

Weak Concentration for First Passage Percolation Times on Graphs and General Increasing Set-valued Processes

David J. Aldous *

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Abstract

A simple lemma bounds $\text{s.d.}(T)/\mathbb{E}T$ for hitting times T in Markov chains with a certain strong monotonicity property. We show how this lemma may be applied to several increasing set-valued processes. Our main result concerns a model of first passage percolation on a finite graph, where the traversal times of edges are independent Exponentials with arbitrary rates. Consider the percolation time X between two arbitrary vertices. We prove that $\text{s.d.}(X)/\mathbb{E}X$ is small if and only if $\Xi/\mathbb{E}X$ is small, where Ξ is the maximal edge-traversal time in the percolation path attaining X .

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1 Introduction

For random variables T arising (loosely speaking) as optimal costs in some “random environment” model, one might not be able to estimate $\mathbb{E}T$ explicitly, for one of two reasons: it may involve a difficult optimization problem (exemplified by the Euclidean TSP over random points [16]), or the model may involve many parameters (the case more relevant to this paper). In such cases the well-known *method of bounded differences* [14] often enables us to bound $T - \mathbb{E}T$ explicitly, and this general topic of *concentration inequalities*

*Department of Statistics, 367 Evans Hall # 3860, U.C. Berkeley CA 94720; aldous@stat.berkeley.edu; www.stat.berkeley.edu/users/aldous. Aldous’s research supported by N.S.F Grant DMS-106998.

has been developed in many directions over the last generation. What we will call (recalling the *weak* law of large numbers) a *weak concentration inequality* is just a result showing $\text{s.d.}(T)/\mathbb{E}T$ is small. The starting point for this paper is a simple technique (section 1.1) for proving weak concentration for hitting times T in Markov chains with a certain strong monotonicity property. At first sight one might doubt that this technique could be more than very narrowly applicable, but its quick use in three diverse contexts (sections 1.2 - 1.4) should partly dispel such doubts. Our main result concerns a more sophisticated analysis of a model of first passage percolation on a finite graph, where the traversal times of edges are independent Exponentials with arbitrary rates. We will describe that result and its background in section 2, and prove it in the subsequent sections.

1.1 A monotonicity condition

The setting is a continuous-time Markov chain (Z_t) on a finite state space Σ , where we study the hitting time

$$T := \inf\{t : Z_t \in \Sigma_0\} \tag{1}$$

for a fixed subset $\Sigma_0 \subset \Sigma$. Assume

$$h(S) := \mathbb{E}_S T < \infty \text{ for each } S \in \Sigma \tag{2}$$

which holds under the natural “reachability” condition. Assume also a rather strong “monotonicity” condition:

$$h(S') \leq h(S) \text{ whenever } S \rightarrow S' \text{ is a possible transition.} \tag{3}$$

In a typical example, the state space Σ will be the space of all subsets S of a given finite set \mathbf{V} , and possible transitions will be of the form $S \rightarrow S \cup \{v\}$. A special case of our simple lemma, which suffices for the quick applications in sections 1.2 - 1.4, is

Lemma 1 *Under condition (3), for any initial state,*

$$\frac{\text{var } T}{\mathbb{E}T} \leq \max\{h(S) - h(S') : S \rightarrow S' \text{ a possible transition}\}.$$

For a more general case we want to allow the possibility of occasional transitions for which $h(S) - h(S')$ is not so small.

Lemma 2 *In the setting of Lemma 1, for arbitrary $\delta, \varepsilon > 0$,*

$$\frac{\text{var } T}{(\mathbb{E}T)^2} \leq \delta + \varepsilon + \frac{\mathbb{E} \int_0^T \mathbb{1}_{\{q_\delta(Z_u) \geq \varepsilon\}} du}{\mathbb{E}T}.$$

where $q_\delta(S)$ is defined at (9) below.

The proofs below involve only very standard martingale analysis. We are not claiming any conceptual novelty in these results, but instead emphasize their applications later.

Proof of Lemma 1. Expectation is relative to some fixed initial state S_0 . Note that $T = \inf\{t : h(Z_t) = 0\}$, that $t \rightarrow h(Z_t)$ is decreasing and so for any reachable state S we have $h(S) \leq h(S_0) = \mathbb{E}T$, facts we use frequently without comment. Consider the martingale

$$M_t := \mathbb{E}(T|Z_t) = h(Z_{t \wedge T}) + t \wedge T. \quad (4)$$

The Doob-Meyer decomposition of M_t^2 into a martingale Q_t and a predictable process is clearly

$$M_t^2 - M_0^2 = Q_t + \int_0^t a(Z_u) du$$

where

$$a(S) := \sum_{S'} q(S, S') (h(S) - h(S'))^2 \quad (5)$$

where $q(S, S')$ are the transition rates. Taking expectation at $t = \infty$ gives

$$\text{var } T = \mathbb{E} \int_0^T a(Z_u) du. \quad (6)$$

On the other hand the martingale property (4) for $\mathbb{E}(T|Z_t)$ corresponds to the identity

$$b(S) := \sum_{S'} q(S, S')(h(S) - h(S')) = 1 \text{ while } S \cap \Sigma_0 = \emptyset \quad (7)$$

and therefore

$$\mathbb{E}T = \mathbb{E} \int_0^T b(Z_u) du. \quad (8)$$

Setting

$$\kappa = \max\{h(S) - h(S') : S \rightarrow S' \text{ a possible transition}\}$$

we clearly have $a(S) \leq \kappa b(S)$; note this is where we use the monotonicity hypothesis (3). The result now follows from (6,8).

