

## **Introduction to Time Series Analysis. Lecture 23.**

1. Lagged regression models.
2. Cross-covariance function, sample CCF.
3. Lagged regression in the time domain: prewhitening.
4. Lagged regression in the frequency domain: Cross spectrum.  
Coherence.

## Lagged regression models

Consider a lagged regression model of the form

$$Y_t = \sum_{h=-\infty}^{\infty} \beta_h X_{t-h} + V_t,$$

where  $X_t$  is an observed input time series,  $Y_t$  is the observed output time series, and  $V_t$  is a stationary noise process.

This is useful for

- Identifying the (best linear) relationship between two time series.
- Forecasting one time series from the other.  
(We might want  $\beta_h = 0$  for  $h < 0$ .)

## Lagged regression models

$$Y_t = \sum_{h=-\infty}^{\infty} \beta_h X_{t-h} + V_t.$$

In the SOI and recruitment example, we might wish to identify how the values of the recruitment series (the number of new fish) is related to the Southern Oscillation Index.

Or we might wish to predict future values of recruitment from the SOI.

## Lagged regression models: Agenda

- Multiple, jointly stationary time series in the time domain: cross-covariance function, sample CCF.
- Lagged regression in the time domain: model the input series, extract the white time series driving it ('prewhitening'), regress with transformed output series.
- Multiple, jointly stationary time series in the frequency domain: cross spectrum, coherence.
- Lagged regression in the frequency domain: Calculate the input's spectral density, and the cross-spectral density between input and output, and find the transfer function relating them, in the frequency domain. Then the regression coefficients are the inverse Fourier transform of the transfer function.

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## Cross-covariance

Recall that the autocovariance function of a stationary process  $\{X_t\}$  is

$$\gamma_x(h) = \mathbf{E} [(X_{t+h} - \mu_x)(X_t - \mu_x)].$$

The *cross-covariance function* of two jointly stationary processes  $\{X_t\}$  and  $\{Y_t\}$  is

$$\gamma_{xy}(h) = \mathbf{E} [(X_{t+h} - \mu_x)(Y_t - \mu_y)].$$

(Jointly stationary = constant means, autocovariances depending only on the lag  $h$ , and cross-covariance depends only on  $h$ .)

## Cross-correlation

The *cross-correlation function* of jointly stationary  $\{X_t\}$  and  $\{Y_t\}$  is

$$\rho_{xy}(h) = \frac{\gamma_{xy}(h)}{\sqrt{\gamma_x(0)\gamma_y(0)}}.$$

Notice that  $\rho_{xy}(h) = \rho_{yx}(-h)$  (but  $\rho_{xy}(h) \neq \rho_{xy}(-h)$ ).

**Example:** Suppose that  $Y_t = \beta X_{t-\ell} + W_t$  for  $\{X_t\}$  stationary and uncorrelated with  $\{W_t\}$ , and  $W_t$  zero mean and white. Then  $\{X_t\}$  and  $\{Y_t\}$  are jointly stationary, with  $\mu_y = \beta\mu_x$ ,

$$\gamma_{xy}(h) = \beta\gamma_x(h + \ell).$$

If  $\ell > 0$ , we say  $x_t$  *leads*  $y_t$ .

If  $\ell < 0$ , we say  $x_t$  *lags*  $y_t$ .

## Sample cross-covariance and sample CCF

$$\hat{\gamma}_{xy}(h) = \frac{1}{n} \sum_{i=1}^{n-h} (x_{t+h} - \bar{x})(y_t - \bar{y})$$

for  $h \geq 0$  (and  $\hat{\gamma}_{xy}(h) = \hat{\gamma}_{yx}(-h)$  for  $h < 0$ ).

The sample CCF is

$$\hat{\rho}_{xy}(h) = \frac{\hat{\gamma}_{xy}(h)}{\sqrt{\hat{\gamma}_x(0)\hat{\gamma}_y(0)}}.$$



## Sample cross-covariance and sample CCF

If either of  $\{X_t\}$  or  $\{Y_t\}$  is white, then  $\hat{\rho}_{xy}(h) \sim AN(0, 1/\sqrt{n})$ .

Notice that we can look for peaks in the sample CCF to identify a leading or lagging relation. (Recall that the ACF of the input series peaks at  $h = 0$ .)

