# **Introduction to Time Series Analysis. Lecture 21.**

- 1. Review: The periodogram, the smoothed periodogram.
- 2. Other smoothed spectral estimators.
- 3. Consistency.
- 4. Asymptotic distribution.

### **Review: Periodogram**

The periodogram is defined as

$$I(\nu) = X_c^2(\nu) + X_s^2(\nu).$$

$$X_c(\nu) = \frac{1}{\sqrt{n}} \sum_{t=1}^n \cos(2\pi t\nu) x_t,$$
$$X_s(\nu) = \frac{1}{\sqrt{n}} \sum_{t=1}^n \sin(2\pi t\nu_j) x_t.$$

Under general conditions,  $X_c(\nu_j)$ ,  $X_s(\nu_j)$  are asymptotically independent and  $N(0, f(\nu_j)/2)$ . Thus,  $\mathbf{E}I(\hat{\nu}^{(n)}) \to f(\nu)$ , but  $\operatorname{Var}(I(\hat{\nu}^{(n)})) \to f(\nu)^2$ .

# **Review: smoothed periodogram**

If  $f(\nu)$  is approximately constant in a band of frequencies  $[\nu_k - L/(2n), \nu_k + L/(2n)]$ , we can average the periodogram over this band:

$$\begin{aligned}
\hat{T}(\nu_k) &= \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} I(\nu_k - l/n) \\
&= \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} \left( X_c^2(\nu_k - l/n) + X_s^2(\nu_k - l/n) \right).
\end{aligned}$$

#### **Review: smoothed periodogram**

Under general conditions, the  $X_c(\nu_k - l/n)$  and  $X_s(\nu_k - l/n)$  are asymptotically independent and  $N(0, f(\nu_k - l/n)/2)$ . Thus,  $\mathbf{E}\hat{f}(\nu^{(n)}) \to f(\nu)$  and  $\operatorname{Var}\hat{f}(\nu^{(n)}) \to f^2(\nu)/L$ .

Notice the *bias-variance trade off*:

**1.** Our assumption that f is approximately constant on  $[\nu - L/(2n), \nu + L/(2n)]$  becomes worse as L increases, so the difference between  $\hat{f}(\hat{\nu}^{(n)})$  and  $f(\nu)$  (the bias) will increase with L.

**2.** The variance of our estimate,  $\operatorname{Var}\hat{f}(\hat{\nu}^{(n)})$  decreases with L.

# **Other smoothed spectral estimators**

Instead of computing an unweighted average of the periodogram at all nearby frequencies, it is common to consider other weighted averages, typically with a smoother weighting function.

Consider the weighted average

$$\hat{f}(\nu) = \sum_{|j| \le L_n} W_n(j) I(\hat{\nu}^{(n)} - j/n),$$

where the bandwidth  $L_n$  is allowed to vary with n, and  $W_n$  is called the *spectral window function*.

### **Other smoothed spectral estimators**

For example, if

$$W_n(j) = \begin{cases} rac{1}{L} & ext{if } |j| < L/2, \\ 0 & ext{otherwise,} \end{cases}$$

then we have the smoothed spectral estimator we consider earlier,

$$\hat{f}(\nu) = \sum_{j} W_n(j) I(\hat{\nu}^{(n)} - j/n)$$
$$= \frac{1}{L} \sum_{|j| < L/2} I(\hat{\nu}^{(n)} - j/n).$$

This is Daniell's estimator (P. J. Daniell, University of Sheffield, 1946).

## **Consistency of nonparametric spectral estimation**

Suppose  $L_n$  and  $W_n$  satisfy

$$L_n \to \infty, \qquad \frac{L_n}{n} \to 0,$$
$$W_n(j) \ge 0, \qquad W_n(j) = W_n(-j), \qquad \sum_{|j| \le L_n} W_n(j) = 1,$$
$$\sum_{|j| \le L_n} W_n^2(j) \to 0$$

as  $n \to \infty$ , then...