Proof of Lemma 2. We continue with the notation above. Fix $\delta > 0$ and write, for a possible transition $S \rightarrow S'$,

$$h(S) - h(S') \leq \delta \mathbb{E}T + \mathbf{1}_{\{h(S) - h(S') > \delta \mathbb{E}T\}} \mathbb{E}T.$$

Using this to bound one term of the product $(h(S) - h(S'))^2$ in the definition (5) of $a(S)$, and comparing with the definition (7) of $b(S)$, we obtain

$$\begin{aligned} a(S) &\leq b(S) \cdot \delta \mathbb{E}T \\ &\quad + \sum_{S': h(S) - h(S') > \delta \mathbb{E}T} q(S, S')(h(S) - h(S')) \cdot \mathbb{E}T \end{aligned}$$

While $S \cap \Sigma_0 = \emptyset$ we have $b(S) = 1$ and so

$$\frac{a(S)}{\mathbb{E}T} \leq \delta + q_\delta(S)$$

where

$$q_\delta(S) := \sum_{S': h(S) - h(S') > \delta \mathbb{E}T} q(S, S')(h(S) - h(S')). \quad (9)$$

Using (6)

$$\frac{\text{var } T}{\mathbb{E}T} \leq \delta \mathbb{E}T + \mathbb{E} \int_0^T q_\delta(Z_u) du.$$

Because $q_\delta(S) \leq b(S) \equiv 1$ we can fix $\varepsilon > 0$ and write $q_\delta(Z_u) \leq \varepsilon + \mathbf{1}_{\{q_\delta(Z_u) \geq \varepsilon\}}$, and the result follows.

Discrete-time chains. Given a discrete-time Markov chain with transition probabilities $p(S, S')$ there is a corresponding continuous-time Markov chain with transition rates $q(S, S') = p(S, S')$. The relation between the hitting times T_{disc} and T_{cont} for these two chains is (using Wald's identity)

$$\mathbb{E}T_{\text{cont}} = \mathbb{E}T_{\text{disc}}; \quad \text{var } T_{\text{cont}} = \text{var } T_{\text{disc}} + \mathbb{E}T_{\text{disc}}. \quad (10)$$

Via this *continuization* device our results may be applied to discrete-time chains.

1.2 A general Markovian growth process

For a first “obvious” use of Lemma 1, we consider a general growth process (Z_t) on the lattice \mathbb{Z}^2 . The states are finite vertex-sets S , the possible transitions are $S \rightarrow S \cup \{v\}$ where v is a vertex adjacent to S . For each such transition, we assume the transition rates are bounded above and below:

$$0 < c_* \leq q(S, S \cup \{v\}) \leq c^* < \infty. \quad (11)$$

Initially $Z_0 = \{\mathbf{0}\}$, where $\mathbf{0}$ denotes the origin. The “monotonicity” condition we impose is that these rates are increasing in S :

$$\text{if } v, v' \text{ are adjacent to } S \text{ then } q(S, S \cup \{v\}) \leq q(S \cup \{v'\}, S \cup \{v, v'\}). \quad (12)$$

Note that we do not assume any kind of spatial homogeneity.

Proposition 3 *Let A be an arbitrary subset of vertices $\mathbb{Z}^2 \setminus \{\mathbf{0}\}$, and consider $T := \inf\{t : Z_t \cap A \text{ is non-empty}\}$. Under assumptions (11, 12),*

$$\text{var } T \leq \mathbb{E}T/c_*.$$

Proof. Condition (12) allows us to couple versions (Z'_t, Z''_t) of the process starting from states $S' \subset S''$, such that in the coupled process we have $Z'_t \subseteq Z''_t$ for all $t \geq 0$. In particular, $h(S) := \mathbb{E}_S T$ satisfies condition (3). To deduce the result from Lemma 1 we need to show that, for any given possible transition $S_0 \rightarrow S_0 \cup \{v_0\}$, we have

$$h(S_0) \leq h(S_0 \cup \{v_0\}) + 1/c_*. \quad (13)$$

Now by running the process started at S_0 until the first time T^* this process contain v_0 , and then coupling the future of that process to the process started at $S_0 \cup \{v_0\}$, we have $h(S_0) \leq \mathbb{E}_{S_0} T^* + h(S_0 \cup \{v_0\})$. And $\mathbb{E}_{S_0} T^* \leq 1/c_*$ by (11), establishing (13).

1.3 A multigraph process

Here is an example which we find more striking. Motivation for this kind of model will be described elsewhere.

Take a finite connected graph (\mathbf{V}, \mathbf{E}) with edge-weights $\mathbf{w} = (w_e)$, where $w_e > 0 \forall e \in \mathbf{E}$. Define a multigraph-valued process as follows. Initially we have the vertex-set \mathbf{V} and no edges. For each vertex-pair $e = (vy) \in \mathbf{E}$, edges vy appear at the times of a Poisson (rate w_e) process, independent over $e \in \mathbf{E}$. So at time t the state of the process, Z_t say, is a multigraph

with $N_e(t) \geq 0$ copies of edge e , where $(N_e(t), e \in \mathbf{E})$ are independent $\text{Poisson}(tw_e)$ random variables.

We study how long until Z_t has various connectivity properties. Specifically, consider

- $T'_k = \inf\{t : Z_t \text{ is } k\text{-edge-connected}\}$
- $T_k = \inf\{t : Z_t \text{ contains } k \text{ edge-disjoint spanning trees.}\}$

Here we regard the $N_e(t)$ copies of e as disjoint edges. Remarkably, Lemma 1 enables us to give a simple proof of a “weak concentration” bound which does not depend on the underlying weighted graph.

Proposition 4

$$\frac{\text{s.d.}(T_k)}{\mathbb{E}T_k} \leq \frac{1}{\sqrt{k}}, \quad k \geq 1.$$

Using the continuization device at (10), the same bound holds in the discrete-time model where edges e arrive IID with probabilities proportional to w_e .