Example: CCF of SOI and recruitment (Figure 1.14 in text) has a peak at  $h = -6$ , indicating that recruitment at  $t$  has its strongest correlation with SOI at time  $t - 6$ . Thus, SOI leads recruitment (by 6 months).

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## Lagged regression in the time domain (Section 5.6)

Suppose we wish to fit a lagged regression model of the form

$$Y_t = \alpha(B)X_t + \eta_t = \sum_{j=0}^{\infty} \alpha_j X_{t-j} + \eta_t,$$

where  $X_t$  is an observed input time series,  $Y_t$  is the observed output time series, and  $\eta_t$  is a stationary noise process, uncorrelated with  $X_t$ .

One approach (pioneered by Box and Jenkins) is to fit ARIMA models for  $X_t$  and  $\eta_t$ , and then find a simple rational representation for  $\alpha(B)$ .

## Lagged regression in the time domain

$$Y_t = \alpha(B)X_t + \eta_t = \sum_{j=0}^{\infty} \alpha_j X_{t-j} + \eta_t,$$

For example:

$$X_t = \frac{\theta_x(B)}{\phi_x(B)} W_t,$$

$$\eta_t = \frac{\theta_\eta(B)}{\phi_\eta(B)} Z_t,$$

$$\alpha(B) = \frac{\delta(B)}{\omega(B)} B^d.$$

Notice the delay  $B^d$ , indicating that  $Y_t$  lags  $X_t$  by  $d$  steps.

## Lagged regression in the time domain

How do we choose all of these parameters?

1. Fit  $\theta_x(B)$ ,  $\phi_x(B)$  to model the input series  $\{X_t\}$ .
2. *Prewhiten* the input series by applying the inverse operator  $\phi_x(B)/\theta_x(B)$ :

$$\tilde{Y}_t = \frac{\phi_x(B)}{\theta_x(B)} Y_t = \alpha(B) W_t + \frac{\phi_x(B)}{\theta_x(B)} \eta_t.$$

## Lagged regression in the time domain

3. Calculate the cross-correlation of  $\tilde{Y}_t$  with  $W_t$ ,

$$\gamma_{\tilde{y},w}(h) = \mathbf{E} \left( \sum_{j=0}^{\infty} \alpha_j W_{t+h-j} W_t \right) = \sigma_w^2 \alpha_h,$$

to give an indication of the behavior of  $\alpha(B)$  (for instance, the delay).

4. Estimate the coefficients of  $\alpha(B)$  and hence fit an ARMA model for the noise series  $\eta_t$ .

## Lagged regression in the time domain

Why prewhiten?

The prewhitening step inverts the linear filter  $X_t = \theta_x(B)/\phi_x(B)W_t$ . Then the lagged regression is between the transformed  $Y_t$  and a white series  $W_t$ .

This makes it easy to determine a suitable lag.

For example, in the SOI/recruitment series, we treat SOI as an input, estimate an AR(1) model, prewhiten it (that is, compute the inverse of our AR(1) operator and apply it to the SOI series), and consider the cross-correlation between the transformed recruitment series and the prewhitened SOI. This shows a large peak at lag -5 (corresponding to the SOI series leading the recruitment series). Examples 5.7 and 5.8 in the text then consider  $\alpha(B) = B^5/(1 - \omega_1 B)$ .

## Lagged regression in the time domain

This sequential estimation procedure ( $\phi_x, \theta_x$ , then  $\alpha$ , then  $\phi_\eta, \theta_\eta$ ) is rather ad hoc. State space methods offer an alternative, and they are also convenient for vector-valued input and output series.



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Coherence.

## Coherence

To analyze lagged regression in the frequency domain, we'll need the notion of *coherence*, the analog of cross-correlation in the frequency domain.

Define the *cross-spectrum* as the Fourier transform of the cross-correlation,

$$f_{xy}(\nu) = \sum_{h=-\infty}^{\infty} \gamma_{xy}(h) e^{-2\pi i \nu h},$$
$$\gamma_{xy}(h) = \int_{-1/2}^{1/2} f_{xy}(\nu) e^{2\pi i \nu h} d\nu,$$

(provided that  $\sum_{h=-\infty}^{\infty} |\gamma_{xy}(h)| < \infty$ ).