### **Consistency of nonparametric spectral estimation**

... for a large class of stationary processes,  $\hat{f}(\nu) \to f(\nu)$  in the mean square sense. In particular,  $\mathbf{E}\hat{f}(\nu) \to f(\nu)$  and

$$\left(\sum_{|j| \le L_n} W_n^2(j)\right)^{-1} \operatorname{Cov}(\hat{f}(\nu_1), \hat{f}(\nu_2)) \to \begin{cases} f^2(\nu_1) & \text{if } \nu_1 = \nu_2 \in (0, 1/2) \\ 0 & \text{if } \nu_1 \neq \nu_2. \end{cases}$$

The conditions on the bandwidth parameter  $L_n$  ensure that, as the sample size grows, the window width goes to zero, but includes an infinite number of terms. The conditions on the spectral window function  $W_n$  ensure that the expectation of  $\hat{f}(\nu)$  converges to  $f(\nu)$  and its variance converges to zero.

# **Consistency of nonparametric spectral estimation**

$$\mathbf{E}\left(\hat{f}(\nu)\right) = \sum_{|j| \le L_n} W_n(j) \mathbf{E}\left(I(\hat{\nu}^{(n)} - j/n)\right)$$
$$\approx \sum_{|j| \le L_n} W_n(j) f(\nu) = f(\nu).$$

$$\begin{split} \operatorname{Var}\left(\hat{f}(\nu)\right) &= \sum_{j,k} W_n(j) W_n(k) \operatorname{Cov}\left(I(\hat{\nu}^{(n)} - j/n), I(\hat{\nu}^{(n)} - k/n)\right) \\ &\approx \sum_j W_n^2(j) \operatorname{Var}\left(I(\hat{\nu}^{(n)} - j/n)\right) \\ &\approx f^2(\nu) \sum_j W_n^2(j) \to 0. \end{split}$$

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# **Nonparametric spectral estimation: asymptotics**

Recall that for Daniell's estimator we have

$$\hat{f}(\nu_k) = \frac{1}{L} \sum_{l=-(L-1)/2}^{(L-1)/2} \left( X_c^2(\nu_k - l/n) + X_s^2(\nu_k - l/n) \right),$$

which is (asymptotically) a sum of 2L independent  $\chi_1^2$  random variables, so

$$\hat{f}(\nu_k) \sim f(\nu_k) \frac{\chi_{2L}^2}{2L}.$$

But for a non-uniform weighting, we form a weighted sum of these  $\chi_1^2$ random variables, so we cannot count up the degrees of freedom in the same way. But we can still approximate a general smoothed spectrum by  $\hat{f}(\nu_k) \sim c_k \chi_d^2$  for some  $c_k$  and d.

## **Nonparametric spectral estimation: asymptotics**

Suppose that  $\hat{f}(\nu_k) \sim c_k \chi_d^2$ . What values should  $c_k$  and d take? We have, for a suitable spectral window  $W_n$ ,

$$f(\nu_k) \approx \mathbf{E}\hat{f}(\nu_k) = c_k d,$$
$$f^2(\nu_k) \sum_{|j| \le L_n} W_n^2(j) \approx \operatorname{Var}\hat{f}(\nu_k) = 2c_k^2 d.$$

Thus, we get

$$c_k = \frac{f(\nu_k)}{d},$$
  

$$2c_k = f(\nu_k) \sum_{|j| \le L_n} W_n^2(j),$$
  

$$d = \frac{2}{\sum_{|j| \le L_n} W_n^2(j)}.$$

## **Nonparametric spectral estimation: asymptotics**

$$c_k = \frac{f(\nu_k)}{d},$$
$$d = \frac{2}{\sum_{|j| \le L_n} W_n^2(j)}$$

This d is often referred to as the *equivalent degrees of freedom* for a smoothed spectrum. Under suitable conditions (and for a slightly different definition of d), it can be shown that, asymptotically,

$$\hat{f}(\nu^{(n)}) \sim f(\nu) \frac{\chi_d^2}{d}.$$

## Nonparametric spectral estimation: the lag window

We can also view smoothing the spectrum in the frequency domain as smoothing in the time domain, via

$$\hat{f}(\nu) = \sum_{|j| \le L_n} w_n(j)\hat{\gamma}(j)e^{-2\pi i\nu j},$$

where  $w_n$  is the inverse Fourier transform of the spectral window. This is known as the *lag window*.

Tapering techniques are also popular: Define  $y_t = h_t x_t$  for some weighting function  $h_t$ . Then the tapered estimator is the smoothed spectral estimator for the tapered series  $y_t$ . For a weighting function  $h_t$  that smoothly diminishes values near the ends of the time series, we see less leakage.

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