|_q$; for further discussion, we again refer the reader to Section 5.2.1, p. 190 of NY [4], as well as to Appendix C of this paper. Also, even though this lower bound requires the oracle to have only bounded variance, our proof actually uses a stochastic oracle based on Bernoulli random variables, for which all moments exist. Consequently, at least in general, our results show that there is no hope of achieving faster rates by restricting to oracles with bounds on higher-order moments. This is an interesting contrast to the case of having less than two moments, in which the rates are slower. For instance, as shown in Section 5.3.1 of NY [4], suppose that the gradient estimates in a stochastic oracle satisfy the moment bound $E\|\tilde{z}(x)\|^b \leq \sigma^2$ for some $b \in [1, 2)$. In this setting, the oracle complexity is lower bounded by $\Omega(T^{-b/(b-1)}L^{b-1})$. Since $T^{b-1} \ll T^{1/2}$ for all $b \in [1, 2)$, there is a significant penalty in convergence rates for having less than two bounded moments.

(d) Even though the results have been stated in a first-order stochastic oracle model, they actually hold in a stronger sense. Indeed, they apply to an infinite-order oracle that responds with a random function $f_t$ such that

$$
E[f_t(x)] = E[f(x)], \quad \text{and} \quad E[\nabla^i f_t(x)] \leq \nabla^i f(x) \quad \text{for all } x \in \mathbb{S} \text{ and } i \text{ such that } \nabla^i f(x) \text{ exists},
$$

along with appropriately bounded second moments of all the derivatives. Consequently, higher-order gradient information cannot improve convergence rates in a worst-case setting. The result is close in spirit to a statistical sample complexity lower bound, since the lower bounds essentially allow the optimization algorithm to see the outcome of the random bits generated by the oracle at any given round. As we will see, the proof techniques are also similar to those employed in statistics, particularly in the areas of non-parametric regression [8, 9] as well as in learning theory [10, 11, 12]. However for purposes of optimization, the metric of interest in which we would like to approximate the functions is very different from classical statistics literature.

### 3.2 Oracle complexity for strongly convex Lipschitz functions

We now turn to the statement of lower bounds over the class of Lipschitz and strongly convex functions $\mathcal{F}_{scv}$ from Definition 3. In all these statements, we assume that $\gamma^2 \leq \frac{8Ld^{-1/p}}{\epsilon}$, as is required for the definition of $\mathcal{F}_{scv}$ to be sensible.

**Theorem 2.** Let $\mathbb{S} = \mathbb{B}_\infty(r)$. Then there exist universal constants $c_1, c_2 > 0$ such that the minimax oracle complexity over the class $\mathcal{F}_{scv}(\mathbb{S}, p; L, \gamma)$ satisfies the following lower bounds:

(a) For $p = 1$, we have

$$
\sup_{\phi \in \mathcal{O}_{p,L}} \epsilon^*_T(\mathcal{F}_{scv}, \phi) \geq \min \left\{ c_1 \frac{L^2}{\gamma^2T}, c_2 Lr \sqrt{\frac{d}{T}}, \frac{L^2}{288\gamma^2d}, \frac{Lr}{144} \right\}.
$$

(11)
(b) For $p > 2$, we have:
\[
\sup_{\phi \in \mathcal{O}_{p,L}} \epsilon^*(\mathcal{F}_{scv}, \phi) \geq \min \left( c_1 \frac{L^2 d^{1-2/p}}{\gamma^2 T}, \frac{c_2 L r^d}{\sqrt{T}}, \frac{L^2 d^{1-2/p}}{288 \gamma^2}, \frac{L r^d}{144} \right).
\] (12)

The results demonstrate an interesting phase transition and have a strongly convex and a convex regime. When the strong convexity parameter $\gamma^2$ is large (and $T$ is sufficiently large), the first term (namely, $\Omega(1/T)$) always dominates. However, for poorly conditioned objectives with $\gamma \approx 0$, the convex rate of $\Omega(1/\sqrt{T})$ is faster than the strongly convex rate. This is natural since it recovers the convex result when $\gamma = c$ as a special case like we expect. As with Theorem 1, these lower bounds are sharp. In particular, for $S = \mathbb{B}_\infty(1)$, stochastic gradient descent achieves the rate $\Omega(1)$ up to logarithmic factors [13], and a closely related algorithm proposed in very recent work [14] matches the lower bound exactly. Again for the purposes of upper bounds, we observe the same phase transition with the best of the strongly convex or the convex rates attainable once the conditioning of the objective is accounted for.

The above theorem applies only to $\mathbb{B}_\infty(r)$ and not arbitrary sets $S$ unlike the previous theorem. The generalization of this result to arbitrary convex, compact sets is an interesting question for open research.

3.3 Oracle complexity for convex Lipschitz functions with sparse optima

Finally, we turn to the oracle complexity of optimization over the class $\mathcal{F}_{sp}$ from Definition 4.

**Theorem 3.** Let $\mathcal{F}_{sp}$ be the class of all convex functions that are $L$-Lipschitz with respect to the $\| \cdot \|_\infty$ norm that have a $k$-sparse optimizer. Let $S \subset \mathbb{R}^d$ be a convex set with $\mathbb{B}_\infty(r) \subseteq S$. Then there exists a universal constant $c > 0$ such that for all $k \leq \lfloor \frac{d}{2} \rfloor$, we have
\[
\sup_{\phi \in \mathcal{O}_{\infty,L}} \epsilon^*(\mathcal{F}_{sp}, \phi) \geq \min \left( c L r \sqrt{\frac{k^2 \log \frac{d}{k}}{T}}, \frac{L kr}{432} \right).
\] (13)

**Remark:** If $k = \mathcal{O}(d^{1-\delta})$ for some $\delta \in (0, 1)$ (so that $\log \frac{d}{k} = \Theta(\log d)$), then this bound is sharp up to constant factors. In particular, suppose that we use mirror descent based on the $\| \cdot \|_{1+\epsilon}$ norm with $\epsilon = 2 \log d / (2 \log d - 1)$. As we discuss in more detail in Appendix C, it can be shown that this technique will achieve a solution accurate to $\mathcal{O}(\sqrt{\frac{k^2 \log d}{T}})$ within $T$ iterations, which matches our lower bound [13] up to constant factors whenever $k = \mathcal{O}(d^{1-\delta})$. To the best of our knowledge, Theorem 3 provides the first tight lower bound on the oracle complexity of sparse optimization.

4 Proofs of results

We now turn to the proofs of our main results. We begin in Section 4.1 by outlining the framework of basic results on which our proofs are based. Sections 4.2 through 4.4 are devoted to the proofs of Theorems 1 through 3 respectively.

4.1 Basic results

We begin by establishing a basic set of results that are exploited in the proofs of the main results. At a high-level, our main idea is to show that the problem of convex optimization is at least as
hard as estimating the parameters of Bernoulli variables—that is, the biases of $d$ independent coins. In order to perform this embedding, for a given error tolerance $\epsilon$, we start with an appropriately chosen subset of the vertices of a $d$-dimensional hypercube, each of which corresponds to some values of the $d$ Bernoulli parameters. For a given function class, we then construct a “difficult” subclass of functions that are indexed by these vertices of the hypercube. We then show that being able to optimize any function in this subclass to $\epsilon$-accuracy requires identifying the hypercube vertex. This is a multiway hypothesis test based on the observations provided by $T$ queries to the stochastic oracle, and we apply the Fano inequality [15] to lower bound the probability of error. In the remainder of this section, we provide more detail on each of steps involved in this embedding.