Notice that  $f_{xy}(\nu)$  can be complex-valued.

Also,  $\gamma_{yx}(h) = \gamma_{xy}(-h)$  implies  $f_{yx}(\nu) = f_{xy}(\nu)^*$ .

## Coherence

The *squared coherence* function is

$$\rho_{y,x}^2(\nu) = \frac{|f_{yx}(\nu)|^2}{f_x(\nu)f_y(\nu)}.$$

Compare this with the correlation  $\rho_{y,x} = \text{Cov}(Y, X) / \sqrt{\sigma_x^2 \sigma_y^2}$ . We can think of the squared coherence at a frequency  $\nu$  as the contribution to squared correlation at that frequency.

(Recall the interpretation of spectral density at a frequency  $\nu$  as the contribution to variance at that frequency.)

## Estimating squared coherence

Recall that we estimated the spectral density using the smoothed squared modulus of the DFT of the series,

$$\begin{aligned}\hat{f}_x(\nu_k) &= \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} |X(\nu_k - l/n)|^2 \\ &= \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} X(\nu_k - l/n)X(\nu_k - l/n)^*.\end{aligned}$$

We can estimate the cross spectral density using the same sample estimate,

$$\hat{f}_{xy}(\nu_k) = \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} X(\nu_k - l/n)Y(\nu_k - l/n)^*.$$

## Coherence

Also, we can estimate the squared coherence using these estimates,

$$\hat{\rho}_{y,x}^2(\nu) = \frac{|\hat{f}_{yx}(\nu)|^2}{\hat{f}_x(\nu)\hat{f}_y(\nu)}.$$

(Knowledge of the asymptotic distribution of  $\hat{\rho}_{y,x}^2(\nu)$  under the hypothesis of no coherence,  $\rho_{y,x}(\nu) = 0$ , allows us to test for coherence.)

## Lagged regression models in the frequency domain

Consider a lagged regression model of the form

$$Y_t = \sum_{j=-\infty}^{\infty} \beta_j X_{t-j} + V_t,$$

where  $X_t$  is an observed input time series,  $Y_t$  is the observed output time series, and  $V_t$  is a stationary noise process.

We'd like to estimate the coefficients  $\beta_j$  that determine the relationship between the lagged values of the input series  $X_t$  and the output series  $Y_t$ .

## Lagged regression models in the frequency domain

The projection theorem tells us that the coefficients that minimize the mean squared error,

$$\mathbf{E} \left[ \left( Y_t - \sum_{j=-\infty}^{\infty} \beta_j X_{t-j} \right)^2 \right]$$

satisfy the orthogonality conditions

$$\mathbf{E} \left[ \left( Y_t - \sum_{j=-\infty}^{\infty} \beta_j X_{t-j} \right) X_{t-k} \right] = 0, \quad k = 0, \pm 1, \pm 2, \dots$$

$$\sum_{j=-\infty}^{\infty} \beta_j \gamma_x(k-j) = \gamma_{yx}(k), \quad k = 0, \pm 1, \pm 2, \dots$$

## Lagged regression models in the frequency domain

We could solve these equations for the  $\beta_j$  using the sample autocovariance and sample cross-covariance. But it is more convenient to use estimates of the spectra and cross-spectrum.

(Convolution with  $\{\beta_j\}$  in the time domain is equivalent to multiplication by the Fourier transform of  $\{\beta_j\}$  in the frequency domain. Let's verify this.)

We replace the autocovariance and cross-covariance with the inverse Fourier transforms of the spectral density and cross-spectral density in the orthogonality conditions,

$$\sum_{j=-\infty}^{\infty} \beta_j \gamma_x(k-j) = \gamma_{yx}(k), \quad k = 0, \pm 1, \pm 2, \dots$$



## Lagged regression models in the frequency domain

This gives, for  $k = 0, \pm 1, \pm 2, \dots$ ,

$$\int_{-1/2}^{1/2} \sum_{j=-\infty}^{\infty} \beta_j e^{2\pi i \nu (k-j)} f_x(\nu) d\nu = \int_{-1/2}^{1/2} e^{2\pi i \nu k} f_{yx}(\nu),$$

$$\int_{-1/2}^{1/2} e^{2\pi i \nu k} B(\nu) f_x(\nu) d\nu = \int_{-1/2}^{1/2} e^{2\pi i \nu k} f_{yx}(\nu),$$

where  $B(\nu) = \sum_{j=-\infty}^{\infty} e^{-2\pi i \nu j} \beta_j$  is the Fourier transform of the coefficient sequence  $\beta_j$ .