We conjecture that some similar result holds for T'_k . But proving this by our methods would require some structure theory (beyond Menger’s theorem) for k -edge-connected graphs, and it is not clear whether relevant theory is known.

Proof of Proposition 4. We will apply Lemma 1. Here the states S are multigraphs over \mathbf{V} , and $h(S)$ is the expectation, starting at S , of the time until the process contains k edge-disjoint spanning trees. What are the possible values of $h(S) - h(S \cup \{e\})$, where $S \cup \{e\}$ denotes the result of adding an extra copy of e to the multigraph S ?

Consider the “min-cut” over proper subsets $S \subset \mathbf{V}$

$$\gamma := \min_S w(S, S^c)$$

where $w(S, S^c) = \sum_{v \in S, y \in S^c} w_{vy}$. Because a spanning tree must have at least one edge across the min-cut,

$$\mathbb{E}T_k \geq k/\gamma. \tag{14}$$

On the other hand we claim

$$h(S) - h(S \cup \{e\}) \leq 1/\gamma. \tag{15}$$

To prove this, take the natural coupling (Z_t, Z_t^+) of the processes started from S and from $S \cup \{e\}$, and run the coupled process until Z_t^+ contains k

edge-disjoint spanning trees. At this time, the process Z_t either contains k edge-disjoint spanning trees, or else contains $k - 1$ spanning trees plus two trees (regard as edge-sets \mathbf{t}_1 and \mathbf{t}_2) such that $\mathbf{t}_1 \cup \mathbf{t}_2 \cup \{e\}$ is a spanning tree. So the extra time we need to run (Z_t) is at most the time until some arriving edge links \mathbf{t}_1 and \mathbf{t}_2 , which has mean at most $1/\gamma$. This establishes (15), and then Lemma 1 establishes the proposition.

1.4 Coverage processes

The topic of *coverage processes* is centered upon spatial or combinatorial variants of the coupon collector’s problem; see the monograph [10] and scattered examples in [1]. Classical theory concerns low-parameter models for which the cover time T_n of a “size n ” model can be shown to have a limit distribution after scaling: $(T_n - a_n)/b_n \rightarrow_d \xi$ for explicit a_n, b_n . In many settings, Lemma 1 can be used to give a weak concentration result for models with much less regular structure. Here is a very simple example, whose one-line proof is left to the reader.

Proposition 5 *Let G be an arbitrary n -vertex graph, and let $(V_i, i \geq 1)$ be IID uniform random vertices. Let T be the smallest t such that every vertex is contained in, or adjacent to, the set $\{V_i, 1 \leq i \leq t\}$. Then $\text{var } T \leq n \mathbb{E}T$.*

For a sequence (G_n) of sparse graphs, $\mathbb{E}T_n$ will be of order $n \log n$, so the bound says that $\text{s.d.}(T_n)/\mathbb{E}T = O(1/\sqrt{\log n})$.

2 The FPP model

As in section 1.3 we start with a finite connected graph (\mathbf{V}, \mathbf{E}) with edge-weights $\mathbf{w} = (w_e)$, where $w_e > 0 \forall e \in \mathbf{E}$, but the model here is different. To the edges $e \in \mathbf{E}$ attach independent Exponential(rate w_e) random variables ξ_e . For each pair of vertices (v', v'') there is a random variable $X(v', v'')$ which can be viewed in two equivalent ways:

viewing ξ_e as the length of edge e , then $X(v', v'')$ is the length of the shortest route from v to v' ;

viewing ξ_e as the time to traverse edge e , then $X(v', v'')$ is the “first passage percolation” time from v to v' .

Taking the latter view, we call this the *FPP model*, and call $X(v', v'')$ the *FPP time* and ξ_e the *traversal time*. This type of model and many generalizations have been studied extensively in several settings, in particular

- FPP with general IID weights on \mathbb{Z}^d [12]

- FPP on classical random graph (Erdős-Rényi or configuration) models [5, 6]
- and a much broader “epidemics and rumors on complex networks” literature [9].

However, this literature invariably starts by assuming some specific graph model; we do not know any “general results” which relate properties of the FPP model to properties of a general underlying graph. As an analogy, another structure that can be associated with a weighted finite graph is a finite reversible Markov chain; the established theory surrounding mixing times of Markov chains [3, 13, 15] does contain “general results” relating properties of the chain to properties of the underlying graph.

This article makes a modest start by studying the “weak concentration” property: when it is true that $X(v', v'')$ is close to its expectation? We can reformulate the FPP model as a set-valued process (see section 2.1(a) for details) and then Lemma 1 immediately implies the following result.

Proposition 6 $\text{var } X(v', v'') \leq \mathbb{E}X(v', v'')/w_*$ for $w_* := \min\{w_e : e \in \mathbf{E}\}$.

Bounds of this type are classical on \mathbb{Z}^d [11],

Proposition 6 implies that on any unweighted graph ($w_e = 1$ for all edges e), the spread of $X = X(v', v'')$ is at most order $\sqrt{\mathbb{E}X}$. For many specific graphs, stronger concentration results are known. For \mathbb{Z}^2 there is extensive literature (see [8] for a recent overview) on the longstanding conjecture that the spread is order $(\mathbb{E}X)^{1/3}$. For the complete graph (e.g. [2] sec. 7.3) and for sparse random graphs on n vertices [5] the spread of $X/\mathbb{E}X$ is typically of order $1/\log n$.