### 4.1.1 Constructing a difficult subclass of functions

Our first step is to construct a subclass of functions $G \subseteq F$ that we use to derive lower bounds. Any such subclass is parameterized by a subset $V \subseteq \{-1, +1\}^d$ of the hypercube, chosen as follows. Recalling that $\Delta_H$ denotes the Hamming metric, we let $V = \{\alpha_1, \ldots, \alpha_M\}$ be a subset of the vertices of the hypercube such that

$$\Delta_H(\alpha_i, \alpha_j) \geq \frac{d}{4}$$

for all $i \neq j$, meaning that $V$ is a $\frac{d}{4}$-packing in the Hamming norm. It is a classical fact (e.g., [18]) that one can construct such a set with cardinality $|V| \geq \left(\frac{2}{\sqrt{e}}\right)^{d/2}$.

Now let $G_{\text{base}} = \{f^+_i, f^-_i, i = 1, \ldots, d\}$ denote some base set of $2d$ functions defined on the convex set $S$, to be chosen appropriately depending on the problem at hand. For a given tolerance $\delta \in (0, \frac{1}{4}]$, we define, for each vertex $\alpha \in V$, the function

$$g_\alpha(x) := \frac{c}{d} \sum_{i=1}^{d} \left\{ (1/2 + \alpha_i \delta) f^+_i(x) + (1/2 - \alpha_i \delta) f^-_i(x) \right\}.$$  

Depending on the result to be proven, our choice of the base functions $\{f^+_i, f^-_i\}$ and the pre-factor $c$ will ensure that each $g_\alpha$ satisfies the appropriate Lipschitz and/or strong convexity properties over $S$. Moreover, we will ensure that all minimizers $x_\alpha$ of each $g_\alpha$ are contained within the set $S$.

Based on these functions and the packing set $V$, we define the function class

$$G(\delta) := \{g_\alpha, \alpha \in V\}.$$  

Note that $G(\delta)$ contains a total of $|V|$ functions by construction, and as mentioned previously, our choices of the base functions etc. will ensure that $G(\delta) \subseteq F$. We demonstrate specific choices of the class $G(\delta)$ in the proofs of Theorems 1 and 2 to follow.

### 4.1.2 Optimizing well is equivalent to function identification

We now claim that if a method can optimize over the subclass $G(\delta)$ up to a certain tolerance, then it must be capable of identifying which function $g_\alpha \in G(\delta)$ was chosen. We first require a measure for the closeness of functions in terms of their behavior near each others’ minima. Recall that we use $x^*_f \in \mathbb{R}^d$ to denote a minimizing point of the function $f$. Given a convex set $S \subseteq \mathbb{R}^d$ and two functions $f, g$, we define

$$\rho(f, g) := \inf_{x \in S} \left[ f(x) + g(x) - f(x^*_f) - g(x^*_g) \right].$$

9
The discrepancy measure is non-negative, symmetric in its arguments, and satisfies $\rho(f, g) = 0$ if and only if $x_f^* = x_g^*$, so that we may refer to it as a premetric. (It does not satisfy the triangle inequality and $\rho(f, g)$ if and only if $f = g$ required to be a metric.)

Given the subclass $G(\delta)$, we quantify how densely it is packed with respect to the pre metric $\rho$ using the quantity

$$
\psi(G(\delta)) := \min_{\alpha \neq \beta \in V} \rho(g_\alpha, g_\beta).
$$

(18)

We denote this quantity by $\psi(\delta)$ when the class $G$ is clear from the context. We now state a simple result that demonstrates the utility of maintaining a separation under $\rho$ among functions in $G(\delta)$.

**Lemma 1.** For any $\bar{x} \in S$, there can be at most one function $g_\alpha \in G(\delta)$ such that

$$
g_\alpha(\bar{x}) - \inf_{x \in S} g_\alpha(x) \leq \frac{\psi(\delta)}{3}.
$$

(19)

Thus, if we have an element $\bar{x} \in S$ that approximately minimizes one function in the set $G(\delta)$ up to tolerance $\psi(\delta)$, then it cannot approximately minimize any other function in the set.

**Proof.** For a given $\bar{x} \in S$, suppose that there exists an $\alpha \in V$ such that $g_\alpha(\bar{x}) - g_\alpha(x_\alpha^*) \leq \frac{\psi(\delta)}{3}$.

From the definition of $\psi(\delta)$ in (18), for any $\beta \in V$, $\beta \neq \alpha$, we have

$$
\psi(\delta) \leq g_\alpha(\bar{x}) - \inf_{x \in S} g_\alpha(x) + g_\beta(\bar{x}) - \inf_{x \in S} g_\beta(x) \leq \frac{\psi(\delta)}{3} + g_\beta(\bar{x}) - \inf_{x \in S} g_\beta(x).
$$

Re-arranging yields the inequality $g_\beta(\bar{x}) - g_\beta(x_\beta^*) \geq \frac{2}{3} \psi(\delta)$, from which the claim (19) follows.

Suppose that for some fixed but unknown function $g_\alpha^* \in G(\delta)$, some method $M_T$ is allowed to make $T$ queries to an oracle with information function $\phi(\cdot; g_\alpha^*)$, thereby obtaining the information sequence

$$
\phi(x_1^T; g_\alpha^*) := \{\phi(x_t; g_\alpha^*), t = 1, 2, \ldots, T\}.
$$
Our next lemma shows that if the method \( \mathcal{M}_T \) achieves a low minimax error over the class \( \mathcal{G}(\delta) \), then one can use its output to construct a hypothesis test that returns the true parameter \( \alpha^* \) at least \( 2/3 \) of the time. (In this statement, we recall the definition (2) of the minimax error in optimization.)

**Lemma 2.** Suppose that based on the data \( \phi(x_1^T; g_\alpha^*) \), there exists a method \( \mathcal{M}_T \) that achieves a minimax error satisfying

\[
\mathbb{E}[\epsilon_T(\mathcal{M}_T, \mathcal{G}(\delta), S, \phi)] \leq \frac{\psi(\delta)}{9}. \tag{20}
\]

Based on such a method \( \mathcal{M}_T \), one can construct a hypothesis test \( \hat{\alpha} : \phi(x_1^T; g_\alpha^*) \rightarrow \mathcal{V} \) such that

\[
\max_{\alpha^* \in \mathcal{V}} \mathbb{P}_\phi[\hat{\alpha} \neq \alpha^*] \leq \frac{1}{3}.
\]

**Proof.** Given a method \( \mathcal{M}_T \) that satisfies the bound (20), we construct an estimator \( \hat{\alpha}(\mathcal{M}_T) \) of the true vertex \( \alpha^* \) as follows. If there exists some \( \alpha \in \mathcal{V} \) such that \( g_\alpha(x_T) - g_\alpha(x_\alpha) \leq \frac{\psi(\delta)}{3} \) then we set \( \hat{\alpha}(\mathcal{M}_T) \) equal to \( \alpha \). If no such \( \alpha \) exists, then we choose \( \hat{\alpha}(\mathcal{M}_T) \) uniformly at random from \( \mathcal{V} \). From Lemma 1 there can exist only one such \( \alpha \in \mathcal{V} \) that satisfies this inequality. Consequently, using Markov’s inequality, we have

\[
\mathbb{P}_\phi[\hat{\alpha}(\mathcal{M}_T) \neq \alpha^*] \leq \mathbb{P}_\phi[\epsilon_T(\mathcal{M}_T, g_\alpha^*, S, \phi) \geq \psi(\delta)/3] \leq \frac{1}{3}.
\]

Maximizing over \( \alpha^* \) completes the proof. 

