

Measurement Error in Spatial Modeling of Environmental Exposures

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Outline

- Spatial exposure estimation and environmental epidemiology
- Spatial modelling of exposure
- Prediction-induced measurement error
- Methods for accounting for measurement error
- Simulation results

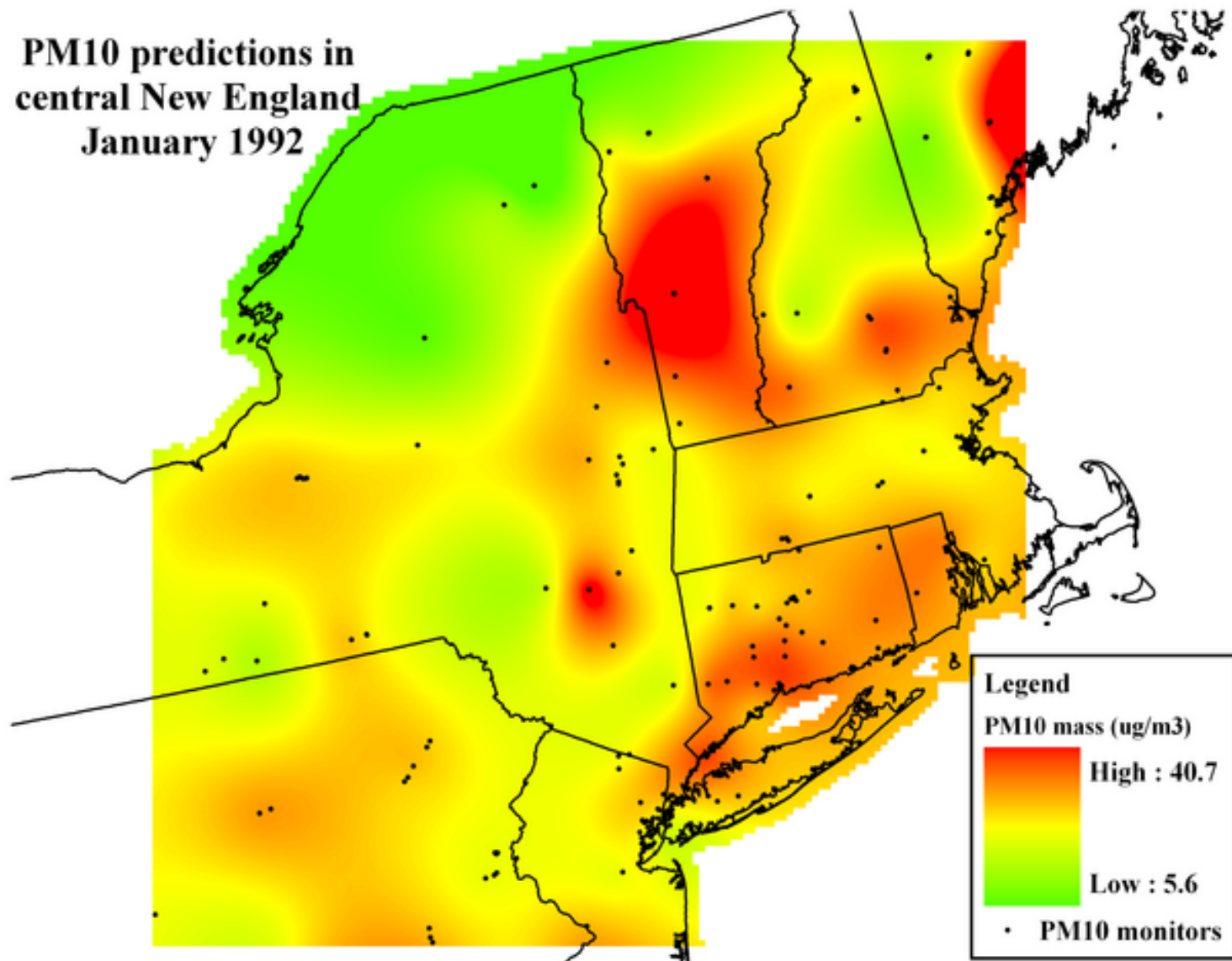
Spatial exposure estimation in environmental health