In contrast, in this paper we study the completely general case where the edge-weights w_e may vary widely over the different edges $e \in \mathbf{E}$. Here is our main result (conjectured in [2] sec. 7.4). Given a pair (v', v'') , there is a random path $\pi(v', v'')$ that attains the FPP time $X(v', v'')$. Define $\Xi(v', v'') := \max\{\xi_e : e \in \pi(v', v'')\}$ as the maximum edge-traversal time in this minimal path. Recall the “ L^0 norm”

$$\|V\|_0 := \inf\{\delta : \mathbb{P}(|V| > \delta) \leq \delta\}.$$

Theorem 7 *There exist functions ψ^+ and $\psi_- : (0, 1] \rightarrow (0, \infty)$ such that $\psi^+(\delta) \downarrow 0$ as $\delta \downarrow 0$, and $\psi_-(\delta) > 0$ for all $\delta > 0$, and such that, for all finite connected edge-weighted graphs and all vertex pairs (v', v'') ,*

$$\psi_- \left(\left\| \frac{\Xi(v', v'')}{\mathbb{E}X(v', v'')} \right\|_0 \right) \leq \frac{\text{s.d.}(X(v', v''))}{\mathbb{E}X(v', v'')} \leq \psi^+ \left(\left\| \frac{\Xi(v', v'')}{\mathbb{E}X(v', v'')} \right\|_0 \right).$$

In words, $X/\mathbb{E}X$ has small spread if and only if $\Xi/\mathbb{E}X$ is small. Intuition for this result comes from the “almost disconnected” case where the path $\pi(v', v'')$ must contain a specific “bridge” edge e with small w_e ; if $1/w_e$ is not $o(\mathbb{E}X)$ then the contribution to X from the traversal time ξ_e is enough to show that X cannot have weakly concentrated distribution.

The proof of the upper bound starts from our “more general” inequality, Lemma 2. See section 3 for an outline.

Theorem 7 is unsatisfactory in that the conditions are not directly on the edge-weights $\mathbf{w} = (w_e)$. By analogy with the bounds ([15] sec. 3.2) for Markov chain mixing times in terms of the *spectral profile*, it seems likely that Proposition 6 can be extended to give more widely applicable upper bounds on spread in terms of extremal flow rates $w(S, S^c) = \sum_{v \in S, y \in S^c} w_{vy}$. However we do not have any conjecture for two-sided bounds analogous to Theorem 7.

A particular case of FPP in our setting – edge-weights w_e varying widely over the different edges – is studied in detail in recent work of Chatterjee-Dey [7]. In their model $\mathbf{V} = \mathbb{Z}^d$ with $w_{xy} = \|y - x\|^{-\alpha+(1)}$. Results and conjectures from that paper are consistent with our Theorem 7, which says that properties $X(\mathbf{0}, nz)/\mathbb{E}X(\mathbf{0}, nz) \rightarrow_p 1$ and $\Xi(\mathbf{0}, nz)/\mathbb{E}X(\mathbf{0}, nz) \rightarrow_p 1$ must either both hold or both fail. They identify several qualitatively different regimes. In their *linear growth regime* ($\alpha > 2d + 1$) they show $\mathbb{E}X(\mathbf{0}, nz)/n$ converges to a nonzero constant and show that both properties hold. In their *super-linear growth regime* ($\alpha \in (2d, 2d + 1)$) they show $X(\mathbf{0}, nz) = n^{\alpha-2d+o(1)}$; here their analysis suggests both properties fail. For $\alpha \in (d, 2d)$ they show $X(\mathbf{0}, nz)$ grows as a power of $\log n$, and their arguments suggest both properties hold. The qualitative behavior and proof techniques in [7] are different in these different regimes, whereas our Theorem 7 is a single result covering all regimes, albeit a less explicit result.

2.1 Some preliminaries

(a) We will view the FPP process started at v' as a process (Z_t) taking values in the (finite) space of subsets $S \subseteq \mathbf{V}$ of vertices, that is as

$$Z_t := \{v : X(v', v) \leq t\}. \quad (16)$$

The assumption of Exponential distributions implies that (Z_t) is the continuous-time Markov chain with $Z_0 = \{v'\}$ and transition rates

$$S \rightarrow S \cup \{y\} : \quad \text{rate } w(S, y) := \sum_{s \in S} w_{sy} \quad (y \notin S).$$

So we are in the setting of Lemmas 1 and 2. Given a target vertex v'' the FPP time $X(v', v'')$ is a stopping time T of the form (1), and the function $h(S)$ at (2) is just the function

$$h(S) := \mathbb{E} \min_{v \in S} X(v, v'').$$

For a possible transition $S \rightarrow S \cup \{y\}$, by considering the first time y is reached in the process started from S we have $h(S) \leq \frac{1}{w(S,y)} + h(S \cup \{y\})$ and so

$$h(S) - h(S \cup \{y\}) \leq \frac{1}{w(S,y)} \leq \frac{1}{w_*}$$

for $w_* := \min\{w_e : e \in \mathbf{E}\}$. Now Lemma 1 does indeed imply Proposition 6, as stated earlier.

Presumably some analog of Theorem 7 remains true when the Exponential assumption is relaxed to some much weaker “shape of distribution” assumption, but we have not investigated that setting.

(b) FPP times such as $X(v', v'')$ are examples of distributions $Y > 0$ with the *submultiplicative* property

$$\mathbb{P}(Y > y_1 + y_2) \leq \mathbb{P}(Y > y_1) \mathbb{P}(Y > y_2); \quad y_1, y_2 > 0.$$

We will need the general bound provided by the following straightforward lemma.

Lemma 8 *There exists a function $\gamma(u) > 0$ with $\gamma(u) \downarrow 0$ as $u \downarrow 0$ such that, for every submultiplicative Y and every event A ,*

$$\frac{\mathbb{E}Y \mathbf{1}_A}{\mathbb{E}Y} \leq \gamma(\mathbb{P}(A)).$$

Note it follows that

$$\mathbb{P}(Y \geq y) \geq \gamma^{-1} \left(\frac{\mathbb{E}Y \mathbf{1}_{\{Y \geq y\}}}{\mathbb{E}Y} \right) \tag{17}$$

for the inverse function $\gamma^{-1}(u) \downarrow 0$ as $u \downarrow 0$.