We have thus shown that having a low minimax optimization error over \( \mathcal{G}(\delta) \) implies that the vertex \( \alpha^* \in \mathcal{V} \) can be identified most of the time.

### 4.1.3 Oracle answers and coin tosses

We now describe stochastic first order oracles \( \phi \) for which the samples \( \phi(x_1^T; g_\alpha) \) can be related to coin tosses. In particular, we associate a coin with each dimension \( i \in \{1, 2, \ldots, d\} \), and consider the set of coin bias vectors lying in the set

\[
\Theta(\delta) = \{ (1/2 + \alpha_1 \delta, \ldots, 1/2 + \alpha_d \delta) \mid \alpha \in \mathcal{V} \}, \quad \tag{21}
\]

Given a particular function \( g_\alpha \in \mathcal{G}(\delta) \)—or equivalently, vertex \( \alpha \in \mathcal{V} \)—we consider two oracles \( \phi \) that present noisy value and gradient samples from \( g_\alpha \), as summarized in the following.

**Oracle A: 1-dimensional unbiased gradients**

(a) Pick an index \( i \in \{1, \ldots, d\} \) uniformly at random.

(b) Draw \( b_i \in \{0, 1\} \) according to a Bernoulli distribution with parameter \( 1/2 + \alpha_i \delta \).

(c) For the given input \( x \in S \), return the value \( \tilde{g}_{\alpha,A}(x) \) and a sub-gradient \( \tilde{z}_{\alpha,A}(x) \in \partial g_{\alpha,A}(x) \) of the function

\[
\tilde{g}_{\alpha,A} := c \left[ b_i f_i^+ + (1 - b_i) f_i^- \right].
\]

By construction, the function value and gradients returned by Oracle A are unbiased estimates of those of \( g_\alpha \). In particular, since each co-ordinate \( i \) is chosen with probability \( 1/d \), we have

\[
\mathbb{E}[\tilde{g}_{\alpha,A}(x)] = \frac{c}{d} \sum_{i=1}^d \left[ b_i f_i^+(x) + (1 - b_i) f_i^-(x) \right] = g_\alpha(x),
\]

11
with a similar relation for the gradient. Furthermore, as long as the base functions \( f_i^+ \) and \( f_i^- \) have gradients bounded by 1, we have \( \mathbb{E}[\|\hat{z}_{A}(x)\|_p] \leq c \) for all \( p \in [1, \infty] \).

Parts of proofs are based on an oracle which responds with function values and gradients that are \( d \)-dimensional in nature.

**Oracle B: \( d \)-dimensional unbiased gradients**

(a) For \( i = 1, \ldots, d \), draw \( b_i \in \{0, 1\} \) according to a Bernoulli distribution with parameter \( 1/2 + \alpha_i \delta \).

(b) For the given input \( x \in \mathbb{S} \), return the value \( \hat{g}_{\alpha,B}(x) \) and a sub-gradient \( \hat{z}_{\alpha,B}(x) \in \partial \hat{g}_{\alpha,B}(x) \) of the function

\[
\hat{g}_{\alpha,B} := \frac{c}{d} \sum_{i=1}^{d} [b_i f_i^+ + (1 - b_i) f_i^-].
\]

As with Oracle A, this oracle returns unbiased estimates of the function values and gradients. We also note that we will be working with constructions where the functions \( f_i^+ \), \( f_i^- \) only depend on the \( i \)-th coordinate \( x(i) \). In such cases, if we assume that \( |\frac{\partial f_i^+}{\partial x(i)}| \leq 1 \) and similarly for \( f_i^- \), we have

\[
\|\nabla \hat{g}_{\alpha,B}(x)\|_p^2 = \frac{c^2}{d^2} \left( \sum_{i=1}^{d} \left| b_i \frac{\partial f_i^+}{\partial x(i)} + (1 - b_i) \frac{\partial f_i^-}{\partial x(i)} \right|^p \right)^{2/p} \leq c^2 d^{2/p-2}. \tag{22}
\]

In our later uses of Oracles A and B, we will choose the pre-factor \( c \) appropriately so as to produce the desired Lipschitz constants.

### 4.1.4 Lower bounds on coin-tossing

Finally, we use information-theoretic methods to lower bound the probability of correctly estimating the true parameter \( \alpha^* \in V \) in our model. At each round of either Oracle A or Oracle B, we can consider a set of \( d \) coin tosses, with an associated vector \( \theta^* = (\frac{1}{2} + \alpha_1^* \delta, \ldots, \frac{1}{2} + \alpha_d^* \delta) \) of parameters. At any round, the output of Oracle A can (at most) reveal the instantiation \( b_i \in \{0, 1\} \) of a randomly chosen index, whereas Oracle B can at most reveal the entire vector \( (b_1, b_2, \ldots, b_d) \). Our goal is to lower bound the probability of estimating the true parameter \( \alpha^* \), based on a sequence of length \( T \). We note that this part of our proof closely resembles statistical sample complexity proofs as observed after Theorem 1.

**Lemma 3.** Suppose that the Bernoulli parameter vector \( \alpha^* \) is chosen uniformly at random from the packing set \( V \), and suppose that the outcome of \( \ell \leq d \) coins chosen uniformly at random is revealed at each round \( t = 1, \ldots, T \). Then for any \( \delta \in (0, 1/4] \), any hypothesis test \( \hat{\alpha} \) satisfies

\[
\mathbb{P}[\hat{\alpha} \neq \alpha^*] \geq 1 - \frac{16\ell T \delta^2 + \log 2}{\frac{1}{2} \log(2/\sqrt{\delta})}, \tag{23}
\]

where the probability is taken over both randomness in the oracle and the choice of \( \alpha^* \).
Note that we will apply the lower bound (23) with \( \ell = 1 \) in the case of Oracle A, and \( \ell = d \) in the case of Oracle B.

Proof. For each time \( t = 1, 2, \ldots, T \), let \( U_t \) denote the randomly chosen subset of size \( \ell \), \( X_{t,i} \) be the outcome of oracle’s coin toss at time \( t \) for coordinate \( i \) and let \( Y_t \in \{-1,0,1\}^d \) be a random vector with entries

\[
Y_{t,i} = \begin{cases} 
X_{t,i} & \text{if } i \in U_t, \\
-1 & \text{if } i \notin U_t.
\end{cases}
\]

By Fano’s inequality \( [15] \), we have the lower bound

\[
\mathbb{P}[\hat{\alpha} \neq \alpha^*] \geq 1 - \frac{I\left(\{U_t, Y_t\}_{t=1}^T; \alpha^*\} + \log 2}{\log |\mathcal{V}|},
\]

where \( I\left(\{U_t, Y_t\}_{t=1}^T; \alpha^*\) denotes the mutual information between the sequence \( \{U_t, Y_t\}_{t=1}^T \) and the random parameter vector \( \alpha^* \). As discussed earlier, we are guaranteed that \( \log |\mathcal{V}| \geq \frac{d}{2} \log (2/\sqrt{e}) \). Consequently, in order to prove the lower bound (23), it suffices to establish the upper bound \( I\left(\{U_t, Y_t\}_{t=1}^T; \alpha^*\} \leq 16T \ell \delta^2 \).

