

Lecture 33

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1 Zero Bias Coupling

Consider a random variable W with $EW = 0$ and $EW^2 = \sigma^2 < \infty$. Define

$$\rho(x) = \frac{1}{\sigma^2} E(W1_{(W \geq x)}).$$

Lemma 1. ρ is a probability density

Proof:

- $\rho(x) \geq 0$ for all x . For $x \geq 0$ this is obvious, for $x < 0$ note $E(W1_{(W \geq x)}) = -E(W1_{(W < x)}) \geq 0$
- $\int \rho(x) dx = 1$:

$$\int \rho(x) dx = \int_{x < 0} -E(W1_{(W < x)}) dx + \int_{x \geq 0} E(W1_{(W \geq x)}) dx$$

Let μ denote the law of W . Then

$$\int_{x \geq 0} E(W1_{(W \geq x)}) dx = \int_{x \geq 0} \int_{y \geq x} y d\mu(y) dx = \int_{y \geq 0} \int_{x \leq y} y dx d\mu(y) = \int_{y \geq 0} y^2 d\mu(y)$$

and similarly for the second term. \square

The distribution corresponding to this density is called the “zero bias transform”. If W^* is a random variable following the zero-bias transform of the law of W , then for all absolutely continuous ϕ we have

$$EW\phi(W) = \sigma^2 E\phi'(W^*)$$

(this is immediate from integration by parts). If $W \sim N(0, 1)$ then

$$\rho(x) = \int_{y \geq x} \frac{ye^{-y^2/2}}{\sqrt{2\pi}} dy = \frac{e^{-x^2/2}}{\sqrt{2\pi}}$$

so that $W \stackrel{d}{=} W^*$, i.e. the standard normal distribution is a fixed point of the zero bias transform.

Example: $W \sim \pm 1$. Then $W^* \sim Uni[-1, 1]$.

Theorem 1. *If $EW = 1$, $EW^2 = 1$, then $Wass(W, Z) \leq 2Wass(W, W^*)$*

Proof: Take any 1-Lipschitz h . Find ϕ such that $\phi'(x) - x\phi(x) = h(x) - Eh(Z)$. By earlier results we know $\|\phi''\|_\infty \leq 2\|h'\|_\infty \leq 2$. Now suppose W and W^* live on the same space. Then

$$|E(h(W)) - E(h(Z))| = |E(\phi'(W) - W\phi(W))| = |E(\phi'(W) - \phi'(W^*))| \leq 2E|W^* - W|$$

Since this is true for any coupling, any h Lipschitz we have

$$Wass(W, Z) \leq 2 \inf_{W, W^*} E|W^* - W| = 2Wass(W, W^*)$$

Example: suppose X_1, \dots, X_n i.i.d. mean 0 variance 1, $W = \frac{\sum X_i}{\sqrt{n}}$. For each i let X_i be independent of everything else. Let $I \sim Unif\{1, \dots, n\}$ independent of everything else. Define

$$W^* := \frac{1}{\sqrt{n}} \left[\sum_{i \neq I} X_i + X_I^* \right]$$

We claim this is a zero bias transform:

$$\begin{aligned} EW\phi(W) &= \frac{1}{\sqrt{n}} \sum EX_i\phi(W) = \\ &= \frac{1}{\sqrt{n}} \sum E \left(X_i \phi \left(\frac{1}{\sqrt{n}} \left[\sum_{j \neq i} X_j + X_i \right] \right) \right) \\ &= \frac{1}{\sqrt{n}} \sum E \left(\phi' \left(\frac{1}{\sqrt{n}} \left[\sum_{j \neq i} X_j + X_i^* \right] \right) \frac{1}{\sqrt{n}} \right) \\ &= E(\phi'(W^*)) \end{aligned}$$

Thus

$$Wass(W, Z) \leq 2E|W^* - W| = 2E \left| \frac{1}{\sqrt{n}}(X_I - X_I^*) \right| = 2E \left| \frac{1}{\sqrt{n}}(X_1 - X_1^*) \right| \leq \dots$$

Theorem 2. (Goldstein-Reinert) *Suppose Y, Y' are an exchangeable pair, $EY = 0$, $EY^2 = \sigma^2$ and $E(Y'|Y) = (1 - \lambda)Y$. Let ν denote the joint distribution of (Y, Y') . Let*

$$d\mu(y, y') = \frac{(y - y')^2}{E(Y - Y')^2} d\nu(y, y').$$

Suppose $(\hat{Y}, \hat{Y}') \sim \mu$. Let $U \sim Uni[0, 1]$ independent of all else. $Y^ = U\hat{Y} + (1 - U)\hat{Y}'$. Then Y^* is a zero bias transform of Y .*

Note that $Ef'(Ua + (1 - U)b) = \frac{f(b) - f(a)}{b - a}$, hence

$$\begin{aligned} \sigma^2 E(f'(Y^*)) &= \sigma^2 E\left(\frac{f(\hat{Y}) - f(\hat{Y}')}{\hat{Y} - \hat{Y}'}\right) \\ &= \frac{\sigma^2}{E(Y - Y')^2} E[(Y - Y')(f(Y) - f(Y'))] = \frac{2\lambda\sigma^2 E[Yf(Y)]}{E(Y - Y')^2} = EYf(Y) \end{aligned}$$

Exercise: Try to get Hoeffding CLT using this method.

Exercise: Suppose $EW\phi(W) = ET\phi'(W)$. Then $E(T|W)$ is the density of $Law(W^*)$ w.r.t. $Law(W)$ evaluated at W . We have Tusnády's lemma based on concentration of T . On the other hand, if W, W^* can be constructed to be close to each other then we can construct W, Z such that $E|W - Z|$ is small. Question: If we know tail bounds on $|W - W^*|$ can we construct (W, Z) with fast decaying tails for $W - Z$?