(c) We will need an elementary stochastic calculus lemma.

Lemma 9 *Let T_1 be a stopping time with (random) intensity η_t – that is, $\mathbb{P}(t < T_1 < t + dt | \mathcal{F}_t) = \eta_t dt$ on $\{T_1 > t\}$. Let ζ be another stopping time such that $\eta_t \geq c$ on $\{\zeta > t\}$, for constant $c > 0$. Then*

$$\mathbb{P}(T_1 \leq \zeta \wedge t_0) \geq (1 - e^{-ct_0}) \mathbb{P}(\zeta > t_0).$$

Proof. Applying the optional sampling theorem to the martingale $\exp(\int_0^t \eta_s ds) \cdot \mathbf{1}_{\{T > t\}}$ and the stopping time $\zeta \wedge t_0$ shows

$$\mathbb{E}[\exp(\int_0^{\zeta \wedge t_0} \eta_s ds) \cdot \mathbf{1}_{\{T > \zeta \wedge t_0\}}] = 1.$$

Now

$$\exp(\int_0^{\zeta \wedge t_0} \eta_s ds) \geq \exp(c(\zeta \wedge t_0)) \geq 1 + (e^{ct_0} - 1)\mathbf{1}_{\{\zeta > t_0\}}$$

and so

$$\mathbb{P}(T_1 > \zeta \wedge t_0) + (e^{ct_0} - 1)\mathbb{P}(T_1 > \zeta \wedge t_0, \zeta > t_0) \leq 1.$$

That is,

$$\begin{aligned} \mathbb{P}(T_1 \leq \zeta \wedge t_0) &\geq (e^{ct_0} - 1)\mathbb{P}(T_1 > \zeta \wedge t_0, \zeta > t_0) \\ &\geq (e^{ct_0} - 1)(\mathbb{P}(\zeta > t_0) - \mathbb{P}(T_1 \leq \zeta \wedge t_0)) \end{aligned}$$

which rearranges to the stated inequality.

3 Proof of upper bound in Theorem 7

Fix v' and as above view the FPP process started at v' as the continuous-time Markov chain (Z_t) at (16) on the space of subsets of \mathbf{V} . Fix a target $v'' \neq v'$ and in the following write S for an arbitrary subset of vertices containing v' . Write

$$h(S) := \mathbb{E} \min_{v \in S} X(v, v'')$$

for the mean percolation time from S to v'' , so $h(S) = 0$ iff $v'' \in S$. So

$$T := X(v', v'') = \inf\{t : h(Z_t) = 0\}$$

is a stopping time for (Z_t) . Note that $t \rightarrow h(Z_t)$ is decreasing and $h(S) \leq \mathbb{E}T$.

Outline of proof. Step 1 is to translate Lemma 2 into our FPP setting; this shows it is enough to prove that in all transitions $S \rightarrow S \cup \{y\}$ the decrements $h(S) - h(S \cup \{y\})$ are $o(\mathbb{E}T)$; so suppose not, that is suppose some are $\Omega(\mathbb{E}T)$. Step 2 shows that in some such transitions, the used edge vy will (with non-vanishing probability) have traversal time ξ_{vy} also of order $\Omega(\mathbb{E}T)$. Step 3 shows that some such edges will (with non-vanishing probability) be in the minimal path.

Details of proof. Step 1. Substituting 2δ for δ , Lemma 2 says that for arbitrary $\delta, \varepsilon > 0$

$$\frac{\text{var } T}{(\mathbb{E}T)^2} \leq 2\delta + \varepsilon + \frac{\mathbb{E} \int_0^T \mathbf{1}_{\{q_\delta(Z_u) \geq \varepsilon\}} du}{\mathbb{E}T}. \quad (18)$$

where

$$q_\delta(S) := \sum_{y: h(S) - h(S \cup \{y\}) > 2\delta \mathbb{E}T} w(S, y)(h(S) - h(S \cup \{y\})).$$

Informally, as stated in the *outline of proof*, this shows it will suffice to prove that in transitions $S \rightarrow S \cup \{y\}$ the decrements $h(S) - h(S \cup \{y\})$ are $o(\mathbb{E}T)$.

Step 2. Now fix $0 < u_1 < u_2$.

Lemma 10 *Condition on $Z_{t_0} = S_0$. The event*

during $[t_0, t_0 + u_2]$ the process (Z_{t_0+u}) makes a transition $S \rightarrow S \cup \{y\}$ such that $h(S) - h(S \cup \{y\}) > \delta \mathbb{E}T$ and using an edge vy for which $\xi_{vy} > u_1$

has probability at least $(1 - \frac{u_2}{\delta \mathbb{E}T})^+ \left(1 - \exp(-\frac{(u_2 - u_1)q(S_0)}{2\mathbb{E}T})\right)$.

