

## Lecture 3

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## 1 Method of Dependency Graphs

We recall the definition of a dependency graph from the previous lecture. For a collection of random variables  $\{X_i, i \in V\}$  indexed by the vertices  $V$  of a graph  $G = (V, E)$  we say that  $G$  is a dependency graph if for any disjoint subsets  $S, T \subseteq V$  with no edges between  $S$  and  $T$  we have  $\{X_i, i \in S\}$  and  $\{X_i, i \in T\}$  independent. Let  $D = 1 + \max \text{degree } G$ .

**Lemma 1** Suppose that  $E(X_i) = 0, \sigma^2 = \text{Var}(\sum X_i), W = \frac{\sum X_i}{\sigma}$  and  $Z \sim N(0, 1)$ . Then

$$\text{Wass}(W, Z) \leq \frac{4}{\sqrt{\pi}\sigma^2} \sqrt{D^3 \sum E|X_i|^4} + \frac{D^2}{\sigma^3} \sum E|X_i|^3.$$

**Proof:** Let  $W_i = \frac{1}{\sigma} \sum_{j \in N_i} X_j$  where  $N_i = \{i\} \cup \{\text{neighbours of } i\}$ . As in the case of iid random variables we have  $X_i$  and  $W_i$  independent but we do not in general have that  $W_i$  and  $W - W_i$  are independent.

Take any  $f$  such that

$$|f| \leq 1, \quad |f'| \leq \sqrt{\frac{2}{\pi}}, |f''| \leq 2.$$

Then

$$EWf(W) = \frac{1}{\sigma} \sum E(X_i f(W)) = \frac{1}{\sigma} \sum E(X_i (f(W) - f(W_i))) = (I) + (II)$$

where

$$(I) = \frac{1}{\sigma} \sum E[X_i (f(W) - f(W_i) - (W - W_i) f'(W))]$$

and

$$(II) = \frac{1}{\sigma} \sum E[X_i (W - W_i) f'(W)].$$

Now

$$(I) \leq \frac{1}{\sigma} \sum \frac{1}{2} E|X_i (W - W_i)^2| |f''|_{\infty} \leq \frac{1}{\sigma^3} \sum E|X_i (\sum_{j \in N_i} X_j)^2|$$

since  $W - W_i = \frac{1}{\sigma} \sum_{j \in N_i} X_j$ . Also

$$(II) = \frac{1}{\sigma} \sum E X_i \left( \sum_{j \in N_i} X_j f(W) \right) = E \left( f'(W) \underbrace{\left[ \frac{1}{\sigma^2} \sum_{j \in N_i} X_i \left( \sum_{j \in N_i} X_j \right) \right]}_T \right).$$

We will proceed by showing that  $T$  is concentrated. Note that since  $E X_i W_i = 0$ ,

$$\frac{1}{\sigma} E \sum X_i (W - W_i) = \frac{1}{\sigma} E \sum X_i W = E W^2 = 1$$

and so

$$|(II) - f'(W)| = |E(f'(W)(T - 1))| \leq |f'|_{\infty} E|T - 1| \leq \sqrt{\frac{2}{\pi}} \sqrt{E(T - 1)^2} = \sqrt{\frac{2}{\pi}} \sqrt{\text{Var}(T)}$$

Combining these results we have

$$|E W f(W) - E f'(W)| \leq \sqrt{\frac{2}{\pi}} \sqrt{\underbrace{\text{Var}\left(\frac{1}{\sigma^2} \sum_{j \in N_i} X_i \left( \sum_{j \in N_i} X_j \right)\right)}_{III} + \underbrace{\frac{1}{\sigma^3} \sum E |X_i \left( \sum_{j \in N_i} X_j \right)^2|}_{IV}}.$$

Now

$$\begin{aligned} IV &\leq \frac{1}{\sigma^3} \sum_i \sum_{j, k \in N_i} E |X_i X_j X_k| \\ &\leq \frac{1}{\sigma^3} \sum_i \sum_{j, k \in N_i} \frac{1}{3} (E |X_i|^3 + E |X_j|^3 + E |X_k|^3) \\ &\leq \frac{D^2}{\sigma^3} \sum E |X_i|^3 \end{aligned}$$

where the second inequality follows from the AM-GM inequality.

Next we need to estimate

$$\text{Var}\left(\sum_{i, j \in N_i} X_i X_j\right).$$

The collection  $\{X_i X_j, i \in V, j \in N_i\}$  is a collection with a dependency graph of maximum degree  $2D^2$ . This can be seen as follows:  $X_i X_j$  is independent of  $X_k X_l$  if neither  $k$  nor  $l$  belongs to  $N_i \cup N_j$ . Now  $|N_i \cup N_j| \leq 2D$  and each vertex in this set has at most  $D$  neighbors so the maximum degree of the new dependency graph is  $2D^2$ . Using the variance bound on sums of dependency graph variables derived in the previous lecture we have

$$\text{Var}\left(\sum_{i, j \in N_i} X_i X_j\right) \leq 2D^2 \sum_{i \sim j} \text{Var}(X_i X_j) \leq 2D^3 \sum E X_i^4$$

using the fact that

$$\text{Var}(X_i X_j) \leq E(X_i^2 X_j^2) \leq \frac{1}{2} E X_i^4 + \frac{1}{2} E X_j^4.$$

The proof is completed by substituting this estimate.  $\square$

**Example 2** Let  $Y_1, \dots, Y_{n+1}$  be iid mean 0 variance 1 random variables and let  $X_i = Y_i Y_{i+1}$ . A dependency graph for the  $X_i$  has edge set  $\{(i, i+1) : 1 \leq i \leq n\}$  and  $D = 3$ . Then  $\text{Var}(\sum X_i) = \sigma^2 = Cn$  so

$$\text{Wass}\left(\frac{1}{\sigma} \sum X_i, Z\right) \leq C \frac{1}{\sigma^2} \sqrt{D^3 \sum E X_i^4} + \frac{D^2}{\sigma^3} \sum E |X_i|^3 \leq \frac{c}{\sqrt{n}}$$

**Exercise 3** In an Erdos-Renyi random graph  $G(n, p)$  let  $T_{n,p}$  be the number of triangles. Using the method of dependency graphs show that for some absolute constant  $C$

$$\text{Wass}\left(\frac{T_{n,p} - E T_{n,p}}{\sqrt{\text{Var}(T_{n,p})}}, Z\right) \leq \frac{C}{np^{9/2}}$$

**Exercise 4** • Find out the best known result for the above problem.

- Show that the CLT can not hold if  $np \not\rightarrow \infty$ .
- Refine the method of dependency graphs to show a CLT when  $np \rightarrow \infty$ .