1. retrieval of spatio-temporal data from monitoring networks or site visits
2. space-time modelling, plus use of GIS-derived covariates
3. prediction at locations of individuals in health study (e.g., large cohort study)
4. epidemiological investigation with exposure predictions as a covariate

Cardiopulmonary disease in the Nurses' Health Study

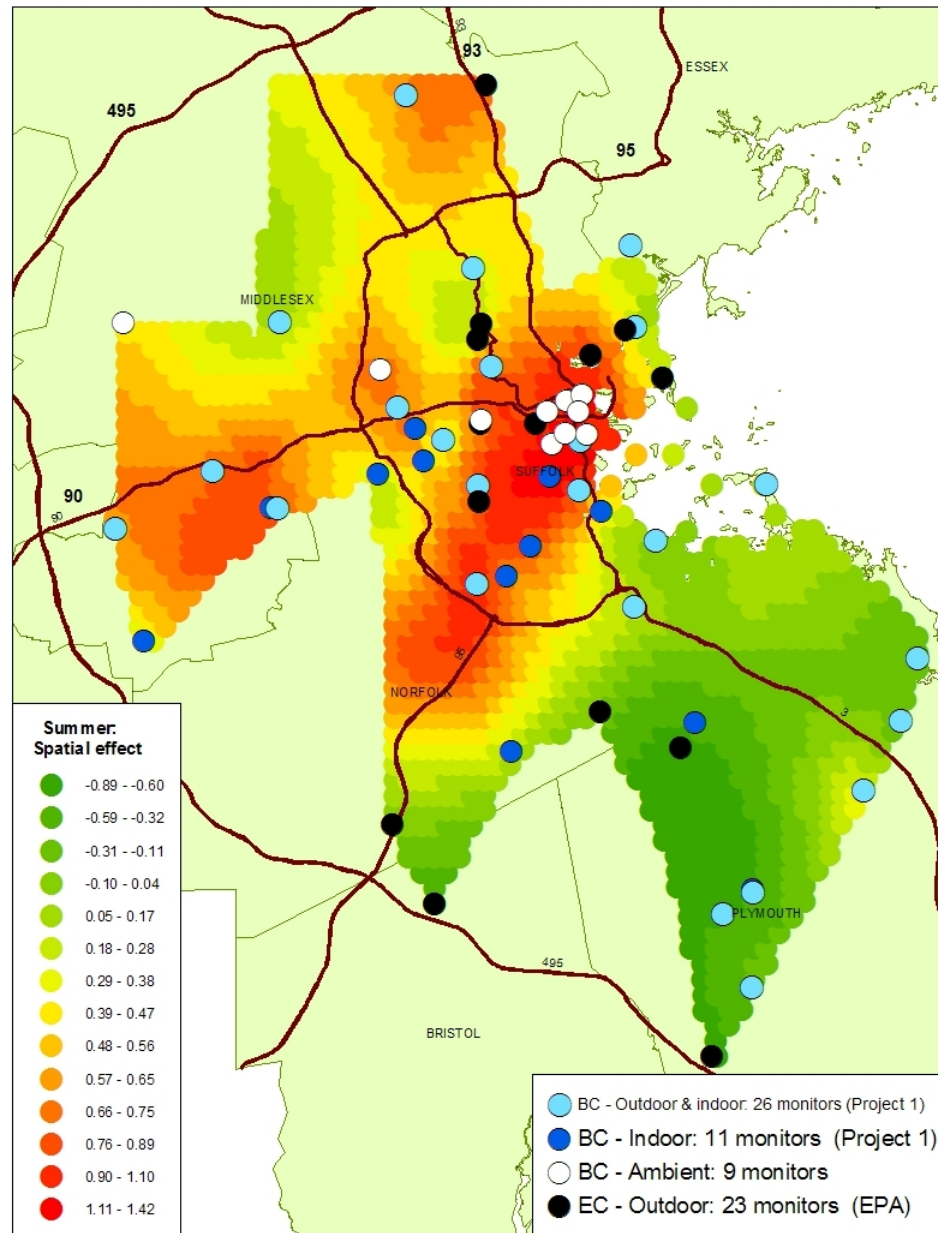
- Hypothesis: coronary and respiratory disease are associated with chronic exposure to particulate matter ($PM_{2.5}$ and PM_{10})
- Prospective cohort study of 122,000 female nurses
- PM data taken from EPA and government monitoring networks: 1985-2002
- Predictive space-time model with GIS-derived covariates
- Predictions made for each nurse's geocoded residence for each month, 1988-2002
- Proportional hazards survival modelling of health outcomes based on predicted exposure and personal covariates