Proof. Define

$$\zeta = \inf\{t > t_0 : h(Z_t) \leq h(S_0) - \delta \mathbb{E}T\}$$

so that, because $t \rightarrow h(Z_t)$ is decreasing,

$$h(S_0) - h(Z_{t_0+u}) \leq \delta \mathbb{E}T \text{ on } \{\zeta > t_0 + u\}. \quad (19)$$

Define

$$q^*(S) := \sum_{y: h(S) - h(S \cup \{y\}) > \delta \mathbb{E}T} w(S_0, y)(h(S) - h(S \cup \{y\})).$$

Note this is the definition of $q_\delta(S)$ modified by changing the constraint “ $> 2\delta \mathbb{E}T$ ” to “ $> \delta \mathbb{E}T$ ” and changing $w(S, y)$ to $w(S_0, y)$. Observe the relation, for $Z_{t_0+u} \supseteq S_0$,

$$\begin{aligned} h(Z_{t_0+u}) - h(Z_{t_0+u} \cup \{y\}) &\geq h(Z_{t_0+u}) - h(S_0 \cup \{y\}) \\ &= [h(S_0) - h(S_0 \cup \{y\})] - [h(S_0) - h(Z_{t_0+u})]. \end{aligned} \quad (20)$$

For y satisfying the constraint

$$h(S_0) - h(S_0 \cup \{y\}) > 2\delta\mathbb{E}T$$

in the definition of $q(S_0)$, and on the event $\{\zeta > t_0 + u\}$, inequalities (19,20) show that y satisfies the constraint in the definition of $q^*(Z_{t_0+u})$, and also show that

$$h(Z_{t_0+u}) - h(Z_{t_0+u} \cup \{y\}) \geq \frac{1}{2}[h(S_0) - h(S_0 \cup \{y\})],$$

implying that

$$q^*(Z_{t_0+u}) \geq \frac{1}{2}q(S_0) \text{ on } \{\zeta > t_0 + u\}. \quad (21)$$

For $S \supseteq S_0$ we have, from the definition of $q^*(S)$, a crude bound

$$q^*(S) \leq (\mathbb{E}T) \sum_{y: h(S) - h(S \cup \{y\}) > \delta\mathbb{E}T} w(S_0, y)$$

and applying this to the left side of (21) gives

$$\sum_{y: h(Z_{t_0+u}) - h(Z_{t_0+u} \cup \{y\}) > \delta\mathbb{E}T} w(S_0, y) \geq \frac{q(S_0)}{2\mathbb{E}T} \text{ on } \{\zeta > t_0 + u\}.$$

Over the time interval $[t_0 + u_1, t_0 + u_2]$ the left side is the intensity of an event which implies the event $(D, \text{ say})$ in Lemma 10 (in particular, an edge vy is used with $v \in S_0 = Z(t_0)$, and so its ‘‘age’’ ξ_{vy} must be at least u_1). We are now in the setting of Lemma 9, which shows that

$$\mathbb{P}(D) \geq \mathbb{P}(\zeta > t_0 + u) \left(1 - \exp\left(\frac{-(u_2 - u_1)q(S_0)}{2\mathbb{E}T}\right)\right).$$

By the martingale property (4) of $h(Z_t) + t$ and Markov’s inequality

$$\mathbb{P}(\zeta \leq t_0 + u) = \mathbb{P}(h(Z_{t_0+u_2}) \leq h(S_0) - \delta\mathbb{E}T) \leq \frac{u_2}{\delta\mathbb{E}T} \quad (22)$$

establishing the bound stated in Lemma 10.

Step 3.

Lemma 11 *Conditional on the process (Z_u) making at time t_0 a transition $S \rightarrow S \cup \{y\}$ using an edge vy , the probability that edge vy is in the minimal path $\pi(v', v'')$ is at least $\gamma^{-1}\left(\frac{h(S) - h(S \cup \{y\})}{\mathbb{E}T}\right)$, for the inverse function $\gamma^{-1}(\cdot) > 0$ at (17).*

Proof. Condition as stated. So after the transition we have

$$T - t_0 = \min(A, B)$$

where $B = X(y, v'')$ and $A = \min_{v \in S} X^*(v, v'')$ where X^* denotes the minimum over paths not using y . The probability in question equals $\mathbb{P}(B < A)$; note that A and B are typically dependent.

It is easy to check that the solution to the problem

given the distribution of $A > 0$, construct a r.v. $B' > 0$ to minimize $\mathbb{P}(B' < A)$ subject to the constraint that $\mathbb{E}A - \mathbb{E} \min(A, B')$ takes a given value

is to take $B' = A \mathbf{1}_{\{A \leq a_0\}}$ for a_0 chosen to satisfy the constraint. So in our setting,

$$\mathbb{P}(B < A) \geq \mathbb{P}(A > a_0)$$

for a_0 chosen to satisfy

$$\mathbb{E}A \mathbf{1}_{\{A > a_0\}} = \mathbb{E}A - \mathbb{E} \min(A, B) = h(S) - h(S \cup \{y\}).$$

Because A has the submultiplicative property, Lemma 8 shows that

$$\mathbb{P}(A > a_0) \geq \gamma^{-1} \left(\frac{h(S) - h(S \cup \{y\})}{\mathbb{E}A} \right).$$

Because $\gamma^{-1}(\cdot)$ is decreasing and $\mathbb{E}A \leq \mathbb{E}T$, we have established the lemma.

Step 4. We now combine the ingredients above. Define $U_1 = \inf\{t \geq 0 : q(Z_t) \geq \varepsilon\}$ and inductively

$$U_{j+1} = \inf\{t \geq U_j + u_2 : q(Z_t) \geq \varepsilon\}.$$

Note that the condition

$$ju_2 \leq \int_0^T \mathbf{1}_{\{q(Z_u) \geq \varepsilon\}} du \tag{23}$$

is sufficient to ensure $U_j < \infty$. Condition on $U_j = t_0, Z_{U_j} = S_0$ and apply Lemmas 10 and 11. We deduce that, with probability at least

$$\left(1 - \frac{u_2}{\delta \mathbb{E}T}\right) \left(1 - \exp\left(\frac{-(u_2 - u_1)\varepsilon}{2\mathbb{E}T}\right)\right) \gamma^{-1}(\delta)$$

the minimal path $\pi(v', v'')$ contains an edge vy with $\xi_{vy} \geq u_1$ and such that y is first reached during $(U_j, U_{j+1}]$. The latter property ensures these

edges are distinct as j varies (but note the corresponding ξ_{vy} are dependent). Summing over j , applying (23) and taking expectation, we find that

$$N(u_1) := |\{e \in \pi(v', v'') : \xi_e \geq u_1\}|$$

satisfies

$$\mathbb{E}N(u_1) \geq \left(1 - \frac{u_2}{\delta \mathbb{E}T}\right)^+ \left(1 - \exp\left(\frac{-(u_2 - u_1)\varepsilon}{2\mathbb{E}T}\right)\right) \gamma^{-1}(\delta) \frac{1}{u_2} \mathbb{E} \int_0^T \mathbf{1}_{\{q(Z_u) \geq \varepsilon\}} du. \quad (24)$$