By the independent and identically distributed nature of the sampling model, we have

\[
I\left(\{U_1, Y_1\}, \ldots, \{U_T, Y_T\}; \alpha^*\) = \sum_{t=1}^T I\left(U_t, Y_t; \alpha^*\) = T I\left(U_1, Y_1; \alpha^*\),
\]

so that it suffices to upper bound the mutual information for a single round. To simplify notation, from here onwards we write \((Y, U)\) to mean the pair \((Y_1, U_1)\); the remainder of our proof is devoted to establishing that \( I(Y; U) \leq 16 \ell \delta^2 \).

By chain rule for mutual information \( [15] \), we have

\[
I\left(U, Y; \alpha^*\) = I\left(Y; \alpha^* \mid U\right) + I\left(\alpha^*; U\right).
\]

Since the subset \( U \) is chosen independently of \( \alpha^* \), we have \( I\left(\alpha^*; U\right) = 0 \), and so it suffices to upper bound the first term. By definition of conditional mutual information \( [15] \), we have

\[
I\left(Y; \alpha^* \mid U\right) = \mathbb{E}_{U} \left[D\left(\mathbb{P}_{Y \mid \alpha^*, U} \mid \mathbb{P}_{Y \mid U}\right)\right]
\]

Since \( \alpha \) has a uniform distribution over \( \mathcal{V} \), we have \( \mathbb{P}_{Y \mid U} = \frac{1}{|\mathcal{V}|} \sum_{\alpha \in \mathcal{V}} \mathbb{P}_{Y \mid \alpha, U} \), so that by convexity of the Kullback-Leibler divergence,

\[
D\left(\mathbb{P}_{Y \mid \alpha^*, U} \mid \mathbb{P}_{Y \mid U}\right) \leq \frac{1}{|\mathcal{V}|} \sum_{\alpha \in \mathcal{V}} D\left(\mathbb{P}_{Y \mid \alpha^*, U} \mid \mathbb{P}_{Y \mid \alpha, U}\right).
\]

Now for any pair \( \alpha^*, \alpha \in \mathcal{V} \), the KL divergence \( D\left(\mathbb{P}_{Y \mid \alpha^*, U} \mid \mathbb{P}_{Y \mid \alpha, U}\right) \) can be at most the KL divergence between \( \ell \) independent pairs of Bernoulli variates with parameters \( \frac{1}{2} + \delta \) and \( \frac{1}{2} - \delta \). Letting \( D(\delta) \) denote the Kullback-Leibler divergence between a single Bernoulli pair, a little calculation yields

\[
D(\delta) = \left(\frac{1}{2} + \delta\right) \log \left(\frac{1}{2} + \delta\right) + \left(\frac{1}{2} - \delta\right) \log \left(\frac{1}{2} - \delta\right)
\]

\[= 2\delta \log \left(1 + \frac{4\delta}{1 - 2\delta}\right) \leq \frac{8\delta^2}{1 - 2\delta}.
\]
Consequently, as long as \( \delta \leq 1/4 \), we have \( D(\delta) \leq 16\delta^2 \). Returning to the bound (24), we conclude that \( D(P_{Y|\alpha^*U} \| P_{Y|U}) \leq 16\delta^2 \). Taking averages over \( U \), we obtain the bound \( I(Y; \alpha^* | U) \leq 16 \ell \delta^2 \), and applying the decomposition (24) yields \( I((U, Y); \alpha^*) \leq 16\ell \delta^2 \), thereby completing the proof. 

Equipped with these tools, we are now prepared to prove our main results.

### 4.2 Proof of Theorem [1]

We begin with oracle complexity for bounded Lipschitz functions, as stated in Theorem [1]. We first prove the result for the set \( S = B_\infty(\frac{1}{2}) \).

**Part (a)—Proof for \( p \in [1, 2] \):** Consider Oracle A that returns the quantities \((\hat{g}_\alpha, A(x)), \hat{z}_\alpha, A(x))\). By definition of the oracle, each round reveals only at most one coin flip, meaning that we can apply Lemma [3] with \( \ell = 1 \), thereby obtaining the lower bound

\[
\Pr[\alpha(M_T) \neq \alpha] \geq 1 - 2\frac{16T \delta^2 + \log 2}{d \log(2/\epsilon)}. \tag{26}
\]

We now seek an upper bound \( \Pr[\alpha(M_T) \neq \alpha] \) using Lemma [2]. In order to do so, we need to specify the base functions \((f_i^+, f_i^-)\) involved. For \( i = 1, \ldots, d \), we define

\[
f_i^+(x) := |x(i) + \frac{1}{2}|, \quad \text{and} \quad f_i^-(x) := |x(i) - \frac{1}{2}|. \tag{27}
\]

Given that \( S = B_\infty(\frac{1}{2}) \), we see that the minimizers of \( g_\alpha \) are contained in \( S \). Also, both the functions are 1-Lipschitz in the \( \ell_1 \)-norm. By the construction (15), we are guaranteed that for any subgradient of \( g_\alpha \), we have

\[
\|z_\alpha(x)\|_p \leq 2c \quad \text{for all} \ p \geq 1.
\]

Therefore, in order to ensure that \( g_\alpha \) is \( L \)-Lipschitz in the dual \( \ell_q \)-norm, it suffices to set \( c = L/2 \).

Let us now lower bound the discrepancy function (17). We first observe that each function \( g_\alpha \) is minimized over the set \( B_\infty(\frac{1}{2}) \) at the vector \( x_\alpha := -\alpha/2 \), at which point it achieves its minimum value

\[
\min_{x \in B_\infty(\frac{1}{2})} g_\alpha(x) = \frac{c}{2} - c\delta.
\]

Furthermore, we note that for any \( \alpha \neq \beta \), we have

\[
g_\alpha(x) + g_\beta(x) = \frac{c}{d} \sum_{i=1}^{d} \left[ (\frac{1}{2} + \alpha_i + \frac{1}{2} + \beta_i) f_i^+(x) + (\frac{1}{2} - \alpha_i + \frac{1}{2} - \beta_i) f_i^-(x) \right]
\]

\[
= \frac{c}{d} \sum_{i=1}^{d} \left[ (1 + \alpha_i + \beta_i) f_i^+(x) + (1 - \alpha_i - \beta_i) f_i^-(x) \right]
\]

\[
= \frac{c}{d} \sum_{i=1}^{d} \left[ (f_i^+(x) + f_i^-(x)) \mathbb{I}(\alpha_i \neq \beta_i) + ((1 + 2\alpha_i) f_i^+(x) + (1 - 2\alpha_i) f_i^-(x)) \mathbb{I}(\alpha_i = \beta_i) \right].
\]

When \( \alpha_i = \beta_i \) then \( x_\alpha(i) = x_\beta(i) = -\alpha_i/2 \), so that this co-ordinate does not make a contribution to the discrepancy function \( \rho(g_\alpha, g_\beta) \). On the other hand, when \( \alpha_i \neq \beta_i \), we have

\[
f_i^+(x) + f_i^-(x) = |x(i) + \frac{1}{2}| + |x(i) - \frac{1}{2}| \geq 1 \quad \text{for all} \ x \in \mathbb{R}.
\]
Consequently, any such co-ordinate yields a contribution of $2c\delta/d$ to the discrepancy. Recalling our packing set (14) with $d/4$ separation in Hamming norm, we conclude that for any distinct $\alpha \neq \beta$ within our packing set,

$$\rho(g_\alpha, g_\beta) = \frac{2c\delta}{d} \Delta_H(\alpha, \beta) \geq \frac{c\delta}{2},$$

so that by definition of $\psi$, we have established the lower bound $\psi(\delta) \geq \frac{c\delta}{2}$. 