Nurses' Health Study prediction



Latent variable modelling of traffic exposure in Boston

- Spatial latent variable model relating several pollutants to latent measure of traffic particles
- Predictive Bayesian space-time model with GIS-derived covariates
- Goal is to relate traffic exposure score to health outcomes in several local cohort studies
 - Birthweights in Boston
 - Diabetes cohort: heart rate variability (HRV) and inflammation markers (CRP/IL6)
 - Normative aging study: HRV and inflammation markers (CRP/IL6)
 - ...



Predictive Modelling

- Nurses' Health Study model:

$$Y_{i,t} \sim N\left(g_t(s_i) + \sum_{p=1}^P f(z_{i,p}), \sigma^2\right)$$

- $g_t(\cdot)$ represented as a thin-plate regression (knot-based) spline
- individual spatial surfaces for each month (large scale heterogeneity)
- Smooth terms of GIS covariates such as distance to roads, land use (small scale heterogeneity)
- fit via `gam()` and backfitting in R

- Boston model:

$$Y_{i,t} \sim N\left(g(s_i) + h(t) + \sum_{p=1}^P f(z_{i,p}), \sigma^2\right)$$

- $g(\cdot)$ represented as a thin-plate smoothing spline
- single spatial surface with smooth terms of time and GIS covariates
- fit via MCMC

Classical and Berkson measurement error

- Classical measurement error:

- covariate, X , is measured with error as W

$$H_i = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \epsilon_i$$

$$W_i = X_i + U_i$$

$$X \perp U$$

$$\text{Var}(W) = \text{Var}(X) + \text{Var}(U)$$

- Berkson measurement error

- covariate, X , is instead centered around a proxy

$$H_i = \beta_0 + \beta_1 X_i + \beta_2 Z_i + \epsilon_i$$

$$X_i = S_i + V_i$$

$$S \perp V$$

$$\text{Var}(X) = \text{Var}(S) + \text{Var}(V)$$

Regression Calibration

- In classical measurement error, replace X with $E(X|W, Z)$
 - simple setting: $E(X|W) = \frac{\sigma_x^2}{\sigma_x^2 + \sigma_u^2} W$
 - linear regression: regression calibration is unbiased for β_1
 - logistic regression:
 - * approximate bias, $\frac{\beta_1}{(1 + \beta_1^2 0.59^2 \sigma_{X|W}^2)^{1/2}}$, is small if β_1 is small
 - survival analysis: bias is small if effect is small (e.g., relative risk < 2)
- Berkson error: $S = E(X)$
 - unbiased in linear regression
 - bias should be small for logistic and Cox regression if effect sizes are small
 - regression calibration produces a Berkson structure ($X = E(X|W) + V$)

Spatial smoothing as regression calibration

- The principle
 - Kriging/Gaussian process modelling/Bayesian smoothing act as regression calibration
 - * $S = E(X|Y)$; $X = S + V$
 - Mixed model prediction acts as regression calibration
 - * the BLUP is the expected value of the spatial random effects, $S = E(X|Y)$
 - Other smoothers are likely to give similar predictions, so should mimic regression calibration
- The practice (in the Nurses' Health Study)

$$\begin{aligned} \text{Var}(X) &= 0.18 & \text{Var}(S) &= 0.15 \\ \text{Cor}(X, U) &= -0.44 & \text{Cor}(S, V) &= -0.18 \end{aligned}$$