Since the Fourier transform is unique, the orthogonality conditions are equivalent to

$$B(\nu) f_x(\nu) = f_{yx}(\nu).$$

## Lagged regression models in the frequency domain

We can write the mean squared error at the solution as follows. (Why?)

$$\begin{aligned} \mathbf{E} \left[ \left( Y_t - \sum_{j=-\infty}^{\infty} \beta_j X_{t-j} \right) Y_t \right] &= \gamma_y(0) - \sum_{j=-\infty}^{\infty} \beta_j \gamma_{xy}(-j) \\ &= \int_{-1/2}^{1/2} (f_y(\nu) - B(\nu) f_{xy}(\nu)) d\nu \\ &= \int_{-1/2}^{1/2} f_y(\nu) \left( 1 - \frac{f_{yx}(\nu) f_{xy}(\nu)}{f_x(\nu) f_y(\nu)} \right) d\nu \\ &= \int_{-1/2}^{1/2} f_y(\nu) \left( 1 - \frac{|f_{yx}(\nu)|^2}{f_x(\nu) f_y(\nu)} \right) d\nu \\ &= \int_{-1/2}^{1/2} f_y(\nu) (1 - \rho_{yx}^2(\nu)) d\nu. \end{aligned}$$

## Lagged regression models in the frequency domain

$$MSE = \int_{-1/2}^{1/2} f_y(\nu) (1 - \rho_{yx}^2(\nu)) d\nu.$$

Thus,  $\rho_{yx}(\nu)^2$  indicates how the component of the variance of  $\{Y_t\}$  at a frequency  $\nu$  is accounted for by  $\{X_t\}$ . Compare this with the corresponding decomposition for random variables:

$$\mathbf{E}(y - \beta x)^2 = \sigma_y^2(1 - \rho_{xy}^2).$$

## Lagged regression models in the frequency domain

We can estimate the  $\beta_j$  in the frequency domain:

$$\hat{B}(\nu_k) = \frac{\hat{f}_{yx}(\nu_k)}{\hat{f}_x(\nu_k)}.$$

We can approximate the inverse Fourier transform of  $\hat{B}(\nu)$ ,

$$\hat{\beta}_j = \int_{-1/2}^{1/2} e^{2\pi i \nu j} \hat{B}(\nu) d\nu$$

via the sum,

$$\hat{\beta}_j = \frac{1}{M} \sum_{k=0}^{M-1} \hat{B}(\nu_k) e^{2\pi i \nu_k j}.$$

This gives a periodic sequence—we might truncate at  $j = M/2$ .

## Lagged regression models in the frequency domain

Here is the approach:

1. Estimate the spectral density and cross-spectral density.
2. Compute the transfer function  $\hat{B}(\nu)$ .
3. Take the inverse Fourier transform to obtain the impulse response function  $\beta_j$ .

## Lagged regression models in the frequency domain

It is often useful to consider both representations

$$Y_t = \sum_{j=-\infty}^{\infty} \alpha_j X_{t-j}, \quad X_t = \sum_{j=-\infty}^{\infty} \beta_j Y_{t-j},$$

since there might be a more parsimonious representation in terms of one than the other. (Just as a small AR model often cannot be well approximated by a small MA model.)

## Lagged regression models in the frequency domain

In the  $X_t = \text{SOI}/Y_t = \text{Recruitment}$  example (Example 4.23), we obtain

$$Y_t = -22X_{t-5} - 15X_{t-6} - 11X_{t-7} - 10X_{t-8} - 7X_{t-9} - \cdots + W_t,$$

$$X_t = 0.012Y_{t+4} - 0.018Y_{t+5} + V_t,$$

and the latter is equivalent to

$$(1 - 0.667B)Y_t = -56B^5 X_t + U_t.$$

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