Setting the target error $\epsilon := \frac{c\delta}{18}$, we observe that this choice ensures that $\epsilon < \frac{\psi(\delta)}{9}$. Recalling the requirement $\delta < 1/4$, we have $\epsilon < c/72$. In this regime, we may apply Lemma 2 to obtain the upper bound $\mathbb{P}_\phi[\hat{\alpha}(M_T) \neq \alpha] \leq \frac{1}{3}$. Combining this upper bound with the lower bound (26) yields the inequality

$$\frac{1}{3} \geq 1 - 2 \frac{16T\delta^2 + \log 2}{d \log(2/\sqrt{e})}.$$ 

Recalling that $c = \frac{L}{2}$, making the substitution $\delta = \frac{18\epsilon}{c} = \frac{36\epsilon}{L}$, and performing some algebra yields

$$T = \Omega\left(\frac{L^2d}{\epsilon^2}\right) \quad \text{for all } d \geq 11 \text{ and for all } \epsilon \leq \frac{L}{144}.$$ 

Combined with Theorem 5.3.1 of NY [4], we conclude that this lower bound holds for all dimensions $d$.

**Part (b)—Proof for $p > 2$:** The preceding proof based on Oracle A is also valid for $p > 2$, but yields a relatively weak result. Here we show how the use of Oracle B yields the stronger claim stated in Theorem (b). When using this oracle, all $d$ coin tosses at each round are revealed, so that Lemma 3 with $\ell = d$ yields the lower bound

$$\mathbb{P}[\hat{\alpha}(M_T) \neq \alpha] \geq 1 - 2 \frac{16T d \delta^2 + \log 2}{d \log(2/\sqrt{e})}. \quad (28)$$

We now seek an upper bound on $\mathbb{P}[\hat{\alpha}(M_T) \neq \alpha]$. As before, we use the set $S = B_\infty(1/2)$, and the previous definitions (27) of $f^+_i(x)$ and $f^-_i(x)$. From our earlier analysis (in particular, equation (22)), the the Lipschitz constant of $g_\alpha(x)$ is at most $cd^{1/p-1}$, so that setting $c = Ld^{1-1/p}$ yields functions that are Lipschitz with parameter $L$.

As before, for any distinct pair $\alpha, \beta \in \mathcal{V}$, we have the lower bound

$$\rho(g_\alpha, g_\beta) = \frac{2c\delta}{d} \Delta_H(\alpha, \beta) \geq \frac{c\delta}{2},$$

so that $\psi(\delta) \geq \frac{c\delta}{2}$. Consequently, if we set the target error $\epsilon := \frac{c\delta}{18}$, then we are guaranteed that $\epsilon < \frac{\psi(\delta)}{9}$, as is required for applying Lemma 2 Application of this lemma yields the upper bound $\mathbb{P}_\phi[\hat{\alpha}(M_T) \neq \alpha] \leq \frac{1}{3}$. Combined with the lower bound (28), we obtain the inequality

$$\frac{1}{3} \geq 1 - 2 \frac{16T d \delta^2 + \log 2}{d \log(2/\sqrt{e})}.$$ 

Substituting $\delta = 18\epsilon/c$ yields the scaling $\epsilon = O\left(\frac{L^2}{d^p}\right)$ for all $d \geq 11$ and $\epsilon \leq c/72$. Recalling that $c = Ld^{1-1/p}$, we obtain the bound (11). Combining this with Theorem 5.3.1 of NY [4] gives the claim for all dimensions.

We have thus completed the proof of Theorem in the special case $S = B_\infty(1/2)$. In order to prove the general claims, which scale with $r$ when $B_\infty(r) \subseteq S$, we note that our preceding proof
required only that \( S \supseteq \mathcal{B}_\infty(\frac{1}{r}) \) so that the minimizing points \( x_\alpha = -\alpha/2 \in S \) for all \( \alpha \). In the general case, we define our base functions to be

\[
 f_i^+(x) = |x(i) + \frac{r}{2}|, \quad \text{and} \quad f_i^-(x) = |x(i) - \frac{r}{2}|.
\]

With this choice, the functions \( g_\alpha(x) \) are minimized at \( x_\alpha = -r\alpha/2 \), and \( \inf_{x \in S} g_\alpha(x) = cd/2 - cr\delta \).

Mimicking the previous steps with \( \theta < \frac{1}{r} \) functions from Definition 3. In this case, we will work with the following base functions parametrized by \( \theta < 1 \):

\[
 f_i^+(x) = r\theta|x(i) + r| + (1 - \theta) \left( x(i) + \frac{r}{2} \right)^2, \quad \text{and} \quad f_i^-(x) = r\theta|x(i) - r| + (1 - \theta) \left( x(i) - \frac{r}{2} \right)^2. \tag{29}
\]

A main component of our proof will be understanding the \( \rho \)-separation of the induced \( g_\alpha \) functions which we show in the following lemma.

**Lemma 4.** With the base functions defined in Equation (29),

\[
 \rho(g_\alpha, \beta) \geq \begin{cases} 
 \frac{4\delta^2 r^2}{(1 - \theta)\rho d} \Delta_H(\alpha, \beta) & \text{if } 1 - \theta > 2\delta \\
 \frac{2\delta^2 r^2}{d} \Delta_H(\alpha, \beta) & \text{if } 1 - \theta \leq 2\delta 
\end{cases}
\]

The proof of the lemma is deferred to Appendix A, we prove the theorem using the lemma here.

**Part (a)—Proof for** \( p = 1 \): We start by setting \( \gamma^2 = Ld^{-1/p}/r = L/d \), where we have recalled that \( p = 1 \). In this case, the base class of Equation (29) specializes to

\[
 f_i^+(x) = \frac{1}{2} \left( x(i) + r \right)^2, \quad \text{and} \quad f_i^-(x) = \frac{1}{2} \left( x(i) - r \right)^2,
\]

for \( i = 1, \ldots, d \). With this choice, the functions \( g_\alpha \) are strongly convex with respect to the Euclidean norm with coefficient \( \gamma = c/d \). Further, each \( f_i^+, f_i^- \) is \( cr \)-Lipschitz, since it depends on only one co-ordinate. Thus the functions \( g_\alpha \) are also \( cr \)-Lipschitz with respect to the \( \| \cdot \|_1 \)-norm as before. We consider Oracle A that returns one-dimensional gradients, so that \( E\| \hat{g}_\alpha(x) \|^2 \leq c^2 \). Once again, we need to lower bound the discrepancy \( \rho(g_\alpha, g_\beta) \) so as to complete the proof. Using Lemma 4 guarantees that \( \psi(\delta) = c\delta^2 r^2/(1 - \theta) \) for \( 1 - \theta > 2\delta \) and \( \psi(\delta) = \delta r^2/2 \) for \( 1 - \theta < 2\delta \). Considering the first regime, we set \( \epsilon = c\delta^2 r^2/(18(1 - \theta)) \) so that \( \epsilon \leq \psi(\delta)/9 \). Recalling that \( \delta \leq 1/4 \), we have for all \( \epsilon \leq cr^2/(288(1 - \theta)) \) and \( d \geq 11 \),

\[
 T = \Omega \left( \frac{cr^2}{(1 - \theta)\epsilon} \right).
\]
In the second regime, we set $\epsilon = c\delta r^2/36$ and simplify just like the proof of Theorem 3(a) obtaining

$$T = \Omega\left(\frac{L^2d}{\epsilon^2}\right).$$

The case for $d < 11$ follows from by considering the special case of just two-way hypothesis testing for $d = 1$, so that $\log |V| = 1$. The statement of the theorem is obtained by substituting $c = L/r$ and $\gamma^2 = L/dr$. To obtain the result for arbitrary values of $\gamma^2 \leq 4L/dr$, we use the general construction from Equation 29 and observe that the resulting functions $g_\alpha$ are $cr$-Lipschitz and $c(1-\theta)/d$-strongly convex. Hence $\gamma^2 = L(1-\theta)/dr$. Hence we can set $\theta \in [0,1]$ to match the desired value of strong convexity. We set $c = L/r$ to match the Lipschitz constant, and substitute $\theta = 1 - \gamma^2dr/L$. Using the bound from Lemma 4 and mimicking the rest of the proof above with these quantities gives the statement of the theorem for general values of $\gamma^2$.