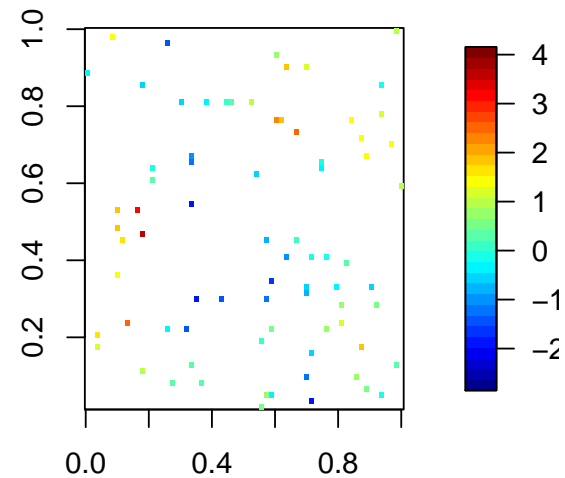
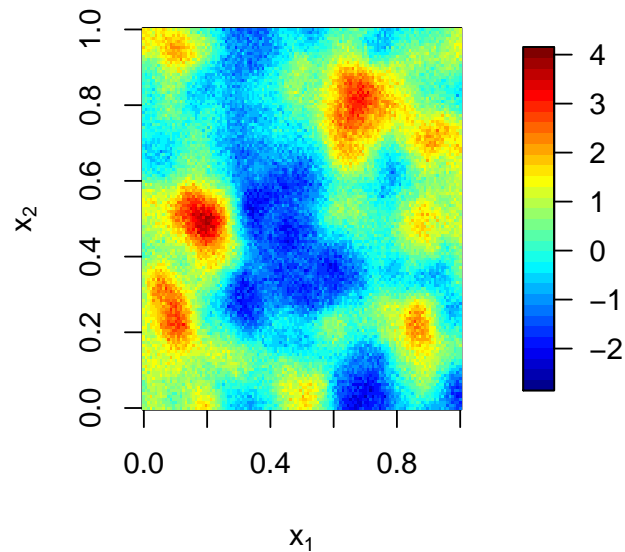
$X = S + V$ is a better model than $S = X + U$

Adjusting for measurement error

1. Use the smoothed estimates directly
2. Joint Bayesian modelling of health outcomes and exposure data, accounting for heteroscedasticity and correlation of smoothed co-variate estimates
3. Sample from the exposure distribution and fit multiple health models to account for uncertainty
 - a bad idea as the sampling moves the situation from Berkson error back to classical error and induces bias
4. Cross-validation to assess under- or over-smoothing and adjust the naive estimate:
possible model: $X = \gamma_0 + \gamma_1 S + V$
 $\hat{\beta}_{1,adj} = \hat{\beta}_1 / \hat{\gamma}_1$ where $\hat{\gamma}_1$ is estimated as the slope from regressing held-out observations on smoothed predictions
 - for NHS, $\hat{\gamma}_1 = 0.88$ (need to adjust for smoothing bias)

Simulation results

	Easy			Harder		
	bias	MSE	coverage	bias	MSE	coverage
true exposure				-0.002	0.003	95.2%
classical smoother	-0.013	0.011	93%	0.070	0.031	74.4%
classical with sampling	0.124	0.028	86%	0.591	0.375	1.2%
classical with γ correction				-0.018	0.028	95.2%
Bayesian	-0.031	0.016	99%	-0.114	0.083	89.0%



Conclusions

- Predictive space-time modelling of exposure induces measurement error
- Error is of the Berkson type, which in principle induces limited bias
- For continuous outcomes, some adjustments can improve estimation, particularly if smoothing problem is hard (e.g., sparse data)
- Further work is needed in the case of survival outcomes