Writing $\Xi = \Xi(v', v'')$, the event $\{N(u_1) \geq 1\}$ is the event $\{\Xi \geq u_1\}$. Also $u_1 N(u_1) \leq T$, so we can write $N(u_1) \leq \frac{T}{u_1} \mathbf{1}_{\{\Xi \geq u_1\}}$ and then

$$\mathbb{E}N(u_1) \leq \frac{\mathbb{E}T}{u_1} \frac{\mathbb{E}T \mathbf{1}_{\{\Xi \geq u_1\}}}{\mathbb{E}T} \leq \frac{\mathbb{E}T}{u_1} \gamma(\mathbb{P}(\Xi > u_1)),$$

the final inequality by Lemma 8. Combining with (24) and rearranging gives a bound for the final term of (18):

$$\frac{\mathbb{E} \int_0^T \mathbf{1}_{\{q(Z_u) \geq \varepsilon\}} du}{\mathbb{E}T} \leq \frac{u_2 \gamma(\mathbb{P}(\Xi > u_1))}{u_1 \left(1 - \frac{u_2}{\delta \mathbb{E}T}\right)^+ \left(1 - \exp\left(\frac{-(u_2 - u_1)\varepsilon}{2\mathbb{E}T}\right)\right) \gamma^{-1}(\delta)}.$$

This becomes a little easier to understand when we set $u_1 = v\mathbb{E}T$ and $u_2 = 2u_1$; then (18) becomes

$$\frac{\text{var } T}{(\mathbb{E}T)^2} \leq 2\delta + \varepsilon + \frac{2 \gamma(\mathbb{P}(\frac{\Xi}{\mathbb{E}T} > v))}{\left(1 - \frac{v}{\delta}\right)^+ \left(1 - \exp\left(\frac{-v\varepsilon}{2}\right)\right) \gamma^{-1}(\delta)}. \quad (25)$$

We could manipulate this into a complicated expression for an upper bound function $\psi^+(\cdot)$ in Theorem 7. But it is simpler to observe that existence of *some* upper bound function is equivalent to the following limit assertion: for weighted graphs $\mathbf{w}^{(n)}$ and FPP times $T^{(n)} = X^{(n)}(v^{(n)'}, v^{(n)'})$

$$\text{if } \frac{\Xi^{(n)}(v^{(n)'}, v^{(n)'})}{\mathbb{E}T^{(n)}} \rightarrow_p 0 \text{ then } \frac{\text{var } T^{(n)}}{(\mathbb{E}T^{(n)})^2} \rightarrow 0.$$

And this is immediate from (25).

4 Proof of lower bound in Theorem 7

We described this lower bound assertion in [2] sec. 7.4 as “intuitively clear (and not hard to prove)”. The *intuition* is that, given a typical realization

of the process where $X = X(v', v'')$ has some value X_0 , and some value Ξ_0 attained by some ξ_e , there are other realizations for which this ξ_e ranges over the interval $[0, \Xi_0]$, for which therefore the value of X ranges over $[X_0 - \Xi_0, X_0]$, so the order of $\text{s.d.}(X)$ should be at least the order of Ξ . Formalizing this intuition is not quite trivial. The proof below is rather crude – likely there is a more elegant argument giving a better bound.

Proof. Fix a pair (v', v'') and write $X = X(v', v'')$ and $\Xi = \Xi(v', v'')$. So X is a function of the traversal times (ξ_e) . If we couple (ξ_e) with a copy (ξ'_e) of (ξ_e) , then by an elementary inequality

$$\text{var } X \geq \frac{1}{4} \mathbb{E}(X' - X)^2. \quad (26)$$

We will use the following coupling. Fix $0 < a < b$ and write $\mathcal{E}(\lambda; a, b)$ for the distribution of an $\text{Exponential}(\lambda)$ random variable conditioned to lie in $[a, b]$. Define (ξ'_e, ξ_e) to be independent as e varies, with

if $\xi_e \notin [a, b]$ then $\xi'_e = \xi_e$

if $\xi_e \in [a, b]$ then ξ'_e is conditionally independent of ξ_e with conditional distribution $\mathcal{E}(w_e; a, b)$.

Consider the set D_{ab} of edges e in the minimal path $\boldsymbol{\pi} = \boldsymbol{\pi}(v', v'')$ for which $\xi_e \in [a, b]$. By considering the same path in the coupled copy,

$$X' - X \leq \sum_{e \in \boldsymbol{\pi}} (\xi'_e - \xi_e) = \sum_{e \in D_{ab}} (\xi'_e - \xi_e). \quad (27)$$

To analyze this expression we will use the following lemma. For $k \geq 1$ and $s > 0$ define

$$F_k(s) := \mathbb{E} \left(\max(0, s - \sum_{i=1}^k U_i) \right)^2 \quad (28)$$

where the U_i are independent $\text{Uniform}(0, 1)$.