Part (a)—Proof for $p > 2$: We start by setting $\gamma^2 = Ld^{-1/p}/r$. We use the same base class as before, but now switch to Oracle B that returns $d$-dimensional gradients. As before, we have

$$E\|\nabla g_\alpha,A(x)\|^2 \leq c^2d^{2/p-2}\epsilon^2.$$  

We set $c = Ld^{-1/p}/r$ so as to obtain functions that are $L$-Lipschitz. In this case, the strong convexity modulus is

$$\gamma^2 = c/d = \frac{Ld^{-1/p}}{r}$$

as desired. Also $\rho(g_\alpha,g_\beta)$ is given by Lemma 4 as before. The remainder of the proof is identical to Theorem 3(b), and we obtain the result by substituting $c = Ld^{-1/p}/r$ and $\gamma^2 = Ld^{-1/p}/r$. Finally, the result for arbitrary settings of $\gamma^2$ is obtained via the same argument as the proof of part (a).

4.4 Proof of Theorem 3

We begin by constructing an appropriate subset of $\mathcal{F}_{sp}(k)$ over which the Fano method can be applied. Let $V(k) := \{\alpha^1, \ldots, \alpha^M\}$ be a set of vectors, such that each $\alpha^j \in \{-1,0,+1\}^d$ satisfies

$$\|\alpha^j\|_0 = k \quad \text{for all } j = 1, \ldots, M, \quad \text{and} \quad \Delta_H(\alpha^j,\alpha^\ell) \geq \frac{k}{2} \quad \text{for all } j \neq \ell.$$

It can be shown that there exists such a packing set with $|V(k)| \geq \exp\left(\frac{k}{2} \log \frac{d-k}{k/2}\right)$ elements (e.g., see Lemma 5 in Raskutti et al. [19]).

For any $\alpha \in V(k)$, we define the function

$$g_\alpha(x) := c \left[ \sum_{i=1}^d \left\{ \left( \frac{1}{2} + \alpha_i \delta \right)|x(i) + r| + \left( \frac{1}{2} - \alpha_i \delta \right)|x(i) - r| \right\} + \delta \sum_{i=1}^d |x(i)| \right].$$  

(30)

In this definition, the quantity $c > 0$ is a pre-factor to be chosen later, and $\delta \in (0,\frac{1}{4}]$ is a given error tolerance. Observe that each function $g_\alpha \in \mathcal{G}(\delta;k)$ is convex, and Lipschitz with parameter $c$ with respect to the $\|\cdot\|_\infty$ norm.

Central to the remainder of the proof is the function class $\mathcal{G}(\delta;k) := \{g_\alpha, \alpha \in V(k)\}$. In particular, we need to control the discrepancy $\psi(\delta;k) := \psi(\mathcal{G}(\delta;k))$ for this class. The following result, proven in Appendix B, provides a suitable lower bound:
Lemma 5. We have
\[
\psi(\delta; k) = \inf_{\alpha \neq \beta \in \mathcal{V}(k)} \rho(g_\alpha, g_\beta) \geq \frac{ck\delta r}{4}. \tag{31}
\]
Using Lemma 5, we may complete the proof of Theorem 3. Define the base functions
\[
f^+_i(x) := d \left( |x(i) + r| + \frac{\delta}{2} |x(i)| \right), \quad \text{and} \quad f^-_i(x) := d \left( |x(i) - r| + \frac{\delta}{2} |x(i)| \right).
\]
Consider Oracle B, which returns \(d\)-dimensional gradients based on the function
\[
\hat{g}_{\alpha,B}(x) = \frac{c}{d} \sum_{i=1}^{d} \left[ b_i f^+_i(x) + (1 - b_i) f^-_i(x) \right],
\]
where \(\{b_i\}\) are Bernoulli variables. By construction, the function \(\hat{g}_{\alpha,B}\) is at most \(3c\)-Lipschitz in \(\ell_\infty\) norm, so that setting \(c = \frac{L}{3}\) yields an \(L\)-Lipschitz function.

Our next step is to use Fano’s inequality [15] to lower bound the probability of error in the multiway testing problem associated with this stochastic oracle, following an argument similar to (but somewhat simpler than) the proof of Lemma 3. Fano’s inequality yields the lower bound
\[
\Pr[\hat{\alpha} \neq \alpha^*] \geq 1 - \frac{1}{\log |\mathcal{V}|} \sum_{\alpha \neq \beta} D(\mathbb{P}_\alpha \| \mathbb{P}_\beta) + \log 2, \tag{32}
\]
(As in the proof of Lemma 3, we have used convexity of mutual information [15] to bound it by the average of the pairwise KL divergences.) By construction, any two parameters \(\alpha, \beta \in \mathcal{V}\) differ in at most \(2k\) places, and the remaining entries are all zeroes in both vectors. The proof of Lemma 3 shows that for \(\delta \in [0, \frac{1}{4}]\), each of these \(2k\) places makes a contribution of at most \(16 \delta^2\). Recalling that we have \(T\) samples, we conclude that \(D(\mathbb{P}_\alpha \| \mathbb{P}_\beta) \leq 32kT\delta^2\). Substituting this upper bound into the Fano lower bound (32) and recalling that the cardinality of \(\mathcal{V}\) is at least \(\exp\left(\frac{d-k}{2} \log \frac{k}{2} \right)\), we obtain
\[
\Pr[\hat{G}(\mathcal{M}_T) \neq \alpha] \geq 1 - 2 \left( \frac{32kT\delta^2 + \log 2}{\frac{2}{k} \log \frac{d-k}{2}} \right) \tag{33}
\]
By Lemma 5 and our choice \(c = L/3\), we have
\[
\psi(\delta) \geq \frac{ck\delta r}{4} = \frac{Lk\delta r}{12}.
\]
Therefore, if we aim for the target error \(\epsilon = \frac{Lk\delta r}{108}\), then we are guaranteed that \(\epsilon \leq \frac{\psi(\delta)}{2}\), as is required for the application of Lemma 2. Recalling the requirement \(\delta \leq 1/4\) gives \(\epsilon \leq Lk\delta r/432\). Now Lemma 2 implies that \(\Pr[\hat{G}(\mathcal{M}_T) \neq \alpha] \leq 1/3\), which when combined with the earlier bound (33) yields
\[
\frac{1}{3} \geq 1 - 2 \left( \frac{32kT\delta^2 + \log 2}{\frac{2}{k} \log \frac{d-k}{2}} \right).
\]
Rearranging yields the lower bound
\[
T = \Omega\left( \frac{\log \frac{d-k}{2}}{\delta^2} \right) = \Omega\left( \frac{L^2 r^2 k^2 \log \frac{d-k}{2}}{\epsilon^2} \right),
\]
where the second step uses the relation \(\delta = \frac{108\epsilon}{Lkr}\). As long as \(k \leq \lfloor d/2 \rfloor\), we have \(\log \frac{d-k}{k^2/2} = \Theta(\log \frac{d}{k})\), which completes the proof.
5 Discussion

In this paper, we have studied the complexity of convex optimization within the stochastic first-order oracle model. We derived lower bounds for various function classes, including convex functions, strongly convex functions, and convex functions with sparse optima. As we discussed, our lower bounds are sharp in general, since there are matching upper bounds achieved by known algorithms, among them stochastic gradient descent and stochastic mirror descent. Our bounds also reveal various dimension-dependent and geometric aspects of the stochastic oracle complexity of convex optimization. An interesting aspect of our proof technique is the features in common statistical minimax theory. In particular, our proofs are based on constructing packing sets, defined with respect to a semi-metric that measures how the degree of separation between the optima of different functions. We then leveraged information-theoretic techniques, in particular Fano’s inequality and its variants, in order to establish lower bounds.

There are various directions for future research. It would be interesting to consider the effect of memory constraints on the complexity of convex optimization, or to derive lower bounds for problems of distributed optimization. We suspect that the proof techniques developed in this paper may be useful for studying these related problems.

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A Proof of Lemma 4

Recalling the definition (17) of the discrepancy $\rho$, we need to compute the quantities $\inf_{x \in \mathbb{S}} g_\alpha(x)$ and $\inf_{x \in \mathbb{S}} \{g_\alpha(x) + g_\beta(x)\}$. Beginning with the former quantity, first observe that for any $x \in B_\infty(r)$, we have

$$|x(i) + r| = x(i) + r \quad \text{and} \quad |x(i) - r| = r - x(i).$$

Consequently, recalling the definitions of the base functions (29), we have

$$\left( \frac{1}{2} + \alpha_i \delta \right) f_i^+(x) + \left( \frac{1}{2} - \alpha_i \delta \right) f_i^-(x) = \frac{1 - \theta}{2} x(i)^2 + 2 \alpha_i \delta r x i + \frac{r^2(1 + \theta)}{2}.$$

This function is minimized at $x(i) = -2 \alpha_i \delta r/(1 - \theta)$ when $1 - \theta > 2 \delta$ and $x(i) = -\alpha_i r$ otherwise. Combining over all coordinates, we get that

$$\inf_{x \in B_\infty(r)} g_\alpha = \begin{cases} 
\frac{-2 \delta^2 r^2 c}{(1 - \theta)} + \frac{r^2(1 + \theta)}{2} & \text{if } 1 - \theta > 2 \delta \\
\frac{r^2 c}{2} & \text{if } 1 - \theta < 2 \delta
\end{cases}.$$

Reasoning similarly, we can compute $\inf_{x \in B_\infty(r)} \{g_\alpha(x) + g_\beta(x)\}$. Combining the two just like the proof of Theorem 1(a) gives the result.
B Proof of Lemma \[5\]

Recalling the definition (17) of the discrepancy \( \rho \), we need to compute the quantities \( \inf_{x \in S} g_\alpha(x) \) and \( \inf_{x \in S}\{g_\alpha(x) + g_\beta(x)\} \). Beginning with the former quantity, first observe that for any \( x \in \mathbb{B}_\infty(r) \), we have

\[
\left[ r + \alpha_i \delta \right] |x(i) + \frac{1}{2}| + \left[ r - \alpha_i \delta \right] |x(i) - \frac{1}{2}| = r + 2\alpha_i \delta x(i). \tag{34}
\]

We now consider one of the individual terms arising in the definition (15) of the function \( g_\alpha \). Using the relation (34), it can be written as

\[
\left( \frac{1}{2} + \alpha_i \delta \right) f_i^+(x) + \left( \frac{1}{2} - \alpha_i \delta \right) f_i^-(x) = \left( \frac{1}{2} + \alpha_i \delta \right) |x(i) + r| + \left( \frac{1}{2} - \alpha_i \delta \right) |x(i) - r| + \delta |x(i)|
\]

\[
= \begin{cases} 
 r + (2\alpha_i - 1)\delta x(i) & \text{if } x(i) \leq 0 \\
 r + (2\alpha_i + 1)\delta x(i) & \text{if } x(i) \geq 0 
\end{cases}
\]

From this representation, we see that whenever \( \alpha_i \neq 0 \), then the \( i \)-th term in the summation defining \( g_\alpha \) minimized at \( x(i) = -r\alpha_i \), at which point it takes on its minimum value \( r(1 - \delta) \). On the other hand, for any term with \( \alpha_i = 0 \), the function is minimized at \( x(i) = 0 \) with associated minimum value of \( r \). Combining these two facts shows that the vector \( -\alpha r \) is an element of the set arg \( \min_{x \in S} g_\alpha(x) \), and moreover that

\[
\inf_{x \in S} g_\alpha(x) = cr(d - k\delta). \tag{35}
\]

We now turn to the computation of \( \inf_{x \in S}\{g_\alpha(x) + g_\beta(x)\} \). From the relation (34) and the definitions of \( g_\alpha \) and \( g_\beta \), some algebra yields

\[
\inf_{x \in S}\{g_\alpha(x) + g_\beta(x)\} = c \inf_{x \in S} \sum_{i=1}^d \left\{ 2r + 2\delta \left[ (\alpha_i + \beta_i)x(i) + |x(i)| \right] \right\}. \tag{36}
\]

Let us consider the minimizer of the \( i \)-th term in this summation. First, suppose that \( \alpha_i \neq \beta_i \), in which case there are two cases to consider.

- If \( \alpha_i \neq \beta_i \) and neither \( \alpha_i \) nor \( \beta_i \) is zero, then we must have \( \alpha_i + \beta_i = 0 \), so that the minimum value of \( 2r \) is achieved at \( x(i) = 0 \).

- Otherwise, suppose that \( \alpha_i \neq 0 \) and \( \beta_i = 0 \). In this case, we see from Equation (36) that it is equivalent to minimizing \( \alpha_i x(i) + |x(i)| \). Setting \( x(i) = -\alpha_i \) achieves the minimum value of \( 2r \).

In the remaining two cases, we have \( \alpha_i = \beta_i \).

- If \( \alpha_i = \beta_i \neq 0 \), then the component is minimized at \( x(i) = -\alpha_i r \) and the minimum value along the component is \( 2r(1 - \delta) \).

- If \( \alpha_i = \beta_i = 0 \), then the minimum value is \( 2r \), achieved at \( x(i) = 0 \).

Consequently, accumulating all of these individual cases into a single expression, we obtain

\[
\inf_{x \in S}\{g_\alpha(x) + g_\beta(x)\} = 2cr \left( d - \delta \sum_{i=1}^d \mathbb{I}[\alpha_i = \beta_i \neq 0] \right). \tag{37}
\]
Finally, combining equations (35) and (37) in the definition of ρ, we find that

\[ ρ(g_α, g_β) = 2cr \left[ d - \delta \sum_{i=1}^{d} \mathbb{I}[α_i = β_i ≠ 0] - (d - kδ) \right] \]

\[ = 2cδr \left[ k - \sum_{i=1}^{d} \mathbb{I}[α_i = β_i ≠ 0] \right] \]

\[ = crδ \Delta_H(α, β), \]

where the second equality follows since α and β have exactly k non-zero elements each. Finally, since V is an k/2-packing set in Hamming distance, we have \( \Delta_H(α, β) ≥ k/2 \), which completes the proof.