Lemma 12 *Let $(V_i, 1 \leq i \leq k)$ be independent $\mathcal{E}(\lambda_i; a, b)$, for arbitrary (λ_i) . Then for arbitrary $v_i \in [a, b]$,*

$$\mathbb{E} \left(\max \left(0, \sum_{i=1}^k v_i - \sum_{i=1}^k V_i \right) \right)^2 \geq (b-a)^2 F_k \left(\sum_{i=1}^k \frac{v_i - a}{b-a} \right). \quad (29)$$

Proof. The distributions $\mathcal{E}(\lambda_i; a, b)$ are stochastically increasing with λ , so the left side is minimized in the $\lambda_i \rightarrow \infty$ limit, which is the uniform distribution on $[a, b]$; the limit value is obtained from (28) by scaling.

We can now combine inequalities (27) and (29) to deduce that, on the event $\{|D_{ab}| = k\}$ for $k \geq 1$,

$$\mathbb{E}((X - X')^2 | \xi_e, e \in \mathbf{E}) \geq (b - a)^2 F_k \left(\sum_{e \in D_{ab}} \frac{\xi_e - a}{b - a} \right).$$

So

$$\mathbb{E}(X - X')^2 \geq (b - a)^2 \sum_{k \geq 1} \mathbb{E} \left[\mathbf{1}_{\{|D_{ab}|=k\}} F_k \left(\sum_{e \in D_{ab}} \frac{\xi_e - a}{b - a} \right) \right]. \quad (30)$$

We need to lower bound the right side. The issue is that, if w_e is large then $\xi_e - a$ may be small. We handle this issue by considering two different values of a . Note that the functions $F_k(s)$ are decreasing in k and increasing in s . Fix $a_1 < a_2 < b$, so that for $1 \leq k \leq K$

$$\begin{aligned} F_k \left(\sum_{e \in D_{a_1 b}} \frac{\xi_e - a_1}{b - a_1} \right) &\geq F_K \left(\sum_{e \in D_{a_1 b}} \frac{\xi_e - a_1}{b - a_1} \right) \quad (\text{decreasing in } k) \\ &\geq F_K \left(\sum_{e \in D_{a_2 b}} \frac{\xi_e - a_1}{b - a_1} \right) \quad (D_{a_2 b} \subseteq D_{a_1 b}) \\ &\geq F_K \left(|D_{a_2 b}| \frac{a_2 - a_1}{b - a_1} \right) \quad (\text{increasing in } s) \\ &\geq F_K \left(\frac{a_2 - a_1}{b - a_1} \right) \mathbf{1}_{\{|D_{a_2 b}| \geq 1\}} \quad (\text{increasing in } s). \end{aligned}$$

Applying (30) with $a = a_1$ and restricting the sum to $1 \leq k \leq K$,

$$\mathbb{E}(X - X')^2 \geq (b - a_1)^2 F_K \left(\frac{a_2 - a_1}{b - a_1} \right) \mathbb{P}(|D_{a_1 b}| \leq K, |D_{a_2 b}| \geq 1). \quad (31)$$

Because $a_1 |D_{a_1 b}| \leq X$, Markov's inequality tells us $\mathbb{P}(|D_{a_1 b}| \geq K) \leq \frac{1}{a_1 K} \mathbb{E}X$. Also the event $\{|D_{a_2 b}| \geq 1\}$ contains the event $\{a_1 \leq \Xi \leq b\}$, so

$$\mathbb{P}(|D_{a_1 b}| \leq K, |D_{a_2 b}| \geq 1) \geq \mathbb{P}(\Xi \geq a_1) - \mathbb{P}(\Xi \geq b) - \frac{1}{a_1 K} \mathbb{E}X.$$

Because $\Xi \leq X$ we have $\mathbb{P}(\Xi \geq b) \leq \mathbb{E}X/b$, and combining with (31) and (26) we conclude

$$\text{var } X \geq \frac{1}{4}(b - a_1)^2 F_K \left(\frac{a_2 - a_1}{b - a_1} \right) \left(\mathbb{P}(\Xi \geq a_1) - \left(\frac{1}{b} + \frac{1}{a_1 K} \right) \mathbb{E}X \right)^+.$$

This holds for arbitrary $0 < a_1 < a_2 < b$ and $K \geq 1$. We now take $1 > \delta > 0$ and choose

$$a_1 = \delta \mathbb{E}X; \quad a_2 = 2a_1; \quad b = 3\mathbb{E}X/\delta$$

and find

$$\frac{\text{var } X}{(\mathbb{E}X)^2} \geq \frac{1}{4} \left(\frac{3}{\delta} - \delta \right)^2 F_K \left(\frac{\delta^2}{3 - \delta^2} \right) \left(\mathbb{P} \left(\frac{\Xi}{\mathbb{E}X} \geq \delta \right) - \left(\frac{\delta}{3} + \frac{1}{\delta K} \right) \right)^+.$$

So finally choosing $K = K(\delta) \geq 3\delta^{-2}$, we see that the lower bound in Theorem 7 holds for

$$\psi_-(\delta) := \sqrt{\frac{1}{4} \left(\frac{3}{\delta} - \delta \right)^2 F_K \left(\frac{\delta^2}{3 - \delta^2} \right) \frac{\delta}{3}}.$$

5 Final remarks

(a). The key point of results such as Proposition 4 and Theorem 7 is that the bounds do not depend on the size n of the graph – any simple use of the method of bounded differences in these models would give bounds that did depend on n .

(b). A somewhat different general approach to proving weak concentration for general coverage processes, assuming IID random subsets, was given in [4], and used to prove weak concentration for Markov chain cover times. The latter result does not seem to follow easily from the methods in this paper.

(c). Theorem 7 went beyond Proposition 6 by using Lemma 2 instead of the simpler Lemma 1, and so one can imagine analogous improvements of the kind of results in sections 1.2 - 1.4.

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