C Upper bounds via mirror descent

This appendix is devoted to background on the family of mirror descent methods. We first describe the basic form of the algorithm and some known convergence results, before showing that different forms of mirror descent provide matching upper bounds for several of the lower bounds established in this paper, as discussed in the main text.

C.1 Background on mirror descent

Mirror descent is a generalization of (projected) stochastic gradient descent, first introduced by Nemirovski and Yudin [4]; here we follow a more recent presentation of it due to Beck and Teboulle [20]. For a given norm \( \| \cdot \| \), let \( Φ : \mathbb{R}^d → \mathbb{R} \cup \{ +∞ \} \) be a differentiable function that is 1-strongly convex with respect to \( \| \cdot \| \), meaning that

\[ Φ(y) ≥ Φ(x) + \langle \nabla Φ(x), y - x \rangle + \frac{1}{2} \| y - x \|^2. \]

We assume that \( Φ \) is a function of Legendre type [21, 22], which implies that the conjugate dual \( Φ^* \) is differentiable on its domain with \( \nabla Φ^* = (\nabla Φ)^{-1} \). For a given proximal function, we let \( D_Φ \) be the Bregman divergence induced by \( Φ \), given by

\[ D_Φ(x, y) := Φ(x) - Φ(y) - \langle \nabla Φ(y), x - y \rangle. \quad (38) \]

With this set-up, we can now describe the mirror descent algorithm based on the proximal function \( Φ \) for minimizing a convex function \( f \) over a convex set \( S \) contained within the domain of \( Φ \). Starting with an arbitrary initial \( x_0 ∈ S \), it generates a sequence \( \{x_t\}_{t=0}^{∞} \) contained within \( S \) via the updates

\[ x_{t+1} = \arg \min_{x ∈ S} \{ η_t \langle x, \nabla f(x_t) \rangle + D_Φ(x, x_t) \}, \quad (39) \]

where \( η_t > 0 \) is a stepsize.

A special case of this algorithm is obtained by choosing the proximal function \( Φ(x) = \frac{1}{2} \| x \|^2 \), which is 1-strongly convex with respect to the Euclidean norm. The associated Bregman divergence \( D_Φ(x, y) = \frac{1}{2} \| x - y \|^2 \) is simply the Euclidean norm, so that the updates (39) correspond to a standard projected gradient descent method. If one receives only an unbiased estimate of the gradient \( \nabla f(x_t) \), then this algorithm corresponds to a form of projected stochastic gradient descent. Moreover, other choices of the proximal function lead to different stochastic algorithms, as discussed below.
In order to state convergence rates for this algorithm, convexity and Lipschitz assumptions about the functions $f$ are required. Following the set-up used in our lower bound analysis, we assume that $\|\nabla \tilde{z}(x_t)\|_* \leq L$ for all $x \in S$, where $\|v\|_* := \sup_{\|x\| \leq 1} (x, v)$ is the dual norm defined by $\| \cdot \|$. Given stochastic mirror descent based on unbiased estimates of the gradient, Beck and Teboulle [20] showed that with the initialization $x_0 = \arg \min_{x \in S} \Phi(x)$ and stepsizes $\eta_t = 1/\sqrt{t}$, the optimization error of the sequence $\{x_t\}$ is bounded as

\[
\frac{1}{T} \sum_{t=1}^{T} \mathbb{E}[f(x_t) - f(x^*)] \leq L \sqrt{D\Phi(x^*,x_1)} \leq L \sqrt{\frac{\Phi(x^*)}{T}} \quad (40)
\]

Note that this averaged convergence is a little different from the convergence of $x_T$ discussed in our lower bounds. In order to relate the two quantities, observe that by Jensen’s inequality

\[
\mathbb{E}\left[f\left(\frac{\sum_{t=1}^{T} x_t}{T}\right)\right] \leq \frac{1}{T} \mathbb{E}[f(x_t)].
\]

Consequently, based on mirror descent for $T - 1$ rounds, we may set $x_T = \frac{1}{T-1} \sum_{t=1}^{T-1} x_t$ so as to obtain the same convergence bounds up to constant factors. In the following discussion, we assume this choice of $x_T$ for comparing the mirror descent upper bounds to our lower bounds.

### C.2 Matching upper bounds

Now consider the form of mirror descent obtained by choosing the proximal function

\[
\Phi_a(x) := \frac{1}{(a - 1)} \|x\|_a^2 \quad \text{for } 1 < a \leq 2. \quad (41)
\]

Note that this proximal function is 1-strongly convex with respect to the $\ell_a$-norm, meaning that

\[
\frac{1}{(a - 1)} \|x\|_a^2 \geq \frac{1}{(a - 1)} \|y\|_a^2 + \left(\nabla \frac{1}{(a - 1)} \|x\|_a^2\right)^T (x - y) + \frac{1}{2} \|x - y\|_a^2.
\]

**Upper bounds for Theorem 1** For this case, we use mirror descent based on the proximal function $\Phi_a$ with $a = q$. Under the condition $\|x^*\|_\infty \leq 1$, a condition which holds in our lower bounds, we obtain

\[
\|x^*\|_q \leq \|x^*\|_\infty \ d^{1/q} = d^{1/q},
\]

which implies that $\Phi_q(x^*) = O(d^{2/q})$. Under the conditions of Theorem 1 we have $\|\nabla f(x_t)\|_p \leq L$ where $p = 1 - 1/q$ defines the dual norm. Note that the condition $1 < q \leq 2$ implies that $p \geq 2$. Substituting this in the upper bound (40) yields

\[
\mathbb{E}[f(x_T) - f(x^*)] = O\left(L \sqrt{d^{2/q}/T}\right) = O\left(L d^{1-1/p} \sqrt{\frac{1}{T}}\right),
\]

which matches the lower bound from Theorem 1(b). Moreover, setting $q = p = 2$ corresponds to stochastic gradient descent, and yields an upper bound to match the lower bound of Theorem 1(a).
Upper bounds for Theorem 3: In order to recover matching upper bounds in this case, we use the function \( \Phi_a \) from equation (41) with \( a = \frac{2 \log d}{2 \log d - 1} \). In this case, the resulting upper bound (40) on the convergence rate takes the form

\[
\mathbb{E} \left[ f(x_T) - f(x^*) \right] = O \left( L \sqrt{\frac{\|x^*\|_2^2}{2(a - 1)T}} \right) = O \left( L \sqrt{\frac{\|x^*\|_2^2 \log d}{T}} \right),
\]

(42)

since \( \frac{1}{a - 1} = 2 \log d - 1 \). Based on the conditions of Theorem 3, we are guaranteed that \( x^* \) is \( k \)-sparse, with every component bounded by 1 in absolute value, so that \( \|x^*\|_2^2 \leq k^{2/a} \leq k^2 \), where the final inequality follows since \( a > 1 \). Substituting this upper bound back into Equation (42) yields

\[
\mathbb{E} \left[ f(x_T) - f(x^*) \right] = O \left( L \sqrt{\frac{k^2 \log d}{T}} \right).
\]

Note that whenever \( k = O(d^{1-\delta}) \) for some \( \delta > 0 \), then we have \( \log d = \Theta(\log d^\frac{1}{\delta}) \), in which case this upper bound matches the lower bound from Theorem 3 up to constant factors, as claimed.

References


